Aleutian Islands Golden King Crab Model Scenarios for May 2022 Assessment

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M.S.M. Siddeek ${ }^{1}$, J. Zheng ${ }^{1}$, C. Siddon ${ }^{1}$, B. Daly ${ }^{2}$, M.J. Westphal ${ }^{3}$, and L. Hulbert ${ }^{1}$<br>${ }^{1}$ Alaska Department of Fish and Game, Division of Commercial Fisheries, P.O. Box 115526, Juneau, Alaska 99811<br>${ }^{2}$ Alaska Department of Fish and Game, Division of Commercial Fisheries, 351 Research Ct., Kodiak, Alaska 99615<br>${ }^{3}$ Alaska Department of Fish and Game, Division of Commercial Fisheries, PO Box 920587, Dutch Harbor, Alaska 99692.

## Preamble

In this report, we provide a set of model scenarios that could be selected for May 2022 assessment, and OFL and ABC determinations for the Aleutian Islands golden king crab stock. The scenarios are based on May 2021 CPT and June 2021 SSC recommendations. This document does not follow the standard SAFE document format. Standard SAFE document will be presented at the May 2022 CPT meeting.

## Highlights:

1. Following May 2021 CPT and June 2021 SSC concerns on currently implemented CPUE standardization procedure,

Several methodological improvements were made, including addressing the number of degrees of freedom in the smoothers and demonstrating spline fits to raw CPUE. Only statistically significant degrees of freedom were considered.
2. Further improvement on Year:Area interaction CPUE analysis was done (Appendix A).

Inclusion of Year:Area interaction addresses the area shrinkage issue as a result of reduction in number of vessels during the post-rationalization period.
3. Three core models were formulated considering different CPUE standardization procedures (main effects CPUE, 21.1a; Year:Area interaction CPUE, 21.1b) and a different set of catchability and additional CVs (three catchability and additional CPUE standard errors, 21.1c). The May /June 2021 accepted model 21.1a was considered as the base model with a few modifications akin to Gmacs model formulation.
4. Three additional models were considered to address the effect of higher (knife-edge) maturity size on mature male abundance estimates and to oblige with a fishing industry request to investigate the effect of omitting one (underperforming) vessel on CPUE indices and reference points in WAG:

Model 21.1a2: Model21.1a + knife-edge maturity was changed from 111 mm carapace length (CL) to 116 mm CL (lower limits of the size bins in which actual size-at-maturity fell).
Model 21.1b2: Model21.1b + knife-edge maturity was changed from $111 \mathrm{~mm} C L$ to 116 mm CL.
Model 21.1d for WAG: Model21.1a + CPUE indices were estimated omitting one vessel.
5. EAG model 21.1a was modified to models 21.6 and 21 _ 7 to implement in Gmacs. Comparison of some results are provided in Appendix E.
6. Following June 2021 SSC request,

Preliminary summary statistics, comparing RACE AIGKC Slope Survey indices with Observer CPUE indices, are presented in Appendix D to solicit CPT and SSC guidance on how to proceed with incorporating slope survey indices into golden king crab assessment model.

For detailed accounts of the Aleutian Islands golden king crab model formulation, fisheries, and biology, we direct you to the stock assessment report presented at the May 2021 CPT and June 2021 SSC meetings (Siddeek et al. 2021).

## Input Data

- The input data presented at the May 2021 CPT meeting were updated after completion of the fisheries. Thus, the time series of data used in the model were retained catch (1981/82-2020/21), total catch (1990/91-2020/21), and groundfish bycatch (1989/902020/21) biomass and size compositions.
- Observer pot sample legal size crab CPUE data were standardized by the generalized linear model (GLM) with the negative binomial link function with variable selection by first CAIC (modified AIC) and followed by R square criterion, separately for 1995/962004/05 (pre-rationalization) and 2005/06-2020/21 (post-rationalization) periods. Fish ticket retained CPUE were standardized by the GLM with the negative binomial link functions for the 1985/86-1998/98 period (see Appendix A).
- A Year and Area interaction was considered in one model, 21.1b, to estimate a set of observer pot sample CPUE indices for the pre- and post-rationalization periods. Area was defined based on observer sample locations within 1 nmi x 1 nmi grids to reflect fishing footprints.
- The 2018/19-2020/21 male golden king crab chela height and carapace length measurement data were re-analyzed to update knife-edge maturity size estimates.

Table A lists a brief description of various models analyzed in this report.

Table A. Features of all model scenarios: Initial condition was estimated in year 1960 by the equilibrium condition; two catchability and two sets of logistic total selectivity curves were used for the pre- and post-rationalization periods for all models except 21.1c; and a common $M$ of $0.21 \mathrm{yr}^{-1}$ based on the estimate from the combined EAG and WAG data was used. The effective sample sizes for size compositions were estimated in two stages: Stage-1: number of vessel days/trips and Stage-2: Francis re-iteration method.

| Model | CPUE data type and modeling consideration | Knife-edge Maturity Size (Lower limit of the size bin) | Period for mean number of recruit calculation for (a) initial equilibrium abundance composition and (b) reference points estimations |
| :---: | :---: | :---: | :---: |
| 21.1a (accepted model in May/June 2021, implemented with up to 2020/21 data) | Observer data from 1995/96-2020/21; fish ticket data from 1985/86-1998/99; two catchability and total selectivity for the 19602004 and 2005-2020 periods, one retention and groundfish bycatch selectivity; and observer and Fish Ticket CPUE standardization by negative binomial models. | 111 mm CL | 1987-2017 |
| 21.1a1 | $21.1 \mathrm{a}+$ consider an $M$ of $0.38 \mathrm{yr}^{-1}$ for years $>1998$ (to address the retrospective issue on EAG assessment) | 111 mm CL | 1987-2017 |
| 21.1 b | 21.1a+ consider observer CPUE standardized with Year:Area interaction. | 111 mm CL | 1987-2017 |
| 21.1c | 21.1a+ consider three catchability and additional CPUE CVs (fish ticket: 1985-1994, observer: 1995-2004, and 2005-2020). | 111 mm CL | 1987-2017 |
| 21.1d | $21.1 \mathrm{a}+$ CPUE indices estimated omitting one vessel for WAG. | 111 mm CL | 1987-2017 |
| 21.1a2 | 21.1a+ higher knife-edge maturity | 116 mm CL | 1987-2017 |
| 21.1 b 2 | $21.1 \mathrm{~b}+$ higher knife-edge maturity | 116 mm CL | 1987-2017 |
| 21.6 \& 21.7 | 21.1a (EAG)+ modified for Gmacs input. | 111 mm CL | 1987-2017 |

## Response to May 2021 CPT comments

## Comment 1:

The analysis of the maturity data should be repeated using, for example, the methods of Olson et al. (2018) and Somerton and Macintosh (1983). The results of the analyses should be presented to the CPT.
Response:
The analysis was repeated following Olson et al. 's approach of directly fitting chela height against carapace length by the bend point analysis package available in $R$. The focus was determining the knife-edge maturity rather than establishing a maturity curve (details are in Appendix B).

## Comment 2:

Consider including the NMFS Aleutian Islands trawl survey as an additional index of abundance. The first step in this process should be to compare the depths at which the survey is conducted to those at which AI golden king crab are found/fished.

Response:
A preliminary comparison figures of NMFS Aleutian Islands trawl survey index of abundance vs observer CPUE index for comparable years and areas/depths was detailed in Appendix D. The purpose was to solicit the CPT and SSC advice on how to incorporate NMFS indices into GKC assessment model.

## Comment 3:

The CPUE standardization for the post rationalization years:
o explore why the index for the WAG is lower in the last three years based on area*year interactions;
o explore why the index for the WAG is more precise in the earlier years based on area*year interactions; and

- better justify the degrees of freedom for smooths and plot the smooths.

Response:
The CPUE standardization procedure was revamped with special attention given to selecting nonsignificant predictor variable coefficients in the final GLM models and ascertaining that the final models' predictor variables were non-collinear. In this process, several (above) concerns were addressed. The degrees offreedom of selected smoother variables had drastically reduced (details are in Appendix A).

The predicted area*year interaction curve was compared with the input area*year interaction curve in a separate plot in Figure 17 for WAG. It indicated that the index was no more precise in early years than later years.

Figures A. 3 to A. 6 in Appendix A depict the fit of smoothers to observed CPUE data for a range of Soak time values at a given set of fixed values of other predictor variables chosen in the final models separately for EAG and WAG. The smoothers in the final model appeared to adequately trace through raw CPUE data. For simplicity, the fits were shown for arbitrarily selected years.

## Comment 4:

The specifications of smooths when analyzing the cooperative survey should be selected using the survey data and not taken from analyses of other indices.
Response:
The cooperative survey data analysis was not taken up in this run. Will address this issue and follow the suggestion at the May2022 CPT.

## Comment 5:

The negative log likelihood for model 21.1b (three total selectivity) is larger than that for model 21.1a model even though model 21.1a is nested within model 21.1b. This should not be, perhaps this model optimized at a local minimum. Furthermore,

Model 21.1 b is unable to provide a better fit to the length-frequency data for the EAG. The reasons for the change in total length-frequency in recent years need to be better understood before new models were formulated. Edward Poulsen noted that the number of vessels in the EAG was less in recent years than before and that the higher CPUE areas tend to have higher abundance of smaller animals, which may be part of the reason for the change in the total length-frequency.

Response:
Previous model 21.1b was no longer considered in the current set of models. The same model name was used for Year:Area interaction CPUE model in the current analysis

## Comment 6:

$92 \%$ of the WAG TAC is taken at the time of the meeting. Adjusting the catches to reflect the final catch is not likely to impact the TAC set by the State (which is usually well below the ABC). However, future assessments should be based on the best projection of total catch when the season is not complete.

Response:

If this situation occurs for the 2021/22 season, we will consider the above recommendation in the assessment.

## Comment 7:

Progress towards further GMACS implementation for this stock is expected for the next cycle in 2022.

## Response:

That is the plan. Current progress is presented in Appendix E.

## Comment 8:

Address SSC concerns that "how many years it takes crab recruited to the model to recruit to the fishery, i.e., size at first selectivity. This could inform the last year of the period to be included for mean R calculation.

## Response:

We will address this issue at the May2022 CPT meeting.

## Comment 9:

Presentational

- Correct the x -axis labeling in Fig. CPT2.
- Colors should be used to distinguish observed and predicted lengthfrequencies in Figures 11-13. However, it would be better to use plots such as Figures 11-13 to show observed length-frequencies and plots of observed vs. predicted length-frequencies (with results shown for multiple models) shown individually by year.
- The rationale for conducting separate assessments for the EAG and WAG should be integrated into the narrative of the assessment.
- Avoid showing fits of models such as 21.1 c to observed data used to fit different models.
- Plot selectivity for all models on the same plot to better allow comparisons.
- Use consistent y-axis ranges in similar figures - see Figure 12a (top panels do not go to 0 vs. bottom panels that do include 0 ).
- Include page numbers in the review draft.
- Increase line width in figures for easier viewing of model runs (e.g., Figures 14 and 32).

Response:

Several suggestions were followed in the current draft report. But the entire set of suggestions will be considered when presenting the final assessment in May 2022.

## Response to June 2021 SSC comments

## Comment 1:

SSC agreed to all the above CPT recommendations.
Response:
Please see responses to above CPT comments.

## Furthermore,

Comment 2:
For this year (2021), the CPT recommended continuing to apply a $25 \%$ buffer. The SSC instead recommends an increase to a $30 \%$ buffer from the maximum permissible ABC , based on: 1) the continued positive retrospective pattern in the EAG Model, 2) continued model convergence concerns indicating remaining parameter confounding (specifically, the jitter analysis for the 19.1 WAG model resulted in multiple solutions for MMB and B35\% at identical total likelihood values), and 3) the CPUE series which included a year:area interaction indicated a steeper decline in recent years than the series used in model 21.1a (the model accepted for harvest specifications).

## Response:

The issue of WAG model producing multiple solutions for MMB and B35\% at identical likelihood values was resolved in the current jitter analysis of model 21.1a. Although the contribution of groundfish bycatch to total removal is small, the likelihood emphasis factor interfered with the optimization. In this run, a higher likelihood weight of 0.5 (instead of 0.2) was used and the model produced identical solutions for MMB, B35\% and OFL at identical total likelihood values.

We will address other issues as soon as possible.

## Comment 3:

The author's rationale for continued use of two separate stock assessment models for the EAG and WAG is very helpful, and the SSC recognizes that this approach is reasonable. However, the SSC notes that sharing biological parameters and basic stock dynamics within a single assessment model that
has two largely independent areas modelled simultaneously may help address recurrent convergence and estimation challenges. The SSC suggests that such an approach be considered further, as either a replacement for the current approach, or as part of a multi-model evaluation.

Response:
Because of time constraint, we plan to do this for the May/June 2022 CPT/SSC meetings.

## Comment 4:

The SSC did not find the logistic fit to the maturity predictions based on the chela height to carapace length relationship to be compelling and supports the CPT's recommendations for additional modelling. However, the SSC also notes that direct observational data on maturity may ultimately be needed to resolve this process and recommends holding studies or any other research be considered.

Response:
Logistic fit to maturity was not considered in this run but focused on improving the method of bend point analysis following other publications (e.g., Olson et al. 2018). We agree with SSC's observation that direct observational data on mating activity are needed to identify the true minimum size at first maturity.

## Introduction

Genetic studies did not show any evidence for separate golden king stocks in the Aleutian Islands. CPUE trends suggest different factors may influence stock productivity in EAG and WAG, which are separated by the $174^{\circ}$ W longitude meridian. Since 1996, the Alaska Department of Fish and Game (ADF\&G) has divided management of the Aleutian Islands golden king fishery into EAG and WAG (ADF\&G 2002). The stocks in the two areas are managed with annual total allowable (retained) catches. Additional management measures include a male-only fishery and a minimum legal-size limit ( 152.4 mm CW , or approximately 136 mm CL ), which is at least one annual molt increment larger than the $50 \%$ maturity length of 120.8 mm CL for males estimated by Otto and Cummiskey (1985).

There is a paucity of information on golden king crab life history characteristics due in part to the deep depth ( $\sim 300-1000 \mathrm{~m}$ ) and extremely rough bottom distribution on the slopes and trenches and the asynchronous nature of life history events, growth, and reproduction (Otto and Cummiskey 1985; Somerton and Otto 1986; Watson et al. 2002).

Figures 1 and 2 provide the historical time series of catch and CPUE for EAG and WAG, respectively. Increases in CPUE were observed during the late 1990s through the early 2000s, and
with the implementation of crab rationalization in 2005. In 2012, the Board of Fisheries of Alaska (BOF) increased the TAC levels to 3.310 million pounds for EAG and 2.980 million pounds for WAG beginning with the 2012/13 fishing year. The below par fishery performance in WAG in middle 2010 years lead to reduction in TAC to 2.235 million pounds, which reflected a $25 \%$ reduction in the TAC for WAG, while the TAC for EAG was kept at the same level 3.31 million pounds for the 2015/16 through 2017/18 fishing seasons. With the improved fishery performance and stock status since 2017/18, the TACs were further increased to 2.5 million pounds for WAG and 3.856 million pounds for EAG in 2018/19 and 2.87 million pounds for WAG and 4.31 million pounds for EAG in 2019/20 fishing years.

A new harvest strategy based on model estimated mature male abundance was accepted by the BOF in March 2019, specifying a $15 \%$ maximum harvest rate for EAG and $20 \%$ maximum harvest rate for WAG, and was implemented first time for the 2019/20 fishery (Daly, et al., 2019). Based on the new harvest strategy, the TACs were set to 2.96 million pounds for WAG and 3.65 million pounds for EAG for the 2020/21 fishery, and to 2.32 million pounds for WAG and 3.61 million pounds for EAG for the 2021/22 fishery.

The EAG and WAG stocks were modelled separately for several reasons:
(a) Fishery catch data (e.g., CPUE magnitude and CPUE temporal trends) suggest that the productivity is different between the two areas.
(b) WAG has wider area of stock distribution compared to limited area distribution in EAG.
(c) The fishing areas are spatially separated with an area gap between EAG and WAG (Siddeek et al. 2021). Regions of low fishery catch suggest that availability of suitable habitat may vary longitudinally.
(d) Tagging studies have shown little mixing between the two areas (Watson and Gish 2002).
(e) Currents are known to be strong around the Aleutian Islands, thus larval mixing between the two regions may occur. Yet needed data to confirm larval drift trajectories or horizontal displacement are lacking. Unlike other king crabs, golden king crab females carry large, yolk-rich, eggs, which hatch into lecithotrophic (non-feeding) larvae that do not require a pelagic distribution for encountering food items. Depth at larval release, the lecithotrophic nature of larvae, and swimming inactivity in lab studies implies benthic distributions, which may limit larval drift between areas if horizontal current velocities are reduced at depth.
(f) Integrating contrasting data in one single model may provide parameter estimates in between the two extremes which would not be applicable to either (Richards 1991; Schnute and Hilborn 1993).
(g) Area specific assessment is superior to a holistic approach for this stock because of patchy nature of golden king crab distribution.
h) Alaska Board of Fisheries decided to manage the two areas with separate total allowable catches. i) Genetic analysis shows no significant differentiation between areas within the Aleutian Island population (Grant and Siddon 2018), thus there is no genetic support for subdividing this population; however, above listed factors support separate stock assessments in the two regions.

## Analytic Approach

The underlying population dynamics model was male-only and length-based (Siddeek et al. 2021). This model combined commercial retained catch, total catch, groundfish (trawl and pot) fishery
discarded catch, standardized observer legal size catch-per-unit-effort (CPUE) and commercial fishery CPUE indices, fishery retained catch size composition, total catch size composition, and tag recaptures by release-recapture length to estimate stock assessment parameters. Tagging data were used to calculate the size transition matrix.

The observer and commercial fishery CPUE indices with GLM estimated standard errors and additional constant standard errors were used in the model fit. The additional constant errors were estimated by the model. There were significant changes in fishing practice due to changes in management regulations (e.g., constant TAC since 1996/97 and crab rationalization since 2005/06), pot configuration (escape web on the pot door increased to 9 -inch since 1999), and improved observer coverage in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two sets of catchability and total selectivity parameters with only one set of retention parameters for the periods 1985/86-2004/05 and 2005/06-2020/21. Three catchability and additional constant CPUE standard errors were also considered in one model, 21.1c.

The equilibrium abundance in 1960 was projected with natural mortality and annual recruitment to create the initial abundance by size at the start of the fishery in 1981. The $\mathrm{R}_{0}$ for equilibrium abundance was determined using the average model estimated number of recruits for a selected period. The standardized CPUE indices, catch, and size composition information were used to determine the stock abundance trends in both regions. The observer and fish ticket CPUE indices were assumed to be linearly related to exploitable abundance. The $M$ was kept constant at 0.21 $\mathrm{yr}^{-1}$. The directed pot fishery discard mortality proportion was assumed at $0.20 \mathrm{yr}^{-1}$, overall groundfish fishery mortality proportion at $0.65 \mathrm{yr}^{-1}$ [mean of groundfish pot fishery mortality ( 0.5 $\mathrm{yr}^{-1}$ ) and groundfish trawl fishery mortality $\left(0.8 \mathrm{yr}^{-1}\right)$ ], and groundfish fishery selectivity at full selection for all length classes (i.e., selectivity $=1.0$ ). Any discard of legal-sized males in the directed pot fishery was not explicitly modeled and assumed to be insignificant.

The numbers of vessel-days were considered as the initial input effective sample sizes (i.e., stage1) for retained and total size compositions and numbers of trips for groundfish discard catch size composition without enforcing any upper limit. The groundfish size composition was not fitted in any model following an earlier CPT suggestion. The stage- 2 effective sample sizes were estimated iteratively from stage-1 effective sample sizes by the Francis (2011) method for all models.

Various weighting factors were used for catch biomass, recruitment deviation, pot fishery F, and groundfish fishery F . The retained catch biomass weight was set to an arbitrarily large value (500.0) because retained catches are more reliable than any other data sets. The total catch biomass weight was scaled in accordance with the observer annual pot sample sizes with a maximum of 250.0. The total catches were derived from observer nominal total CPUE and effort. In some years, observer sample sizes were low (Tables 3). A small groundfish bycatch weight was chosen based on the September 2015 CPT suggestion to lower its weight ( 0.2 for EAG and 0.5 for WAG). The best fit to groundfish bycatch data criteria was used to choose the lower weight for the groundfish bycatch. A higher weight for WAG groundfish bycatch likelihood was chosen to get the global maximum log likelihood in the jitter runs (see Appendix C). Note that groundfish bycatch of Aleutian Islands golden king crab was very low (Table 2). The CPUE weights were set to 1.0 for all models. The Burnham et al. (1987) suggested formula was used for $\ln (C P U E)$ [and $\ln (M M B)]$
variance estimation (formula given in Siddeek et al. 2021)). The CPUE index variances estimated from the negative binomial with additional constant variances appeared to have adequately fitted the model, as confirmed by the fit diagnostics (Fox and Weisberg 2011).

The AD Model Builder (Fournier et al. 2012) was used for model fitting.

## Results

Model equations and weights for different data sets are provided in Siddeek et al. (2021). These weights (with the corresponding coefficient of variations) adequately fitted various data under integrated model setting. All models considered molt probability parameters in addition to the linear growth increment and normal growth variability parameters to determine the size transition matrix.

In May 2019 assessment and before, the length-weight relationship of $W=a L^{b}$, based on 1991 weight vs. CL data, where $\mathrm{a}=3.725^{*} 10^{-4}, \mathrm{~b}=3.0896$, was used for biomass calculation from number of crab by length. The length-weight relationship parameters were updated in 2020 using cooperative survey collected data during 2018/19 with $\mathrm{a}=1.095^{*} 10^{-4}, \mathrm{~b}=3.35923$. Furthermore, the crab weight in a size bin was calculated using Beyer's (1987) formula, which appropriately considers integration through lower $\left(\mathrm{CL}_{1}\right)$ limit to upper $\left(\mathrm{CL}_{u}\right)$ limit of a size bin:
$W_{l}=\left(\frac{1}{C L_{u}-C L_{l}}\right)\left(\frac{a}{1+b}\right)\left(C L_{u}^{b+1}-C L_{l}^{b+1}\right)$
The CPT/SSC/Council plan is to bring all crab assessment models into the generalized Gmacs framework. Some results from Gmacs implementation of model 21.1a for EAG were compared with that of the original 21.1a model in Appendix E.

## Tables of input values and parameter estimates

a. Historical GHL, TAC, catch, effort, CPUE, and mean crab weight are summarized in Table 1 for EAG and WAG.
b. Time series of retained and total catch and groundfish fishery discard mortality are summarized in Table 2 for EAG and WAG.
c. Time series of pot fishery and observer nominal retained and total CPUE, annual pot fishing effort, observer sample size, estimated observer CPUE indices are listed in Table 3 for EAG and WAG.
d. The estimated commercial fishery (fish ticket) CPUE indices are provided in Table 4 for EAG and WAG. The CPUE index estimation methods and fits are described in Appendix A.
e. The parameter estimates with coefficient of variation for three models 21.1 a (base), 21.1b, and 21.1c are summarized respectively in Tables 5 for EAG and 10 for WAG. The boundaries for parameter searches are also provided in those tables, and the estimates are within the bounds.
f. The mature male and legal male abundance time series for models 21.1a (base), 21.1b, and 21.1c are summarized in Tables 6, 7, and 8 for EAG and for models 21.1a (base), 21.1b, 21.1 c , and 21.1 d are summarized in Tables $11,12,13$, and 14 for WAG.
g. The recruitment estimates for those model scenarios are summarized in Tables 6 to 8 for EAG and Tables 11 to 14 for WAG.
h. The likelihood component values and the total likelihood values for models 21.1a, 21.1b, 21.1c, 21.1a2, and 21.1b2 are summarized in Table 9 for EAG and for models 21.1a, 21.1b, 21.1c, 21.1d, 21.1a2, and 21.1b2 are summarized in Table 15 for WAG.
i. The Tier level, $\mathrm{MMB}_{35 \%}$, current MMB, current MMB/MMB $35 \%$, $M$, $\mathrm{F}_{\mathrm{FL}}, \mathrm{F}_{35 \%}$, OFL, and ABC (under $25 \%$ and $30 \%$ buffers) for EAG, WAG, and the entire Aleutian Islands (AI) are listed in Table 16 (models 21.1a, 21.1b, 21.1c, 21.1d, 21.1a2, and 21.1b2 for WAG; and 21.1a, 21.1b, 21.1c, 21.1a2, and 21.1b2 for EAG). The status of the stock in EAG is estimated to be in Tier 3a for all models except model 21.1c whereas the status of the stock in WAG is determined to be in Tier 3b for all models. The respective reference points are added disregarding the stock status to estimate the reference points for the entire AI.

## Graphs of estimates

a. The retained length composition fits are provided in Figures 3a, 3b, and 3c for EAG and Figures 13a, 13b, and 13c for WAG, total length composition fits in Figures 4a, 4b, and 4c for EAG and Figures 14a, 14b, and 14c for WAG, and groundfish discarded catch length composition fits in Figures 5a, 5b, and 5c for EAG and Figures 15a, 15b, and 15c for WAG for $21.1 \mathrm{a}, 21.1 \mathrm{~b}$, and 21.1 c models, respectively. The retained and total catch size composition fits appear satisfactory for most years but the fits to groundfish bycatch size compositions are bad.
b. The pre- and post-rationalization periods' total and retained selectivity curves are provided in Figure 6 for EAG and Figure 16 for WAG for 21.1a, 21.1b, and 21.1c models. Total selectivity for the pre-rationalization period is used in the tagging model. The groundfish bycatch selectivity appeared flat in the preliminary analysis, indicating that all size groups are vulnerable to this gear. This is also shown in the size compositions of groundfish bycatch (Figures $5 \mathrm{a}-\mathrm{c}$ and $15 \mathrm{a}-\mathrm{c}$ ).
c. The CPUE fits by 21.1a, 21.1b, and 21.1c models are provided in Figure 7 for EAG and CPUE fits by 21.1a, 21.1b, 21.1c, and 21.1d models are given in Figure 17 for WAG. The CPUE trend of model 21.1c differed from those of other models in both management areas.
d. The recruitment trends for 21.1a, 21.1b, and 21.1c model fits are shown in Figure 8 for EAG and that for 21.1a, 21.1b, 21.1c, and 21.1d model fits are given in Figure 18 for WAG. The recruitment pulse peaked in 1988 and was high during 2016-2019 for all model fits in EAG. On the other hand, large recruitment pulses occurred during 1984-1989 but stabilized in recent years for all model fits in WAG.
e. The fits to retained catch, total catch, and groundfish discarded catch by 21.1a, 21.1b, and 21.1c models are provided in Figure 9 for EAG and that by 21.1a, 21.1b, 21.1c, and 21.1d models are given in Figure 19 for WAG. The retained and ground fish bycatch fits are adequate, but the total catch fits showed some discrepancy.
f. The fits to pre -1985 retained catches by $21.1 \mathrm{a}, 21.1 \mathrm{~b}$, and 21.1 c models are shown in Figure 10 for EAG and that by 21.1a, 21.1b, 21.1c, and 21.1d models are given in Figure

20 for WAG. All models adequately fitted the 1981/82-1984/85 retained catches in both areas.
g. Pot fishery total fishing mortality (F) plots for 21.1a, 21.1b, and 21.1c models for EAG (left) and for 21.1a, 21.1b, 21.1c, and 21.1d models for WAG (right) are shown in Figure 11. The F values increased during 1988-1992 and 1995 and systematically declined thereafter in the EAG. Slight increases in F were observed from 2015 to 2019, followed by a decline in 2020 in the EAG. On the other hand, the F in the WAG increased in 19861992 and 1994-2001, declined in late 2000s, and slightly increased in 2019 and 2020.
h. The MMB trends for 21.1a, 21.1b, and 21.1c models for EAG (left) and that for 21.1a, 21.1b, 21.1c, and 21.1d models for WAG (right) are shown in Figure 12. The MMB plots for the long time series (1960/61-2020/21) is shown at the top and for the short time series (2005/06-2020/21) is depicted at the bottom. The MMB systematically increased since 2017 in the EAG, but the increase was mild in the WAG.
i. The retrospective pattern of MMB has been an issue for EAG. It is investigated by comparing the status quo model (i.e., 21.1a) retrospective trends of MMB with that of higher $M$ model (21.1a1) and different catchability and additional CPUE standard error model (21.1c). Model 21.1a1 assumes a high $M$ of $0.38 \mathrm{yr}^{-1}$ (Siddeek et al. 2002) for years $>1998$. The six-year retrospective patterns for models 21.1a, 21.1a1, and 21.1c for EAG are compared in Figure 21. The Mohn rho values for the three models ranged from 0.4011 to 0.5092 with the lowest value determined by model 21.1a1. These values suggest that there is no significant improvement achieved in the retrospective patterns over the base model.

## Specification of the Tier level

The OFL and ABC for Aleutian Islands golden king crab stocks are determined under Tier 3 level. The calculation procedures are described below:

The critical assumptions for $M M B_{M S Y}$ reference point estimation of Aleutian Islands golden king crab are:
a. Natural mortality is constant, $0.21\left(\mathrm{yr}^{-1}\right)$.
b. Growth transition matrix is fixed and estimated using tagging data with the molt probability sub-model.
c. Total fishery selectivity and retention curves are length dependent and the 2005/062020/21 period selectivity estimates are used.
d. Groundfish bycatch fishery selectivity is kept constant at 1.0 for all length groups.
e. Model estimated recruits (in millions of crab) are averaged for the period 1987-2017.
f. Model estimated groundfish bycatch mortality values are averaged for the period 2011/12 - 2020/21 (10 years).
g. Knife-edge minimum maturity size of 111 mm CL is used for MMB estimation for all models except 21.1a2 and 21.1b2, which considered 116 mm CL.

## Method:

We simulated the population abundance starting from the model estimated terminal year stock size by length, model estimated parameter values, a fishing mortality value ( F ), and adding a constant number of annual recruits. Once the stock dynamics were stabilized (we used the $99^{\text {th }}$ year
estimates) for an F , we calculated the $\mathrm{MMB} / \mathrm{R}$ for that F . We computed the relative $M M B / R$ in percentage, $\left(\frac{M M B}{R}\right)_{x \%}\left(\right.$ where $\mathrm{x} \%=\frac{\frac{M M B_{F}}{R}}{\frac{M M B_{0}}{R}} \times 100$ and $M M B_{0} / R$ is the virgin $M M B / R$ ) for different F values.
$F_{35 \%}$ is the F value that produces the $\mathrm{MMB} / \mathrm{R}$ value equal to $35 \%$ of $M M B_{0} / R$.
$M M B_{35 \%}$ is estimated using the following formula:
$M M B_{35 \%}=\left(\frac{M M B}{R}\right)_{35} \times \bar{R}$, where $\bar{R}$ is the mean number of estimated recruits for a selected period.

Specification of the OFL:
We determined $F_{O F L}$ using the following equation with an iterative procedure accounting for intervening total crab catch removals. The formula for removal of catches and groundfish discards are given in Siddeek et al. (2021).

If,
$M M B_{\text {current }}>M M B_{35 \%}, F_{\text {OFL }}=F_{35 \%}$
If,
$M M B_{\text {current }} \leq M M B_{35 \%}$ and $M M B_{\text {current }}>0.25 M M B_{35 \%}$,
$F_{O F L}=F_{35 \%} \frac{\left(\frac{M M B_{\text {current }}}{M M B_{35 \%}}-\alpha\right)}{(1-\alpha)}$
If,
$M M B_{\text {current }} \leq 0.25 M M B_{35 \%}$,
$F_{O F L}=0$.
where $\alpha$ is a parameter, $\mathrm{MMB}_{\text {current }}$ is the mature male biomass in the current year, and $M M B_{35 \%}$ is the proxy $M M B_{M S Y}$ for Tier 3 stocks. We set $\alpha$ at 0.1 .

## Calculation of ABC:

The cumulative probability distribution of OFL, assuming a log normal distribution of OFL, was used to estimate OFL at the 0.5 probability and the ABC using $25 \%$ and $30 \%$ buffers on estimated OFL.

The OFL and ABC estimates for various models under Tier 3 are summarized separately for EAG, WAG, and the entire Aleutian Islands (AI) in Table 16.

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Table 1. Commercial fishery history for the Aleutian Islands golden king crab fishery 1981/82-2020/21: number of vessels, guideline harvest level (GHL; established in lb, converted to t) for 1996/97-2004/05, total allowable catch (TAC; established in lb, converted to $t$ ) for 2005/06-2020/21, weight of retained catch (harvest; $t$ ), number of retained crab, pot lifts, fishery catch-per-uniteffort (CPUE; retained crab per pot lift), and average weight ( kg ) of landed crab. The values are separated by EAG and WAG beginning in 1996/97.

| Crab <br> Fishing <br> Season | Vessels | GHL/TAC | Harvest ${ }^{\text {a }}$ | Crab | Pot Lifts | CPUE ${ }^{\text {b }}$ | Average Weight ${ }^{\text {c }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1981/82 | 14-20 | - | 599 | 240,458 | 27,533 | 9 | $2.5{ }^{\text {d }}$ |
| 1982/83 | 99-148 | - | 4,169 | 1,737,109 | 179,472 | 10 | $2.4{ }^{\text {d }}$ |
| 1983/84 | 157-204 | - | 4,508 | 1,773,262 | 256,393 | 7 | $2.5{ }^{\text {d }}$ |
| 1984/85 | 38-51 | - | 2,132 | 971,274 | 88,821 | 11 | $2.2{ }^{\text {e }}$ |
| 1985/86 | 53 | - | 5,776 | 2,816,313 | 236,601 | 12 | $2.1{ }^{\text {f }}$ |
| 1986/87 | 64 | - | 6,685 | 3,345,680 | 433,870 | 8 | $2.0{ }^{\text {f }}$ |
| 1987/88 | 66 | - | 4,199 | 2,177,229 | 307,130 | 7 | $1.9{ }^{\text {f }}$ |
| 1988/89 | 76 | - | 4,820 | 2,488,433 | 321,927 | 8 | $1.9{ }^{\text {f }}$ |
| 1989/90 | 68 | - | 5,453 | 2,902,913 | 357,803 | 8 | $1.9{ }^{\text {f }}$ |
| 1990/91 | 24 | - | 3,153 | 1,707,618 | 215,840 | 8 | $1.9{ }^{\text {f }}$ |
| 1991/92 | 20 | - | 3,494 | 1,847,398 | 234,857 | 8 | $1.9{ }^{\text {f }}$ |
| 1992/93 | 22 | - | 2,854 | 1,528,328 | 203,221 | 8 | $1.9{ }^{\text {f }}$ |
| 1993/94 | 21 | - | 2,518 | 1,397,530 | 234,654 | 6 | $1.8{ }^{\text {f }}$ |
| 1994/95 | 35 | - | 3,687 | 1,924,271 | 386,593 | 5 | $1.9{ }^{\text {f }}$ |


| Crab <br> Fishing <br> Season | Vessels |  | GHL/TAC |  | Harvest ${ }^{\text {a }}$ |  | Crab |  | Pot Lifts |  | CPUE ${ }^{\text {b }}$ |  | Average Weight ${ }^{\text {c }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995/96 | 28 |  | - |  | 3,157 |  | 1,582,333 |  | 293,021 |  | 5 |  | $2.0{ }^{\text {f }}$ |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1996/97 | 14 | 13 | 1,452 | 1,225 | 1,493 | 1,145 | 731,909 | 602,968 | 113,460 | 99,267 | 7 | 6 | $2.04{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 1997/98 | 13 | 9 | 1,452 | 1,225 | 1,588 | 1,109 | 780,610 | 569,550 | 106,403 | 86,811 | 7 | 7 | $2.04{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 1998/99 | 14 | 3 | 1,361 | 1,225 | 1,473 | 768 | 740,011 | 410,018 | 83,378 | 35,975 | 9 | 11 | $2.00^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 1999/00 | 15 | 15 | 1,361 | 1,225 | 1,392 | 1,256 | 709,332 | 676,558 | 79,129 | 107,040 | 9 | 6 | $1.95{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2000/01 | 15 | 12 | 1,361 | 1,225 | 1,422 | 1,308 | 704,702 | 705,613 | 71,551 | 101,239 | 10 | 7 | $2.00^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2001/02 | 19 | 9 | 1,361 | 1,225 | 1,442 | 1,243 | 730,030 | 686,738 | 62,639 | 105,512 | 12 | 7 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2002/03 | 19 | 6 | 1,361 | 1,225 | 1,280 | 1,198 | 643,886 | 664,823 | 52,042 | 78,979 | 12 | 8 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2003/04 | 18 | 6 | 1,361 | 1,225 | 1,350 | 1,220 | 643,074 | 676,633 | 58,883 | 66,236 | 11 | 10 | $2.09^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2004/05 | 19 | 6 | 1,361 | 1,225 | 1,309 | 1,219 | 637,536 | 685,465 | 34,848 | 56,846 | 18 | 12 | $2.04{ }^{\text {f }}$ | $1.77{ }^{\text {f }}$ |
| 2005/06 | 7 | 3 | 1,361 | 1,225 | 1,300 | 1,204 | 623,971 | 639,368 | 24,569 | 30,116 | 25 | 21 | $2.09^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2006/07 | 6 | 4 | 1,361 | 1,225 | 1,357 | 1,030 | 650,587 | 527,734 | 26,195 | 26,870 | 25 | 20 | $2.09^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2007/08 | 4 | 3 | 1,361 | 1,225 | 1,356 | 1,142 | 633,253 | 600,595 | 22,653 | 29,950 | 28 | 20 | $2.13{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2008/09 | 3 | 3 | 1,361 | 1,286 | 1,426 | 1,150 | 666,946 | 587,661 | 24,466 | 26,200 | 27 | 22 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2009/10 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,253 | 679,886 | 628,332 | 29,298 | 26,489 | 26 | 24 | $2.09^{\text {f }}$ | $2.00^{\text {f }}$ |
| 2010/11 | 3 | 3 | 1,429 | 1,286 | 1,428 | 1,279 | 670,983 | 626,246 | 25,851 | 29,994 | 26 | 21 | $2.13{ }^{\text {f }}$ | $2.04{ }^{\text {f }}$ |


| Crab | Vessels | GHL/TAC | Harvest $^{\text {a }}$ | Crab | Pot Lifts | CPUE $^{\text {b }}$ | Average <br> Weight ${ }^{\text {c }}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Fishing |  |  |  |  |  |  |  |


|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2011 / 12$ | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,276 | 668,828 | 616,118 | 17,915 | 26,326 | 37 | 23 | $2.13^{f}$ | $2.09^{f}$ |
| $2012 / 13$ | 3 | 3 | 1,501 | 1,352 | 1,504 | 1,339 | 687,666 | 672,916 | 20,827 | 32,716 | 33 | 21 | $2.18^{f}$ | $2.00^{\mathrm{f}}$ |
| $2013 / 14$ | 3 | 3 | 1,501 | 1,352 | 1,546 | 1,347 | 720,220 | 686,883 | 21,388 | 41,835 | 34 | 16 | $2.13^{\mathrm{f}}$ | $1.95^{\mathrm{f}}$ |
| $2014 / 15$ | 3 | 2 | 1,501 | 1,352 | 1,554 | 1,217 | 719,064 | 635,312 | 17,002 | 41,548 | 42 | 15 | $2.18^{\mathrm{f}}$ | $1.91^{\mathrm{f}}$ |
| $2015 / 16$ | 3 | 2 | 1,501 | 1,352 | 1,590 | 1,139 | 763,604 | 615,355 | 19,376 | 41,108 | 39 | 15 | $2.09^{\mathrm{f}}$ | $1.85^{\mathrm{f}}$ |
| $2016 / 17$ | 3 | 3 | 1,501 | 1,014 | 1,578 | 1,015 | 793,983 | 543,796 | 24,470 | 38,118 | 32 | 14 | $1.99^{\mathrm{f}}$ | $1.87^{\mathrm{f}}$ |
| $2017 / 18$ | 3 | 3 | 1,501 | 1,014 | 1,571 | 1,014 | 802,610 | 519,051 | 25,516 | 30,885 | 31 | 17 | $1.96^{\mathrm{f}}$ | $1.95^{\mathrm{f}}$ |
| $2018 / 19$ | 3 | 3 | 1,749 | 1,134 | 1,830 | 1,135 | 940,336 | 578,221 | 25,553 | 29,156 | 37 | 20 | $1.95^{\mathrm{f}}$ | $1.96^{\mathrm{f}}$ |
| $2019 / 20$ | 3 | 3 | 1,955 | 1,302 | 2,031 | 1,288 | $1,057,464$ | 649,832 | 30,998 | 42,924 | 34 | 15 | $1.92^{\mathrm{f}}$ | $1.98^{\mathrm{f}}$ |
| $2020 / 21$ | 3 | 3 | 1,656 | 1,343 | 1,733 | 1,267 | 902,122 | 682,107 | 30,072 | 46,701 | 30 | 15 | $1.92^{\mathrm{f}}$ | $1.86^{\mathrm{f}}$ |

## Note:

a. Includes deadloss.
b. Number of crab per pot lift.
c. Average weight of landed crab, including dead loss.
d. Managed with $6.5^{\prime \prime}$ carapace width (CW) minimum size limit.
e. Managed with $6.5^{\prime \prime} \mathrm{CW}$ minimum size limit west of $171^{\circ} \mathrm{W}$ longitude and $6.0^{\prime \prime}$ minimum size limit east of $171^{\circ} \mathrm{W}$ longitude.
f. Managed with 6.0" minimum size limit.

Catch and effort data include cost recovery fishery.

Table 2. Annual weight of total fishery mortality to Aleutian Islands golden king crab, 1981/82 2020/21, partitioned by source of mortality: retained catch, bycatch mortality during crab fisheries, and bycatch mortality during groundfish fisheries. For bycatch in the federal groundfish fisheries, historical data (1991-2008) are not available for areas east and west of 174 W , and are listed for federal groundfish reporting areas 541,542 , and 543 combined. The 2009- present data are available by separate EAG and WAG fisheries and are listed as such. A mortality rate of $20 \%$ was applied for crab fisheries bycatch, and a mortality rate of $50 \%$ for groundfish pot fisheries and $80 \%$ for the trawl fisheries were applied.

| Season | Retained <br> Catch ( t ) |  | Bycatch Mortality by Fishery Type (t) |  |  |  | Total Fishery Mortality (t) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Crab |  | Groundfish |  |  |  |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | Entire AI |
| 1981/82 | 490 | 95 |  |  |  |  |  |  | 585 |
| 1982/83 | 1,260 | 2,655 |  |  |  |  |  |  | 3,914 |
| 1983/84 | 1,554 | 2,991 |  |  |  |  |  |  | 4,545 |
| 1984/85 | 1,839 | 424 |  |  |  |  |  |  | 2,263 |
| 1985/86 | 2,677 | 1,996 |  |  |  |  |  |  | 4,673 |
| 1986/87 | 2,798 | 4,200 |  |  |  |  |  |  | 6,998 |
| 1987/88 | 1,882 | 2,496 |  |  |  |  |  |  | 4,379 |
| 1988/89 | 2,382 | 2,441 |  |  |  |  |  |  | 4,823 |
| 1989/90 | 2,738 | 3,028 |  |  |  |  |  |  | 5,766 |
| 1990/91 | 1,623 | 1,621 |  |  |  |  |  |  | 3,244 |
| 1991/92 | 2,035 | 1,397 | 515 | 344 |  | 0 |  |  | 4,291 |
| 1992/93 | 2,112 | 1,025 | 1,206 | 373 |  | 0 |  |  | 4,716 |
| 1993/94 | 1,439 | 686 | 383 | 258 |  | 4 |  |  | 2,770 |
| 1994/95 | 2,044 | 1,540 | 687 | 823 |  | 1 |  |  | 5,095 |
| 1995/96 | 2,259 | 1,203 | 725 | 530 |  | 2 |  |  | 4,719 |
| 1996/97 | 1,738 | 1,259 | 485 | 439 |  | 5 |  |  | 3,926 |
| 1997/98 | 1,588 | 1,083 | 441 | 343 |  | 1 |  |  | 3,455 |
| 1998/99 | 1,473 | 955 | 434 | 285 |  | 1 |  |  | 3,149 |
| 1999/00 | 1,392 | 1,222 | 313 | 385 |  | 3 |  |  | 3,316 |
| 2000/01 | 1,422 | 1,342 | 82 | 437 |  | 2 |  |  | 3,285 |
| 2001/02 | 1,442 | 1,243 | 74 | 387 |  | 0 |  |  | 3,146 |
| 2002/03 | 1,280 | 1,198 | 52 | 303 |  | 8 |  |  | 2,850 |
| 2003/04 | 1,350 | 1,220 | 53 | 148 |  | 0 |  |  | 2,792 |
| 2004/05 | 1,309 | 1,219 | 41 | 143 |  | 1 |  |  | 2,715 |
| 2005/06 | 1,300 | 1,204 | 22 | 73 |  | 2 |  |  | 2,601 |
| 2006/07 | 1,357 | 1,022 | 28 | 81 |  | 8 |  |  | 2,506 |
| 2007/08 | 1,356 | 1,142 | 24 | 114 |  | 9 |  |  | 2,695 |
| 2008/09 | 1,426 | 1,150 | 61 | 102 |  | 3 |  |  | 2,772 |
| 2009/10 | 1,429 | 1,253 | 111 | 108 | 18 | 5 | 1,558 | 1,366 | 2,923 |
| 2010/11 | 1,428 | 1,279 | 123 | 124 | 49 | 3 | 1,600 | 1,407 | 3,006 |
| 2011/12 | 1,429 | 1,276 | 106 | 117 | 25 | 4 | 1,560 | 1,398 | 2,957 |
| 2012/13 | 1,504 | 1,339 | 118 | 145 | 9 | 6 | 1,631 | 1,491 | 3,122 |


| $2013 / 14$ | 1,546 | 1,347 | 113 | 174 | 5 | 7 | 1,665 | 1,528 | 3,192 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2014 / 15$ | 1,554 | 1,217 | 127 | 175 | 9 | 5 | 1,691 | 1,397 | 3,088 |
| $2015 / 16$ | 1,590 | 1,139 | 165 | 157 | 23 | 2 | 1,778 | 1,298 | 3,076 |
| $2016 / 17$ | 1,578 | 1,015 | 203 | 145 | 101 | 4 | 1,882 | 1,164 | 3,046 |
| $2017 / 18$ | 1,571 | 1,014 | 219 | 126 | 47 | 2 | 1,837 | 1,142 | 2,979 |
| $2018 / 19$ | 1,830 | 1,135 | 240 | 140 | 24 | 3 | 2,094 | 1,278 | 3,372 |
| $2019 / 20$ | 2,031 | 1,288 | 275 | 112 | 18 | 6 | 2,327 | 1,406 | 3,733 |
| $2020 / 21$ | 1,733 | 1,267 | 241 | 147 | 40 | 17 | 2,014 | 1,431 | 3,444 |

Table 3. Time series of nominal annual pot fishery retained, observer retained, and observer total catch-per-unit-effort (CPUE, number of crabs per pot lift), total pot fishing effort (number of pot lifts), observer sample size (number of sampled pots), and GLM estimated observer CPUE Index (for non-interaction model 21.1a) for the EAG and WAG golden king crab stocks, 1985/862020/21. Observer retained CPUE includes retained and non-retained legal-size crabs.

| Year | Pot Fishery Nominal Retained CPUE |  | Obs. Nominal Retained CPUE |  | Obs. Nominal Total CPUE |  | Pot Fishery Effort (no.pot lifts) |  | Obs. Sample Size (no.pot lifts) |  | Obs. CPUE Index |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1985/86 | 11.90 | 11.90 |  |  |  |  | 117,718 | 118,563 |  |  |  |  |
| 1986/87 | 8.42 | 7.32 |  |  |  |  | 155,240 | 277,780 |  |  |  |  |
| 1987/88 | 7.03 | 7.15 |  |  |  |  | 146,501 | 160,229 |  |  |  |  |
| 1988/89 | 7.52 | 7.93 |  |  |  |  | 155,518 | 166,409 |  |  |  |  |
| 1989/90 | 8.49 | 7.83 |  |  |  |  | 155,262 | 202,541 |  |  |  |  |
| 1990/91 | 8.90 | 7.00 | 6.84 | 8.34 | 13.00 | 26.67 | 106,281 | 108,533 | 138 | 340 |  |  |
| 1991/92 | 8.20 | 7.40 | 9.84 | 6.14 | 36.91 | 19.17 | 133,428 | 101,429 | 377 | 857 |  |  |
| 1992/93 | 8.40 | 5.90 | 10.44 | 4.26 | 38.52 | 16.83 | 133,778 | 69,443 | 199 | 690 |  |  |
| 1993/94 | 7.80 | 4.40 | 5.91 | 12.75 | 20.81 | 17.23 | 106,890 | 127,764 | 31 | 174 |  |  |
| 1994/95 | 5.90 | 4.10 | 4.66 | 6.62 | 12.91 | 19.23 | 191,455 | 195,138 | 127 | 1,270 |  |  |
| 1995/96 | 5.90 | 4.70 | 6.03 | 6.03 | 16.98 | 14.28 | 177,773 | 115,248 | 6,388 | 5,598 | 0.71 | 1.00 |
| 1996/97 | 6.50 | 6.10 | 6.02 | 5.90 | 13.81 | 13.54 | 113,460 | 99,267 | 8,360 | 7,194 | 0.72 | 0.93 |
| 1997/98 | 7.30 | 6.60 | 7.99 | 6.72 | 18.25 | 15.03 | 106,403 | 86,811 | 4,670 | 3,985 | 0.80 | 0.99 |
| 1998/99 | 8.90 | 11.40 | 9.82 | 9.43 | 25.77 | 23.09 | 83,378 | 35,975 | 3,616 | 1,876 | 0.95 | 1.07 |
| 1999/00 | 9.00 | 6.30 | 10.28 | 6.09 | 20.77 | 14.49 | 79,129 | 107,040 | 3,851 | 4,523 | 0.93 | 0.92 |
| 2000/01 | 9.90 | 7.00 | 10.40 | 6.46 | 25.39 | 16.64 | 71,551 | 101,239 | 5,043 | 4,740 | 0.88 | 0.80 |
| 2001/02 | 11.70 | 6.50 | 11.73 | 6.04 | 22.48 | 14.66 | 62,639 | 105,512 | 4,626 | 4,454 | 1.18 | 0.86 |
| 2002/03 | 12.40 | 8.40 | 12.70 | 7.47 | 22.59 | 17.37 | 52,042 | 78,979 | 3,980 | 2,509 | 1.33 | 0.97 |
| 2003/04 | 10.90 | 10.20 | 11.34 | 9.33 | 19.43 | 18.17 | 58,883 | 66,236 | 3,960 | 3,334 | 1.16 | 1.28 |
| 2004/05 | 18.30 | 12.10 | 18.34 | 11.14 | 28.48 | 22.45 | 34,848 | 56,846 | 2,206 | 2,619 | 1.74 | 1.30 |
| 2005/06 | 25.40 | 21.20 | 29.52 | 23.89 | 38.55 | 36.23 | 24,569 | 30,116 | 1,193 | 1,365 | 0.97 | 1.18 |
| 2006/07 | 24.80 | 19.60 | 25.13 | 23.93 | 33.39 | 33.47 | 26,195 | 26,870 | 1,098 | 1,183 | 0.81 | 1.15 |
| 2007/08 | 28.00 | 20.00 | 31.10 | 21.01 | 40.38 | 32.46 | 22,653 | 29,950 | 998 | 1,082 | 0.90 | 1.00 |
| 2008/09 | 27.30 | 22.40 | 29.97 | 24.50 | 38.23 | 38.16 | 24,466 | 26,200 | 613 | 979 | 0.88 | 1.17 |
| 2009/10 | 25.90 | 23.70 | 26.60 | 26.54 | 35.88 | 34.08 | 26,298 | 26,489 | 408 | 892 | 0.72 | 1.24 |
| 2010/11 | 26.00 | 20.90 | 26.40 | 22.43 | 37.10 | 29.05 | 25,851 | 29,994 | 436 | 867 | 0.75 | 1.08 |
| 2011/12 | 37.30 | 23.40 | 39.48 | 23.63 | 52.04 | 31.13 | 17,915 | 26,326 | 361 | 837 | 1.08 | 1.11 |
| 2012/13 | 33.02 | 20.57 | 37.82 | 22.88 | 47.57 | 30.76 | 20,827 | 32,716 | 438 | 1,109 | 1.04 | 1.09 |
| 2013/14 | 33.67 | 16.42 | 35.94 | 16.89 | 46.16 | 25.01 | 21,388 | 41,835 | 499 | 1,223 | 1.01 | 0.82 |
| 2014/15 | 42.29 | 15.29 | 47.01 | 15.25 | 60.00 | 22.67 | 17,002 | 41,548 | 376 | 1,137 | 1.33 | 0.73 |
| 2015/16 | 39.41 | 14.97 | 43.27 | 15.81 | 58.68 | 22.14 | 19,376 | 41,108 | 478 | 1,296 | 1.26 | 0.75 |
| 2016/17 | 32.45 | 14.29 | 36.89 | 16.65 | 52.82 | 24.41 | 24,470 | 38,118 | 617 | 1,060 | 1.06 | 0.86 |
| 2017/18 | 31.46 | 16.81 | 35.18 | 19.30 | 54.62 | 25.54 | 25,516 | 30,885 | 585 | 760 | 1.01 | 0.99 |
| 2018/19 | 36.80 | 19.83 | 41.57 | 22.90 | 62.97 | 30.69 | 25,553 | 29,156 | 475 | 688 | 1.23 | 1.21 |
| 2019/20 | 34.11 | 15.10 | 40.88 | 16.30 | 57.46 | 22.73 | 30,998 | 42,963 | 540 | 967 | 1.15 | 0.98 |
| 2020/21 | 30.00 | 14.61 | 36.15 | 15.71 | 57.21 | 22.82 | 30,072 | 46,701 | 567 | 1,137 | 1.05 | 0.86 |

Table 4. Time series of GLM estimated CPUE indices and standard errors [standard error of $\ln$ (CPUE index)] for fish ticket based retained catch-per-pot lift (CPUE) for the EAG and WAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data.

|  | EAG <br> Negative <br> Binomial | Standard <br> Error of <br> ln <br> Year | WAG Negative <br> Binomial | Standard <br> Error of ln <br> (CPUE) |
| :--- | :---: | :---: | :---: | :---: |
|  | Index | CPUE Index | (CPUE) |  |
| $1985 / 86$ | 1.58 | 0.19 | 1.37 |  |
| $1986 / 87$ | 0.58 | 0.57 | 1.56 | 0.10 |
| $1987 / 88$ | 0.79 | 0.50 | 1.05 | 0.08 |
| $1988 / 89$ | 1.60 | 0.16 | 1.49 | 0.08 |
| $1989 / 90$ | 0.78 | 0.14 | 1.15 | 0.04 |
| $1990 / 91$ | 1.15 | 0.15 | 0.90 | 0.03 |
| $1991 / 92$ | 1.08 | 0.12 | 0.81 | 0.04 |
| $1992 / 93$ | 0.79 | 0.15 | 0.65 | 0.04 |
| $1993 / 94$ | 1.28 | 0.13 | 0.77 | 0.04 |
| $1994 / 95$ | 0.94 | 0.11 | 0.86 | 0.06 |
| $1995 / 96$ | 0.54 | 0.16 | 0.97 | 0.04 |
| $1996 / 97$ | 0.89 | 0.11 | 0.90 | 0.04 |
| $1997 / 98$ | 1.47 | 0.11 | 0.84 | 0.03 |
| $1998 / 99$ | 1.30 | 0.10 | 1.16 | 0.03 |

Table 5. Parameter estimates and coefficient of variations (CV) with the 2020 MMB (MMB estimated on 15 Feb 2021) for models 21.1a, 21.1b, and 21.1c for the golden king crab data from the EAG, 1985/86-2020/21. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

|  | Model 21.1a |  | Model 21.1b |  | Model 21.1c |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | CV | Estimate | CV | Limits |
| $\log _{-} \omega_{1}$ (growth incr. intercept) | 2.53 | 0.01 | 2.53 | 0.01 | 2.53 | 0.01 | 1.0, 4.5 |
| $\omega_{2}$ (growth incr. slope) | -9.78 | 0.18 | -9.77 | 0.18 | -9.77 | 0.18 | -12.0,-5.0 |
| $\log _{\text {_ }} \mathrm{a}$ (molt prob. slope) | -2.56 | 0.02 | -2.56 | 0.02 | -2.57 | 0.02 | -4.61,-1.39 |
| $\log _{\text {_ }} \mathrm{b}$ (molt prob. L50) | 4.94 | 0.001 | 4.94 | 0.001 | 4.94 | 0.001 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.68 | 0.03 | 3.68 | 0.03 | 3.68 | 0.03 | 0.1,12.0 |
| log_total sel delta $\theta$, 1985-04 | 3.44 | 0.02 | 3.44 | 0.02 | 3.44 | 0.02 | 0.,4.4 |
| $\log _{\text {_ }}$ total sel delta $\theta, 2005-19$ | 2.97 | 0.02 | 2.98 | 0.02 | 2.99 | 0.02 | 0.,4.4 |
| log_ret. sel delta $\theta$, 1985-19 | 1.86 | 0.02 | 1.86 | 0.02 | 1.86 | 0.02 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.88 | 0.002 | 4.88 | 0.002 | 4.88 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-19$ | 4.92 | 0.002 | 4.92 | 0.002 | 4.92 | 0.002 | 4.0,5.0 |
| log_ret. sel $\theta_{50}, 1985-19$ | 4.92 | 0.0003 | 4.92 | 0.0003 | 4.92 | 0.0003 | 4.0,5.0 |
| $\log \_\beta_{\mathrm{r}}$ (rec.distribution par.) | -1.01 | 0.14 | -1.02 | 0.14 | -1.02 | 0.14 | -12.0, 12.0 |
| Logq1 (fishery catchability 1985-04) |  |  |  |  | -0.63 | 16132.31 | -9.0, 2.25 |
| $\operatorname{logq} 2$ (fishery catchability 1985-04 / observer catchability |  |  |  |  |  |  |  |
| 1995-04) | -0.49 | 0.19 | -0.50 | 0.17 | -0.52 | 0.16 | -9.0, 2.25 |
| logq3 (observer catchability 2005-20) | -0.83 | 0.14 | -0.72 | 0.16 | -0.56 | 0.17 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 1.04 | 0.04 | 1.02 | 0.04 | 1.01 | 0.04 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -0.99 | 0.07 | -0.97 | 0.07 | -0.95 | 0.07 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -9.32 | 0.08 | -9.30 | 0.08 | -9.28 | 0.08 | -15.0, -1.6 |
| $\sigma_{e}^{2} \quad$ (Fishery CPUE additional log standard deviation, 19851998) | -1.09 | 0.25 | -1.07 | 0.25 | -1.46 | 0.17 | -8.0, 1.0 |

$\sigma_{e}^{2} \quad$ (observer CPUE additional log standard deviation, 19952004)
$\sigma_{e}^{2} \quad$ (observer CPUE additional log standard deviation, 20052020)

2020 MMB

|  |  |  |  | -2.09 | 0.12 | $-8.0,0.15$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |
| -1.61 | 0.11 | -3.39 | 0.91 | -1.10 | 0.23 | $-8.0,0.15$ |
| 12,561 | 0.20 | 10,820 | 0.20 | 8,280 | 0.26 |  |

Table 6. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1a for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to 1985-2020. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq 101$ mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=25,859 \\ & M M B_{35 \%}=9,250 \end{aligned}$ |  |  |  |
| 1985 | 2.60 | 10,380 | 0.03 | 10,335 | 0.05 |
| 1986 | 1.52 | 7,932 | 0.03 | 8,707 | 0.04 |
| 1987 | 4.39 | 7,446 | 0.04 | 6,769 | 0.04 |
| 1988 | 9.97 | 8,536 | 0.04 | 5,739 | 0.04 |
| 1989 | 2.41 | 6,976 | 0.05 | 5,038 | 0.05 |
| 1990 | 4.42 | 6,974 | 0.04 | 4,737 | 0.06 |
| 1991 | 5.77 | 7,214 | 0.04 | 5,024 | 0.06 |
| 1992 | 3.75 | 6,857 | 0.04 | 4,811 | 0.05 |
| 1993 | 3.16 | 7,086 | 0.03 | 4,760 | 0.05 |
| 1994 | 3.87 | 6,570 | 0.03 | 5,300 | 0.03 |
| 1995 | 3.50 | 5,760 | 0.04 | 4,807 | 0.03 |
| 1996 | 3.11 | 5,802 | 0.04 | 4,079 | 0.04 |
| 1997 | 4.24 | 6,112 | 0.04 | 4,189 | 0.04 |
| 1998 | 3.88 | 6,672 | 0.05 | 4,256 | 0.05 |
| 1999 | 3.78 | 7,386 | 0.05 | 4,699 | 0.05 |
| 2000 | 3.63 | 8,027 | 0.05 | 5,404 | 0.05 |
| 2001 | 2.62 | 8,351 | 0.06 | 6,071 | 0.06 |
| 2002 | 3.14 | 8,746 | 0.06 | 6,671 | 0.06 |
| 2003 | 2.70 | 9,028 | 0.07 | 7,093 | 0.07 |
| 2004 | 2.16 | 9,018 | 0.07 | 7,418 | 0.07 |
| 2005 | 2.89 | 9,092 | 0.07 | 7,617 | 0.07 |
| 2006 | 2.46 | 9,169 | 0.07 | 7,549 | 0.08 |
| 2007 | 2.32 | 9,119 | 0.07 | 7,596 | 0.08 |
| 2008 | 3.27 | 9,267 | 0.07 | 7,644 | 0.08 |
| 2009 | 2.50 | 9,437 | 0.06 | 7,562 | 0.08 |
| 2010 | 2.03 | 9,265 | 0.06 | 7,754 | 0.07 |
| 2011 | 2.42 | 9,004 | 0.06 | 7,845 | 0.07 |
| 2012 | 2.26 | 8,718 | 0.06 | 7,613 | 0.07 |
| 2013 | 1.93 | 8,237 | 0.06 | 7,319 | 0.06 |
| 2014 | 2.94 | 8,016 | 0.07 | 6,955 | 0.07 |
| 2015 | 3.55 | 8,216 | 0.07 | 6,520 | 0.07 |
| 2016 | 3.54 | 8,637 | 0.09 | 6,345 | 0.08 |
| 2017 | 4.30 | 9,433 | 0.10 | 6,590 | 0.09 |
| 2018 | 5.51 | 10,661 | 0.13 | 7,126 | 0.10 |
| 2019 | 4.44 | 11,716 | 0.17 | 7,830 | 0.13 |
| 2020 | 3.22 | 12,561 | 0.20 | 8,891 | 0.17 |

Table 7. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1 b for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to 1985-2020. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the <br> Model ( $\geq 101$ mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=25,511 \\ & M M B_{35 \%}=9,177 \end{aligned}$ |  |  |  |
| 1985 | 2.60 | 10,384 | 0.03 | 10,315 | 0.05 |
| 1986 | 1.50 | 7,932 | 0.03 | 8,704 | 0.04 |
| 1987 | 4.41 | 7,447 | 0.04 | 6,772 | 0.04 |
| 1988 | 10.00 | 8,549 | 0.04 | 5,738 | 0.04 |
| 1989 | 2.38 | 6,986 | 0.05 | 5,039 | 0.05 |
| 1990 | 4.42 | 6,978 | 0.04 | 4,744 | 0.06 |
| 1991 | 5.78 | 7,220 | 0.04 | 5,028 | 0.06 |
| 1992 | 3.75 | 6,865 | 0.04 | 4,812 | 0.05 |
| 1993 | 3.15 | 7,091 | 0.03 | 4,764 | 0.05 |
| 1994 | 3.87 | 6,573 | 0.03 | 5,304 | 0.03 |
| 1995 | 3.52 | 5,769 | 0.04 | 4,809 | 0.03 |
| 1996 | 3.12 | 5,820 | 0.04 | 4,081 | 0.04 |
| 1997 | 4.23 | 6,129 | 0.04 | 4,199 | 0.04 |
| 1998 | 3.85 | 6,675 | 0.05 | 4,272 | 0.05 |
| 1999 | 3.76 | 7,373 | 0.05 | 4,709 | 0.05 |
| 2000 | 3.57 | 7,984 | 0.05 | 5,399 | 0.05 |
| 2001 | 2.60 | 8,276 | 0.06 | 6,047 | 0.06 |
| 2002 | 3.20 | 8,679 | 0.06 | 6,615 | 0.06 |
| 2003 | 2.71 | 8,984 | 0.06 | 7,018 | 0.06 |
| 2004 | 2.18 | 8,991 | 0.07 | 7,361 | 0.07 |
| 2005 | 2.90 | 9,084 | 0.06 | 7,580 | 0.07 |
| 2006 | 2.50 | 9,184 | 0.06 | 7,532 | 0.07 |
| 2007 | 2.34 | 9,160 | 0.06 | 7,596 | 0.07 |
| 2008 | 3.18 | 9,289 | 0.06 | 7,668 | 0.07 |
| 2009 | 2.36 | 9,380 | 0.06 | 7,604 | 0.07 |
| 2010 | 1.99 | 9,140 | 0.06 | 7,756 | 0.07 |
| 2011 | 2.37 | 8,840 | 0.06 | 7,766 | 0.06 |
| 2012 | 2.16 | 8,502 | 0.05 | 7,480 | 0.06 |
| 2013 | 1.86 | 7,967 | 0.06 | 7,147 | 0.06 |
| 2014 | 2.88 | 7,708 | 0.06 | 6,732 | 0.06 |
| 2015 | 3.32 | 7,816 | 0.06 | 6,253 | 0.06 |
| 2016 | 3.34 | 8,099 | 0.07 | 6,034 | 0.07 |
| 2017 | 3.95 | 8,701 | 0.09 | 6,170 | 0.08 |
| 2018 | 4.81 | 9,555 | 0.12 | 6,563 | 0.09 |
| 2019 | 3.98 | 10,191 | 0.17 | 7,032 | 0.12 |
| 2020 | 3.16 | 10,820 | 0.20 | 7,692 | 0.17 |

Table 8. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1c for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to 1985-2020. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the <br> Model ( $\geq 101$ mm CL) | $\begin{gathered} \hline \text { Mature Male } \\ \text { Biomass } \\ \text { (Bent-Point fit) } \end{gathered}$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=25,224 \\ & M M B_{35 \%}=9,056 \end{aligned}$ |  |  |  |
| 1985 | 2.59 | 10,418 | 0.03 | 10,354 | 0.05 |
| 1986 | 1.49 | 7,957 | 0.03 | 8,745 | 0.04 |
| 1987 | 4.47 | 7,484 | 0.04 | 6,804 | 0.04 |
| 1988 | 9.95 | 8,580 | 0.04 | 5,763 | 0.04 |
| 1989 | 2.36 | 7,019 | 0.05 | 5,075 | 0.05 |
| 1990 | 4.38 | 6,999 | 0.04 | 4,785 | 0.06 |
| 1991 | 5.75 | 7,230 | 0.04 | 5,064 | 0.06 |
| 1992 | 3.75 | 6,874 | 0.04 | 4,836 | 0.05 |
| 1993 | 3.15 | 7,098 | 0.03 | 4,779 | 0.05 |
| 1994 | 3.87 | 6,580 | 0.03 | 5,315 | 0.03 |
| 1995 | 3.52 | 5,776 | 0.04 | 4,817 | 0.03 |
| 1996 | 3.12 | 5,825 | 0.04 | 4,090 | 0.04 |
| 1997 | 4.21 | 6,127 | 0.04 | 4,206 | 0.04 |
| 1998 | 3.81 | 6,655 | 0.05 | 4,277 | 0.05 |
| 1999 | 3.77 | 7,341 | 0.05 | 4,704 | 0.05 |
| 2000 | 3.59 | 7,958 | 0.05 | 5,376 | 0.05 |
| 2001 | 2.58 | 8,251 | 0.06 | 6,019 | 0.06 |
| 2002 | 3.14 | 8,631 | 0.06 | 6,595 | 0.06 |
| 2003 | 2.68 | 8,908 | 0.06 | 6,992 | 0.06 |
| 2004 | 2.16 | 8,896 | 0.07 | 7,309 | 0.07 |
| 2005 | 2.87 | 8,976 | 0.06 | 7,503 | 0.07 |
| 2006 | 2.45 | 9,058 | 0.06 | 7,442 | 0.07 |
| 2007 | 2.29 | 9,008 | 0.06 | 7,495 | 0.07 |
| 2008 | 3.15 | 9,117 | 0.06 | 7,548 | 0.07 |
| 2009 | 2.35 | 9,203 | 0.06 | 7,458 | 0.07 |
| 2010 | 1.95 | 8,958 | 0.06 | 7,598 | 0.07 |
| 2011 | 2.29 | 8,634 | 0.05 | 7,605 | 0.06 |
| 2012 | 2.11 | 8,269 | 0.05 | 7,310 | 0.06 |
| 2013 | 1.83 | 7,721 | 0.05 | 6,952 | 0.06 |
| 2014 | 2.78 | 7,429 | 0.06 | 6,516 | 0.06 |
| 2015 | 3.26 | 7,496 | 0.07 | 6,023 | 0.06 |
| 2016 | 3.13 | 7,696 | 0.08 | 5,765 | 0.07 |
| 2017 | 3.71 | 8,156 | 0.11 | 5,856 | 0.08 |
| 2018 | 4.37 | 8,776 | 0.16 | 6,145 | 0.11 |
| 2019 | 3.64 | 9,134 | 0.22 | 6,456 | 0.15 |
| 2020 | 3.09 | 8,280 | 0.26 | 6,863 | 0.22 |

Table 9. Negative log-likelihood values of the fits for models 21.1a (base), 21.1b, 21.1c, 21.1a2, and 21.1b2 for golden king crab in the EAG. Likelihood components with zero entry in the entire rows are omitted.

| Likelihood Component | 21.1a | 21.1b | 21.1c | 21.1a2 | 21.1b2 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Number of free parameters | 152 | 152 | 154 | 152 | 152 |
| Retlencomp | -2090.3300 | -2099.7800 | -2101.0800 | -2090.3300 | -2099.7800 |
| Totallencomp | -1508.7500 | -1506.4100 | -1505.6600 | -1508.7500 | -1506.4100 |
| Observer cpue | -28.7077 | -30.9382 | -25.7316 | -28.7077 | -30.9382 |
| Fishery cpue | -15.1798 | -15.0393 | -13.3399 | -15.1798 | -15.0393 |
| RetdcatchB | 8.4172 | 8.4805 | 8.3674 | 8.4172 | 8.4805 |
| TotalcatchB | 24.4882 | 24.4759 | 24.4362 | 24.4882 | 24.4759 |
| GdiscdcatchB | 0.0003 | 0.0002 | 0.0002 | 0.0003 | 0.0002 |
| Rec_dev | 28.7325 | 28.2168 | 27.4714 | 28.7325 | 28.2168 |
| Pot F_dev | 0.0131 | 0.0130 | 0.0131 | 0.0131 | 0.0130 |
| Gbyc_F_dev | 0.0271 | 0.0274 | 0.0279 | 0.0271 | 0.0274 |
| Tag | 2690.4200 | 2690.4800 | 2690.5800 | 2690.4200 | 2690.4800 |
| RetcatchN | 0.0083 | 0.0079 | 0.0074 | 0.0083 | 0.0079 |
| Total | -890.8550 | -900.4620 | -894.9110 | -890.8550 | -900.4620 |

Table 10. Parameter estimates and coefficient of variations (CV) with the 2020 MMB (MMB estimated on 15 Feb 2021) for models 21.1a, 21.1b, 21.1c, and 21.1d for the golden king crab data from the WAG, 1985/86-2020/21. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

| Parameter | Model 21.1a |  | Model 21.1b |  | Model 21.1c |  | Model 21.1d |  | Limits |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate | CV | Estimate | CV | Estimate | CV | Estimate | CV |  |
| $\log _{-} \omega_{1}$ (growth incr. intercept) | 2.53 | 0.01 | 2.53 | 0.01 | 2.53 | 0.01 | 2.52 | 0.01 | 1.0, 4.5 |
| $\omega_{2}$ (growth incr. slope) | -8.55 | 0.20 | -8.62 | 0.20 | -8.59 | 0.20 | -8.79 | 0.20 | -12.0-5.0 |
| $\log _{\text {a }} \mathrm{a}$ (molt prob. slope) | -2.69 | 0.03 | -2.69 | 0.03 | -2.69 | 0.03 | -2.67 | 0.03 | -4.61,-1.39 |
| log_b (molt prob. L50) | 4.94 | 0.001 | 4.94 | 0.001 | 4.94 | 0.001 | 4.93 | 0.001 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.69 | 0.03 | 3.69 | 0.03 | 3.69 | 0.03 | 3.69 | 0.03 | 0.1,12.0 |
| log_total sel delta $\theta$, 1985-04 | 3.42 | 0.01 | 3.43 | 0.01 | 3.41 | 0.01 | 3.45 | 0.01 | 0.,4.4 |
| $\log _{\text {_ }}$ total sel delta $\theta, 2005-19$ | 2.88 | 0.02 | 2.88 | 0.02 | 2.91 | 0.02 | 2.86 | 0.02 | 0.,4.4 |
| log_ret. sel delta $\theta, 1985-19$ | 1.81 | 0.02 | 1.81 | 0.02 | 1.80 | 0.02 | 1.81 | 0.02 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.91 | 0.002 | 4.92 | 0.002 | 4.91 | 0.002 | 4.92 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-19$ | 4.91 | 0.001 | 4.91 | 0.001 | 4.91 | 0.001 | 4.90 | 0.001 | 4.0,5.0 |
| log_ret. sel $\theta_{50}, 1985-19$ | 4.92 | 0.0002 | 4.92 | 0.000 2 | 4.92 | 0.0002 | 4.92 | $\begin{gathered} 0.000 \\ 2 \end{gathered}$ | 4.0,5.0 |
| $\log _{\Omega} \beta_{\mathrm{r}}$ (rec.distribution par.) | -1.02 | 0.13 | -1.01 | 0.13 | -1.04 | 0.13 | -1.00 | 0.13 | -12.0, 12.0 |
| Logq1 (fishery catchability |  |  |  |  |  |  |  |  |  |
| 1985-04) |  |  |  |  | -3.38 | 3952.89 |  |  | -9.0, 2.25 |
| $\operatorname{logq} 2$ (fishery catchability |  |  |  |  |  |  |  |  |  |
| 1985-04 / observer catchability 1995-04) | 0.11 | 0.68 | 0.11 | 0.80 | -0.13 | 0.56 | 0.25 | 0.39 | -9.0, 2.25 |
| $\operatorname{logq3}$ (observer catchability 2005-20) | -0.43 | 0.21 | -0.42 | 0.24 | -0.24 | 0.32 | -0.50 | 0.18 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.95 | 0.04 | 0.95 | 0.04 | 0.94 | 0.04 | 0.97 | 0.04 | 0.01, 5.0 |
| $\log _{\text {_ }}$ mean_Fpot (Pot fishery F) | -0.74 | 0.08 | -0.73 | 0.08 | -0.72 | 0.08 | -0.77 | 0.08 | -15.0, -0.01 |
| $\log _{-}$mean_Fground (GF byc. F) | -8.56 | 0.09 | -8.55 | 0.09 | -8.53 | 0.09 | -8.58 | 0.09 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (fishery CPUE additional log standard deviation, 19851998) | -1.22 | 0.21 | -1.23 | 0.20 | -2.05 | 0.18 | -0.96 | 0.28 | -8.0, 1.0 |

$\sigma_{e}^{2} \quad$ (observer CPUE additional log standard deviation, 19952004)
$\sigma_{e}^{2} \quad$ (observer CPUE additional log standard deviation, 20052020)

2020 MMB
-2.23
0.08

6,460
0.14
$-2.905$
0.13
$-8.0,0.15$
0.19
,527
0.12

6,85
$-8.0,0.15$

Table 11. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1a for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to 1985-2020. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the <br> Model ( $\geq 101$ <br> mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ <br> mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,519 \\ & M M B_{35 \%}=7,358 \end{aligned}$ |  |  |  |
| 1985 | 4.40 | 12,811 | 0.04 | 10,415 | 0.09 |
| 1986 | 6.19 | 10,601 | 0.04 | 9,760 | 0.06 |
| 1987 | 4.73 | 9,871 | 0.04 | 6,661 | 0.06 |
| 1988 | 3.54 | 8,714 | 0.04 | 6,233 | 0.05 |
| 1989 | 4.34 | 6,408 | 0.04 | 5,651 | 0.04 |
| 1990 | 3.31 | 5,583 | 0.04 | 3,510 | 0.05 |
| 1991 | 2.33 | 4,986 | 0.04 | 3,164 | 0.05 |
| 1992 | 2.64 | 4,903 | 0.04 | 3,027 | 0.05 |
| 1993 | 2.79 | 5,636 | 0.03 | 3,142 | 0.05 |
| 1994 | 2.90 | 5,084 | 0.03 | 3,823 | 0.03 |
| 1995 | 2.92 | 4,938 | 0.03 | 3,071 | 0.03 |
| 1996 | 2.49 | 4,666 | 0.03 | 2,997 | 0.03 |
| 1997 | 2.69 | 4,699 | 0.03 | 3,018 | 0.04 |
| 1998 | 2.65 | 5,078 | 0.03 | 3,073 | 0.04 |
| 1999 | 3.36 | 5,116 | 0.04 | 3,415 | 0.03 |
| 2000 | 3.61 | 5,205 | 0.04 | 3,330 | 0.04 |
| 2001 | 3.62 | 5,643 | 0.04 | 3,314 | 0.04 |
| 2002 | 3.50 | 6,252 | 0.05 | 3,633 | 0.05 |
| 2003 | 2.43 | 6,532 | 0.05 | 4,192 | 0.05 |
| 2004 | 2.89 | 6,694 | 0.06 | 4,744 | 0.05 |
| 2005 | 2.36 | 6,950 | 0.06 | 4,991 | 0.06 |
| 2006 | 3.01 | 7,497 | 0.05 | 5,265 | 0.06 |
| 2007 | 2.09 | 7,664 | 0.05 | 5,682 | 0.06 |
| 2008 | 1.73 | 7,432 | 0.05 | 5,983 | 0.06 |
| 2009 | 2.29 | 7,076 | 0.05 | 6,046 | 0.05 |
| 2010 | 2.03 | 6,818 | 0.05 | 5,681 | 0.05 |
| 2011 | 1.48 | 6,300 | 0.05 | 5,400 | 0.05 |
| 2012 | 2.20 | 5,751 | 0.05 | 5,081 | 0.05 |
| 2013 | 3.04 | 5,665 | 0.05 | 4,478 | 0.05 |
| 2014 | 2.24 | 5,844 | 0.06 | 4,020 | 0.06 |
| 2015 | 2.28 | 5,960 | 0.06 | 4,181 | 0.06 |
| 2016 | 1.99 | 6,138 | 0.06 | 4,404 | 0.06 |
| 2017 | 2.09 | 6,261 | 0.06 | 4,697 | 0.06 |
| 2018 | 2.48 | 6,363 | 0.07 | 4,868 | 0.06 |
| 2019 | 2.29 | 6,341 | 0.10 | 4,850 | 0.07 |
| 2020 | 2.80 | 6,460 | 0.14 | 4,788 | 0.09 |

Table 12. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1 b for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to $1985-2020$. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the <br> Model ( $\geq \mathbf{1 0 1}$ mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=19,954 \\ & M M B_{350}=5,810 \end{aligned}$ |  |  |  |
| 1985 | 4.35 | 12,824 | 0.04 | 10,428 | 0.09 |
| 1986 | 6.21 | 10,612 | 0.04 | 9,773 | 0.06 |
| 1987 | 4.75 | 9,896 | 0.04 | 6,664 | 0.06 |
| 1988 | 3.51 | 8,731 | 0.04 | 6,236 | 0.05 |
| 1989 | 4.29 | 6,404 | 0.04 | 5,658 | 0.04 |
| 1990 | 3.32 | 5,574 | 0.04 | 3,512 | 0.05 |
| 1991 | 2.35 | 4,984 | 0.04 | 3,153 | 0.05 |
| 1992 | 2.64 | 4,899 | 0.04 | 3,011 | 0.05 |
| 1993 | 2.72 | 5,602 | 0.03 | 3,127 | 0.05 |
| 1994 | 2.94 | 5,037 | 0.03 | 3,801 | 0.03 |
| 1995 | 2.94 | 4,902 | 0.03 | 3,025 | 0.03 |
| 1996 | 2.48 | 4,625 | 0.03 | 2,947 | 0.03 |
| 1997 | 2.78 | 4,679 | 0.03 | 2,971 | 0.03 |
| 1998 | 2.69 | 5,093 | 0.03 | 3,026 | 0.04 |
| 1999 | 3.37 | 5,151 | 0.04 | 3,394 | 0.03 |
| 2000 | 3.65 | 5,255 | 0.04 | 3,340 | 0.04 |
| 2001 | 3.59 | 5,691 | 0.05 | 3,339 | 0.04 |
| 2002 | 3.49 | 6,288 | 0.05 | 3,669 | 0.05 |
| 2003 | 2.33 | 6,529 | 0.06 | 4,220 | 0.05 |
| 2004 | 2.96 | 6,681 | 0.06 | 4,757 | 0.06 |
| 2005 | 2.51 | 7,002 | 0.06 | 4,970 | 0.06 |
| 2006 | 3.05 | 7,617 | 0.05 | 5,250 | 0.07 |
| 2007 | 2.09 | 7,804 | 0.05 | 5,732 | 0.06 |
| 2008 | 1.69 | 7,558 | 0.05 | 6,083 | 0.06 |
| 2009 | 2.24 | 7,164 | 0.05 | 6,162 | 0.05 |
| 2010 | 2.02 | 6,870 | 0.05 | 5,778 | 0.05 |
| 2011 | 1.52 | 6,347 | 0.05 | 5,460 | 0.05 |
| 2012 | 2.24 | 5,814 | 0.05 | 5,113 | 0.05 |
| 2013 | 2.96 | 5,709 | 0.06 | 4,512 | 0.05 |
| 2014 | 2.17 | 5,834 | 0.06 | 4,064 | 0.06 |
| 2015 | 2.19 | 5,883 | 0.06 | 4,197 | 0.06 |
| 2016 | 1.88 | 5,982 | 0.07 | 4,369 | 0.07 |
| 2017 | 2.03 | 6,040 | 0.08 | 4,594 | 0.07 |
| 2018 | 2.34 | 6,067 | 0.10 | 4,691 | 0.08 |
| 2019 | 2.10 | 5,930 | 0.14 | 4,614 | 0.10 |
| 2020 | 2.76 | 5,930 | 0.19 | 4,471 | 0.13 |

Table 13. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1c for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to 1985-2020. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | $\begin{aligned} & \text { Recruits to the } \\ & \text { Model }(\geq 101 \\ & \text { mm CL) } \end{aligned}$ | Mature Male Biomass (Bent-Point fit) | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,246 \\ & M M B_{35 \%}=7,213 \end{aligned}$ |  |  |  |
| 1985 | 4.02 | 12,991 | 0.04 | 10,771 | 0.08 |
| 1986 | 6.25 | 10,725 | 0.04 | 10,086 | 0.06 |
| 1987 | 4.59 | 10,007 | 0.04 | 6,873 | 0.06 |
| 1988 | 3.42 | 8,818 | 0.04 | 6,431 | 0.05 |
| 1989 | 4.07 | 6,472 | 0.04 | 5,834 | 0.04 |
| 1990 | 3.14 | 5,618 | 0.04 | 3,676 | 0.05 |
| 1991 | 2.27 | 5,027 | 0.04 | 3,300 | 0.05 |
| 1992 | 2.51 | 4,917 | 0.04 | 3,133 | 0.05 |
| 1993 | 2.73 | 5,600 | 0.03 | 3,222 | 0.05 |
| 1994 | 2.91 | 5,034 | 0.03 | 3,854 | 0.03 |
| 1995 | 2.88 | 4,876 | 0.03 | 3,056 | 0.03 |
| 1996 | 2.50 | 4,599 | 0.03 | 2,967 | 0.03 |
| 1997 | 2.70 | 4,643 | 0.03 | 2,976 | 0.04 |
| 1998 | 2.64 | 5,031 | 0.03 | 3,032 | 0.04 |
| 1999 | 3.34 | 5,072 | 0.04 | 3,385 | 0.03 |
| 2000 | 3.57 | 5,147 | 0.04 | 3,309 | 0.04 |
| 2001 | 3.56 | 5,558 | 0.04 | 3,292 | 0.04 |
| 2002 | 3.40 | 6,118 | 0.05 | 3,596 | 0.04 |
| 2003 | 2.39 | 6,357 | 0.05 | 4,125 | 0.05 |
| 2004 | 2.68 | 6,434 | 0.05 | 4,632 | 0.05 |
| 2005 | 2.13 | 6,556 | 0.05 | 4,832 | 0.06 |
| 2006 | 3.08 | 7,059 | 0.05 | 5,017 | 0.06 |
| 2007 | 2.09 | 7,262 | 0.05 | 5,322 | 0.06 |
| 2008 | 1.70 | 7,056 | 0.05 | 5,611 | 0.05 |
| 2009 | 2.38 | 6,760 | 0.05 | 5,703 | 0.05 |
| 2010 | 2.00 | 6,558 | 0.04 | 5,369 | 0.05 |
| 2011 | 1.40 | 6,045 | 0.04 | 5,153 | 0.05 |
| 2012 | 2.14 | 5,478 | 0.05 | 4,873 | 0.05 |
| 2013 | 2.92 | 5,348 | 0.05 | 4,266 | 0.05 |
| 2014 | 2.26 | 5,509 | 0.05 | 3,784 | 0.05 |
| 2015 | 2.36 | 5,671 | 0.05 | 3,905 | 0.06 |
| 2016 | 1.98 | 5,896 | 0.05 | 4,126 | 0.06 |
| 2017 | 1.88 | 5,976 | 0.05 | 4,464 | 0.05 |
| 2018 | 2.30 | 5,973 | 0.06 | 4,663 | 0.05 |
| 2019 | 2.20 | 5,875 | 0.08 | 4,588 | 0.06 |
| 2020 | 2.68 | 5,527 | 0.12 | 4,433 | 0.07 |

Table 14. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1d for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961-2021 are restricted to 1985-2020. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | $\begin{aligned} & \text { Recruits to the } \\ & \text { Model ( } \geq 101 \\ & \text { mm CL) } \end{aligned}$ | Mature Male Biomass (Bent-Point fit) | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,779 \\ & M M B_{35 \%}=7,422 \end{aligned}$ |  |  |  |
| 1985 | 4.15 | 12,880 | 0.04 | 10,734 | 0.08 |
| 1986 | 6.17 | 10,601 | 0.04 | 9,941 | 0.06 |
| 1987 | 4.71 | 9,872 | 0.04 | 6,730 | 0.06 |
| 1988 | 3.52 | 8,702 | 0.04 | 6,256 | 0.05 |
| 1989 | 4.34 | 6,386 | 0.04 | 5,652 | 0.04 |
| 1990 | 3.35 | 5,565 | 0.04 | 3,501 | 0.05 |
| 1991 | 2.35 | 4,971 | 0.04 | 3,145 | 0.05 |
| 1992 | 2.62 | 4,873 | 0.04 | 3,005 | 0.05 |
| 1993 | 2.70 | 5,559 | 0.03 | 3,116 | 0.05 |
| 1994 | 3.03 | 5,008 | 0.03 | 3,777 | 0.03 |
| 1995 | 2.99 | 4,907 | 0.03 | 2,990 | 0.03 |
| 1996 | 2.47 | 4,645 | 0.03 | 2,936 | 0.03 |
| 1997 | 2.79 | 4,710 | 0.04 | 2,990 | 0.04 |
| 1998 | 2.80 | 5,171 | 0.03 | 3,058 | 0.04 |
| 1999 | 3.44 | 5,291 | 0.04 | 3,440 | 0.03 |
| 2000 | 3.69 | 5,448 | 0.04 | 3,439 | 0.04 |
| 2001 | 3.64 | 5,930 | 0.05 | 3,496 | 0.04 |
| 2002 | 3.57 | 6,583 | 0.05 | 3,872 | 0.05 |
| 2003 | 2.36 | 6,869 | 0.06 | 4,470 | 0.05 |
| 2004 | 2.90 | 7,012 | 0.06 | 5,057 | 0.06 |
| 2005 | 2.38 | 7,248 | 0.06 | 5,295 | 0.06 |
| 2006 | 3.05 | 7,790 | 0.06 | 5,542 | 0.07 |
| 2007 | 2.11 | 7,946 | 0.05 | 5,940 | 0.06 |
| 2008 | 1.75 | 7,693 | 0.05 | 6,230 | 0.06 |
| 2009 | 2.36 | 7,333 | 0.05 | 6,278 | 0.06 |
| 2010 | 2.05 | 7,072 | 0.05 | 5,896 | 0.06 |
| 2011 | 1.52 | 6,546 | 0.05 | 5,616 | 0.05 |
| 2012 | 2.27 | 6,007 | 0.05 | 5,292 | 0.05 |
| 2013 | 3.07 | 5,933 | 0.06 | 4,689 | 0.06 |
| 2014 | 2.26 | 6,112 | 0.06 | 4,247 | 0.06 |
| 2015 | 2.33 | 6,234 | 0.06 | 4,419 | 0.07 |
| 2016 | 2.03 | 6,422 | 0.06 | 4,643 | 0.07 |
| 2017 | 2.12 | 6,547 | 0.07 | 4,941 | 0.07 |
| 2018 | 2.55 | 6,657 | 0.08 | 5,115 | 0.07 |
| 2019 | 2.33 | 6,646 | 0.11 | 5,100 | 0.08 |
| 2020 | 2.90 | 6,854 | 0.16 | 5,049 | 0.10 |

Table 15. Negative log-likelihood values of the fits for models 21.1a (base), 21.1b, 21.1c, 21.1d, 21.1a2, and 21.1b2 for golden king crab in the WAG. Likelihood components with zero entry in the entire rows are omitted.

| Likelihood Component | 21.1a | 21.1b | 21.1c | 21.1d | 21.1a2 | 21.1b2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of free parameters | 152 | 152 | 154 | 152 | 152 | 152 |
| Retlencomp | -2045.3400 | -2043.7300 | -2030.2200 | -2023.9300 | -2045.3400 | -2043.7300 |
| Totallencomp | -1600.0700 | -1626.2200 | -1592.1800 | -1624.8000 | -1600.0700 | -1626.2200 |
| Observer cpue | -43.1882 | -28.6586 | -49.1695 | -27.6625 | -43.1882 | -28.6586 |
| Fishery cpue | -17.2600 | -17.3922 | -21.2682 | -13.6626 | -17.2600 | -17.3922 |
| RetdcatchB | 5.7765 | 6.1631 | 7.2070 | 6.9146 | 5.7765 | 6.1631 |
| TotalcatchB | 43.9060 | 44.8060 | 44.9945 | 43.8921 | 43.9060 | 44.8060 |
| GdiscdcatchB | 0.0003 | 0.0002 | 0.0005 | 0.0001 | 0.0003 | 0.0002 |
| Rec_dev | 23.8881 | 23.9941 | 23.3176 | 23.1456 | 23.8881 | 23.9941 |
| Pot F_dev | 0.0274 | 0.0276 | 0.0262 | 0.0285 | 0.0274 | 0.0276 |
| Gbyc_F_dev | 0.0477 | 0.0476 | 0.0490 | 0.0470 | 0.0477 | 0.0476 |
| Tag | 2691.8400 | 2692.1700 | 2691.9200 | 2691.5700 | 2691.8400 | 2692.1700 |
| RetcatchN | 0.0087 | 0.0083 | 0.0062 | 0.0067 | 0.0087 | 0.0083 |
| Total | -940.3570 | -948.7790 | -925.3140 | -924.4570 | -940.3570 | -948.7790 |

Table 16. Stock status, reference biomass and fishing mortality, OFL (total catch), and ABC for various models for EAG, WAG, and AI golden king crab stock.

EAG: Biomass, OFL, and ABC are in t. Current $\mathrm{MMB}=\mathrm{MMB}$ in 2021.

|  |  |  | Current | MMB/ |  | $\mathrm{M}\left(\mathrm{yr}^{-1}\right)$ | OFL | MaxABC | ABC | ABC <br> $\left(\mathrm{P}^{*}=0.49\right)$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model | Tier | $M M B_{35 \%}$ | MMB | $M M B_{35 \%}$ | $F_{\text {OFL }}$ | $F_{35 \%}$ |  |  | $(0 F L)$ | $(0.70 * \mathrm{OFL})$ |  |
| 21.1a | 3a | 9,298 | 11,039 | 1.19 | 0.64 | 0.64 | 0.21 | 3,795 | 3,775 | 2,846 | 2,657 |
| 21.1b | 3a | 9,157 | 9,834 | 1.07 | 0.65 | 0.65 | 0.21 | 3,212 | 3,195 | 2,409 | 2,248 |
| 21.1c | 3b | 8,974 | 8,279 | 0.92 | 0.60 | 0.66 | 0.21 | 2,204 | 2,182 | 1,653 | 1,543 |
| 21.1a2 | 3a | 8,999 | 10,668 | 1.19 | 0.56 | 0.56 | 0.21 | 3,416 | 3,398 | 2,562 | 2,391 |
| 21.1b2 | 3a | 8,848 | 9,417 | 1.06 | 0.57 | 0.57 | 0.21 | 2,897 | 2,881 | 2,172 | 2,028 |

WAG: Biomass, OFL, and ABC are in t. Current $\mathrm{MMB}=\mathrm{MMB}$ in 2021.

|  |  |  | Current | MMB/ |  | $\mathrm{M}\left(\mathrm{yr}^{-1}\right)$ | OFL | MaxABC | ABC | ABC <br> Model |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Tier | $M M B_{35 \%}$ | MMB | $M M B_{35 \%}$ | $F_{\text {OFL }}$ | $F_{35 \%}$ |  |  | $\left(\mathrm{P}^{*}=0.49\right)$ | $(0.75 * \mathrm{OFL})$ | $(0.70 * \mathrm{OFL})$ |
| 21.1a | 3b | 7,370 | 6,702 | 0.91 | 0.57 | 0.63 | 0.21 | 1,669 | 1,659 | 1,252 |
| 21.1b | 3b | 7,354 | 6,378 | 0.87 | 0.54 | 0.63 | 0.21 | 1,446 | 1,435 | 1,085 |
| 21.1c | 3 b | 7,089 | 6,069 | 0.86 | 0.55 | 0.66 | 0.21 | 1,314 | 1,307 | 986 |
| 21.1d | 3 b | 7,441 | 6,960 | 0.94 | 0.58 | 0.62 | 0.21 | 1,836 | 1,825 | 1,377 |
| 21.1a2 | 3 b | 7,157 | 6,260 | 0.87 | 0.46 | 0.54 | 0.21 | 1,422 | 1,414 | 1,067 |
| 21.1b2 | 3b | 7,078 | 5,897 | 0.83 | 0.45 | 0.55 | 0.21 | 1,247 | 1,237 | 935 |

AI: OFL and ABC are in t .

|  | OFL | ABC <br> $\left(0.75^{*} \mathrm{OFL}\right)$ | ABC <br> $(0.7 * \mathrm{OFL})$ |
| :---: | :---: | :---: | :---: |
| 21.1 a | 5,464 | 4,098 | 3,826 |
| 21.1 b | 4,658 | 3,494 | 3,260 |
| 21.1 c | 3,518 | 2,639 | 2,463 |
| 21.1 a 2 | 4,838 | 3,629 | 3,387 |
| 21.1 b 2 | 4,144 | 3,107 | 2,901 |



Figure 1. Historical commercial harvest (from fish tickets; metric tons) and catch-per-unit effort (CPUE, number of crab per pot lift) of golden king crab in the EAG, 1985/86-2020/21 fisheries (note: 1985 refers to the 1985/86 fishing year).


Figure 2. Historical commercial harvest (from fish tickets; metric tons) and catch-per-unit effort (CPUE, number of crab per pot lift) of golden king crab in the WAG, 1985/86-2020/21 fisheries (note: 1985 refers to the 1985/86 fishing year).


Figure 3a. Predicted (black line) vs. observed (bar) retained catch relative length frequency distributions for model 21.1a for golden king crab in the EAG, 1985/86 to 2020/21.


Figure 3b. Predicted (black line) vs. observed (bar) retained catch relative length frequency distributions for model 21.1b for golden king crab in the EAG, 1985/86 to 2020/21.


Figure 3c. Predicted (black line) vs. observed (bar) retained catch relative length frequency distributions for model 21.1c for golden king crab in the EAG, 1985/86 to 2020/21.


Figure 4a. Predicted (black line) vs. observed (bar) total catch relative length frequency distributions for model 21.1a for golden king crab in the EAG, 1990/91 to 2020/21.


Figure 4b. Predicted (black line) vs. observed (bar) total catch relative length frequency distributions for model 21.1b for golden king crab in the EAG, 1990/91 to 2020/21.


Figure 4c. Predicted (black line) vs. observed (bar) total catch relative length frequency distributions for model 21.1c for golden king crab in the EAG, 1990/91 to 2020/21.


Figure 5a. Predicted (black line) vs. observed (bar) groundfish discarded catch relative length frequency distributions for model 21.1a for golden king crab in the EAG, 1989/90 to 2020/21.


Figure 5b. Predicted (black line) vs. observed (bar) groundfish discarded catch relative length frequency distributions for model 21.1b for golden king crab in the EAG, 1989/90 to 2020/21.


Figure 5c. Predicted (black line) vs. observed (bar) groundfish discarded catch relative length frequency distributions for model 21.1c for golden king crab in the EAG, 1989/90 to 2020/21.


Figure 6. Estimated total (black solid line) and retained selectivity (red dotted line) for pre- and post- rationalization periods for models 21.1a, 21.1b, and 21.1c fits to golden king crab data in the EAG.


Figure 7. Comparison of input CPUE indices [black open circles with $+/-2$ SE for model 21.1a (left) and model 21.1b (right)] with predicted CPUE indices (colored solid lines) by 21.1a, 21.1b, and 21.1c, model fits for EAG golden king crab data, 1985/86-2020/21. Model estimated additional standard error was added to each input standard error.


Figure 8. Estimated number of male recruits (millions of crab $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) for 21.1a, 21.1b, and 21.1c model fits to EAG golden king crab data, 1961-2021.


Figure 9. Observed (open circle) vs. predicted (solid line) retained catch (top left), total catch (top right), and groundfish bycatch (bottom left) of golden king crab for 21.1a, 21.1b, and 21.1c model fits to EAG data, 1981/82-2020/21.

## Retained Catch



Figure 10. Observed (open circle) vs. predicted (solid line) retained catch of golden king crab for 21.1a, 21.1b, and 21.1c model fits to EAG data, 1981/82-1984/85. Note: Input retained catches to the model during pre-1985 fishery period was in number of crab.

## EAG



Figure 11. Trends in pot fishery full selection total fishing mortality of golden king crab for 21.1a, 21.1b, and 21.1c model fits to EAG (left) and for 21.1a, 21.1b, 21.1c, and 21.1d model fits to WAG (right) data, 1981/82-2020/21.


Figure 12. Trends in golden king crab mature male biomass for 21.1a, 21.1b, and 21.1c model fits to EAG (left) and for 21.1a, 21.1b, 21.1c, and 21.1d model fits to WAG (right) data. Top: 1960/61-2020/21, bottom: 2005/06-2020/21. Model21.1a estimate has two standard error confidence limits.


Figure 13a. Predicted (black line) vs. observed (bar) retained catch relative length frequency distributions for model 21.1a for golden king crab in the WAG, 1985/86 - 2020/21.


Figure 13b. Predicted (black line) vs. observed (bar) retained catch relative length frequency distributions for model 21.1b for golden king crab in the WAG, 1985/86-2020/21.


Figure 13c. Predicted (black line) vs. observed (bar) retained catch relative length frequency distributions for model 21.1c for golden king crab in the WAG, 1985/86 - 2020/21.


Figure 14a. Predicted (black line) vs. observed (bar) total catch relative length frequency distributions for model 21.1a for golden king crab in the WAG, 1990/91-2020/21.


Figure 14b. Predicted (black line) vs. observed (bar) total catch relative length frequency distributions for model 21.1b for golden king crab in the WAG, 1990/91-2020/21.


Figure 14c. Predicted (black line) vs. observed (bar) total catch relative length frequency distributions for model 21.1c for golden king crab in the WAG, 1990/91-2020/21.


Figure 15a. Predicted (black line) vs. observed (bar) groundfish discard catch relative length frequency distributions for model 21.1a for golden king crab in the WAG, 1989/90 - 2020/21.


Figure 15b. Predicted (black line) vs. observed (bar) groundfish discard catch relative length frequency distributions for model 21.1b for golden king crab in the WAG, 1989/90 - 2020/21.


Figure 15c. Predicted (black line) vs. observed (bar) groundfish discard catch relative length frequency distributions for model 21.1c for golden king crab in the WAG, 1989/90 - 2020/21.


Figure 16. Estimated total (black solid line) and retained selectivity (red dotted line) for pre- and post- rationalization periods for models 21.1a, 21.1b, and 21.1c fits to golden king crab data in the WAG.


Figure 17. Comparison of input CPUE indices [black open circles with $+/-2$ SE for model 21.1a (left) and model 21.1b (right)] with predicted CPUE indices (colored solid lines) by 21.1a, 21.1b, 21.1c, and 21.1d, model fits for WAG golden king crab data, 1985/86$2020 / 21$. Model estimated additional standard error was added to each input standard error.


Figure 18. Estimated number of male recruits (millions of crab $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) for 21.1a, 21.1b, 21.1c, and 21.1d model fits to WAG golden king crab data, 1961-2021.


Figure 19. Observed (open circle) vs. predicted (solid line) retained catch (top left), total catch (top right), and groundfish bycatch (bottom left) of golden king crab for 21.1a, 21.1b, 21.1c, and 21.1d model fits to WAG data, 1981/82-2020/21.

Retained Catch


Figure 20. Observed (open circle) vs. predicted (solid line) retained catch of golden king crab for 21.1a, 21.1b, 21.1c, and 21.1d model fits to WAG data, 1981/82-1984/85. Note: Input retained catches to the model during pre-1985 fishery period was in number of crab.


Figure 21. Comparison of 21.1a, 21.1a1 (a higher $M$ of $0.38 \mathrm{yr}^{-1}$ was used for years $>1998$ in the assessment), and 21.1c models' retrospective fits for EAG.

## Appendix A: Catch and CPUE data

## Observer data collection protocol:

Observer data have been collected since 1988 (Moore et al. 2000; Barnard et al. 2001; Barnard and Burt 2004; Gaeuman 2011), but data were not comprehensive in the initial years, so a shorter time series of data for the period 1990/91-2020/21 was selected for this analysis. During 1990/91-1994/95, observers were only deployed on catcher-processor vessels. During 1995/96-2004/05, observers were deployed on all fishing vessels during fishing activity. Observers have been deployed on all fishing vessels since 2005/06, but catcher-only vessels are only required to carry observers for a minimum of $50 \%$ of their fishing activity during a season; catcher-processor vessels are still required to carry observers during all fishing activity. Onboard observers sample seven pots per day (may be different numbers of pots per string) and count and measure all crabs caught and categorize catch as females, sublegal males, retained legal males, and non-retained legal males in a sampled pot. Prior to the 2009/10 season, depending on season, area, and type of fishing vessel, observers were also instructed to sample additional pots in which all crab were only counted and categorized as females, sublegal males, retained legal males, and non-retained legal males, but were not measured. Annual mean nominal CPUEs of retained and total crabs were estimated considering all sampled pots within each season (Table 3). The observer CPUE data collection improved over the years and the data since 1995/96 are more reliable. Thus, for model fitting, the observer CPUE time series was restricted to 1995/96-2020/21. The 1990/91-2020/21 observer database consists of 118,552 records and that of 1995/96-2020/21 contains 114,273 records.

We detected some computational errors in raw size frequency summary data preparation (observer and fish ticket sampling) for 2016-2019 and rectified errors in relative retained and total size frequency computations in the current analysis. The correction of errors did not affect retained catch crab distribution by size bins but caused minor changes to allocation of total catch crab into size bins.

Length-specific CPUE data collected by observers provides information on a wider size range of the stock than did the commercial catch length frequency data obtained from mostly legalsized landed males.

## Retained catch by size-class:

The commercial catch and length frequency distribution were estimated from ADF\&G landing records and dockside sampling (Bowers et al. 2008, 2011). The annual retained catch, total catch, and groundfish (or trawl) discarded mortality are provided in Tables 1, 2, and 2 b for EAG and WAG. The weighted length frequency data were used to distribute the catch into 5mm size intervals. The length frequency data for a year were weighted by each sampled vessel's catch as follows. The $i$-th length-class frequency was estimated as:

$$
\begin{equation*}
\sum_{j=1}^{k} C_{j} \frac{L F_{j, i}}{\sum_{i=1}^{n} L F_{j, i}} \tag{A.1}
\end{equation*}
$$

where $k=$ number of sampled vessels in a year, $L F_{j, i}=$ number of crabs in the $i$-th length-class in the sample from $j$-th vessel, $n=$ number of size classes, $C_{j}=$ number of crabs caught by $j$-th
vessel. Then the relative frequency for the year was calculated and applied to the annual retained catch (in number of crabs) to obtain retained catch by length-class.

## Total catch by size-class:

The annual total catch (in number of crabs) was estimated by the observer nominal (unstandardized) total CPUE considering all vessels multiplied by the total fishing effort (number of pot lifts). The weighted length frequency of the observer samples across the fleet was estimated using Equation A.1. Observer measurement of crab ranged from 20 to 220 mm CL. To restrict the total number of crabs to the model assumed size range ( $101-185+\mathrm{mm}$ CL), the proportion of observer total relative length frequency corresponding to this size range was multiplied by the total catch (number of crabs). This total number of crabs was distributed into length-classes using the weighted relative length frequency. Thus, crab sizes $<101 \mathrm{~mm}$ CL were excluded from the model. In addition, all crab $>185 \mathrm{~mm}$ CL were pooled into a plus length class. Note that the total crab catches by size that went into the model did not consider retained and discard components separately. However, once the model estimated the annual total catch, then retained catch was deducted from this total and multiplied by handling mortality [we used a $20 \%$ handling mortality (Siddeek et al. 2005) to obtain the directed fishery discarded (dead) catch].

## Observer CPUE standardization:

For CPUE standardization, observer data were restricted to the 1995/96-2020/21 period for reliability of data and further reduced by $5 \%$ cutoff of Soak time and $1 \%$ cutoff of Depth on both ends of the variable range to remove unreliable data or data from dysfunctional pot operations and restricting to vessels which have made five trips per year for at least three years during 1985/86-2020/21. There were significant changes in fishing practice due to changes in management regulations (e.g., since 1996/97 constant TAC and since 2005/06 crab rationalization), pot configuration (escape web on the pot door increased to 9 " since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two separate observer CPUE time series, 1995/962004/05 and 2005/06-2020/21, to estimate CPUE indices for model input.

## Fishery CPUE standardization:

To include a long time series of CPUE indices for stock abundance contrast, we also considered the 1985/86-1998/99 legal size standardized CPUE (i.e., fishery CPUE indices) as a separate likelihood component in all scenarios. Because of the lack of soak time data before 1990, we estimated the CPUE index considering a limited set of explanatory variables (e.g., vessel, captain, area, month) and fitting the negative binomial GLM model to fish ticket data (Tables 4 and 14).

When using CPUE indices in the model fit, we compared the predicted with the observed legal male CPUE in the observer CPUE likelihoods because legal male (retained plus non-retained) data are more reliable than total in the observer samples.

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2018). Following a suggestion from the CIE reviewers in June 2018 we reduced the number of gear codes in the database after consulting with the fishing industry
(Rip Carlton, Chad Hoefer, and Scott Goodman, personal communication December 2018; Table B1). Following an SSC suggestion in October 2018, we used a hybrid procedure: First, we selected a scope of variables set by Akike Information Criterion, AIC (Burnham and Anderson 2002). An increase of more than 2 units in the AIC was used to identify the variable to be included successively (stepAIC program, R Core Team 2020). Then, the model parsimony was improved further by successively removing the term that explained the least proportion of deviance ( $\mathrm{R}^{2}<0.01$ ) (stepCPUE R function was used, Siddeek et al. 2018). Feenstra, et al. (2019) used a similar hybrid approach.

Table A.1. Updated gear codes for observer data analysis. Only gear codes \# 5, 6, 7, 8, and 13 were considered following crab industry suggestion. Note: Identical codes were given to those gear codes with similar catchability/selectivity. X indicates gear codes that were ignored.

| Original <br> Gear code | Pot gear description | Mark X against the code that can be ignored | Number encountered by observers during $1990-2016$ | Updated gear code |
| :---: | :---: | :---: | :---: | :---: |
| 1 | Dungeness crab pot, small \& round | X | - 2 | X |
| 2 | Pyramid pot, tunnel openings usually on sides, stackable | X | 2121 | X |
| 3 | Conical pot, opening at top of cone, stackable | X | 2000 | X |
| 4 | $4^{\prime} \mathrm{X} 4$ ' rectangular pot |  | 60 | X |
| 5 | $5{ }^{\prime} \mathrm{X} 5$ ' rectangular pot |  | 18032 | 5 |
| 6 | $6{ }^{\prime} \mathrm{X} 6$ ' rectangular pot |  | 17508 | 6 |
| 7 | $7{ }^{\prime} \mathrm{X} 7$ ' rectangular pot |  | 23806 | 7 |
| 8 | $8^{\prime} \mathrm{X} 8{ }^{\prime}$ rectangular pot |  | 1936 | 8 |
| 9 | $51 / 2^{\prime} \mathrm{X} 51 / 2^{\prime}$ rectangular pot |  | 6934 | 5 |
| 10 | $61 / 2^{\prime} \mathrm{X} 61 / 2^{\prime}$ rectangular pot |  | 22085 | 6 |
| 11 | $71 / 2^{\prime} \mathrm{X} 71 / 2^{\prime}$ rectangular pot |  | 387 | 7 |
| 12 | Round king crab pot, enlarged version of Dungeness crab pot |  | 8259 | X |
| 13 | $10{ }^{\prime} \mathrm{X} 10{ }^{\prime}$ rectangular pot |  | 466 | 13 |
| 14 | 9' X 9 ' rectangular pot | X | 1 | X |
| 15 | $81 / 2^{\prime} \mathrm{X} 81 / 2^{\prime}$ rectangular pot | X | 1 | X |
| 16 | $91 / 2^{\prime} \mathrm{X} 91 / 2^{\prime}$ rectangular pot | X | Not used | X |
| 17 | $8^{\prime} \mathrm{X} 9^{\prime}$ rectangular pot | X | 1 | X |
| 18 | $8^{\prime} \mathrm{X} 10{ }^{\prime}$ rectangular pot | X | 1 | X |
| 19 | $9^{\prime} \mathrm{X} 10{ }^{\text {' rectangular pot }}$ |  | Not used | X |
| 20 | $7{ }^{\prime} \mathrm{X} \mathrm{8}$ ' rectangular pot | X | 252 | X |
| 21 | Hair crab pot, longlined and small, stackable |  | Not used | X |
| 22 | snail pot | X | 1 | X |
| 23 | Dome-shaped pot, tunnel opening on top, often longlined in deep-water fisheries | X | 6756 | X |
| 24 | ADF\&G shellfish research 7’ X 7' X34" rectangular pot with 2.75 " stretch mesh and no escapement rings or mesh |  | Research pot | X |

All scenarios used CPUE indices estimated by the hybrid GLM method. Following a January 2019 CPT request, we considered a Year:Area interaction factor as a special case for a CPUE standardization scenario.

Thus we estimated two sets of observer CPUE indices for model input, 21.1a (reduced number of gear codes and no interaction), and 21.1c (reduced number of gear codes and Year:Area interaction).

## Observer CPUE index by GLM

## a. Non-interaction GLM model

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2016b). We considered the negative binomial GLM on positive and zero catches to select the explanatory variables. The response variable CPUE is the observer sample catch record for a pot haul. The negative binomial model uses the log link function for the GLM fit.

For the non-interaction model, we assumed the null model to be:

$$
\begin{equation*}
\ln \left(\text { CPUE }_{\mathrm{i}}\right)=\text { Year }_{\mathrm{y}_{\mathrm{i}}} \tag{A.2}
\end{equation*}
$$

where Year is a factorial variable.
The maximum set of model terms offered to the stepwise selection procedure was:
$\ln \left(\right.$ CPUE $\left._{\mathrm{I}}\right)=$ Year $_{\mathrm{y}_{\mathrm{i}}}+\mathrm{ns}\left(\right.$ Soak $\left._{\text {si }}, \mathrm{df}\right)+$ Month $_{\mathrm{m}_{\mathrm{i}}}+$ Vessel $_{\mathrm{vi}}+$ Captain $_{\mathrm{ci}}+$ Area $_{\mathrm{ai}}+$ Gear $_{\text {gi }}+\mathrm{ns}\left(\right.$ Depth $\left._{\text {di }}, \mathrm{df}\right)$,
where Soak is in unit of days and is numeric; Month, Area (Block) code, Vessel code, Captain code, and Gear code are factorial variables; Depth in fathom is a numeric variable; ns=cubic spline, and $\mathrm{df}=$ degree of freedom.

We used a log link function and a dispersion parameter $(\theta)$ in the GLM fitting process. We used the $\mathrm{R}^{2}$ criterion for predictor variable selection (Siddeek et al. 2016b).

We calculated appropriate degrees of freedom and dispersion parameters by calculating AICs for a range of values and locating the best values at the minimum AIC (Figures A. 1 and A.2, respectively).


Figure A.1. AIC vs degrees of freedom for soak time and depth during pre- and postrationalization periods for EAG (top) and WAG (bottom). Vertical lines identify the optimum degrees of freedom values chosen for CPUE standardization.


Figure A.2. AIC vs theta (dispersion parameter) during pre- and post-rationalization periods for EAG (top) and WAG (bottom). Vertical lines identify the optimum theta values chosen for CPUE standardization.

Figures A. 3 to A. 6 depict the fit of smoothers to observed CPUE data for a range of Soak time values at a given set of fixed values of other predictor variables chosen in the final model. For simplicity, the fits are shown for a single year.


Figure A.3. Smoother fit to 2004 observed CPUE data for EAG.


Figure A.4. Smoother fit to 2014 observed CPUE data for EAG.


Figure A.5. Smoother fit to 2004 observed CPUE data for WAG.


Figure A.6. Smoother fit to 2014 observed CPUE data for WAG.

Instead of using the traditional AIC ( $-2 \log$ _likelihood +2 p ) we used the Consistent Akaike Information Criteria (CAIC) (Bozdogan 1987) $\left\{-2 \log _{\text {_ }}\right.$ likelihood $\left.+[\ln (\mathrm{n})+1]^{*} \mathrm{p}\right\}$ for variable selection by StepAIC, where $n=$ number of observations and $p=$ number of parameters to be estimated. The number of selected variables were further reduced for parsimony, if feasible, by the $\mathrm{R}^{2}$ criterion using the StepCPUE function. i.e., a hybrid selection procedure (Feenstra et al. 2019).

AIC selected high values of smoother functions' degrees of freedom for some data sets were criticized by the CPT/SSC in May/June 2021. We addressed this concern following a different approach by selecting the final CPUE model predictor variables after removing several nonsignificant variable subcomponents. In particular, the degrees of freedom of smoothers were readjusted to obtain significant degrees of freedom parameter estimates for the final CPUE model (see Tables A. 2 to A. 5 below).

Example R codes used for main effect GLM fitting are as follows:
For EAG 1995_04 CPUE indices:
library(MASS)
library(splines)
Step 1:

$$
\text { glm.object }<- \text { glm(Legals } \sim \text { Year,family }=\text { negative.binomial(1.38),data=datacore })
$$

epotsampleoutAIC<-stepAIC(glm.object,scope=list(upper=
$\sim($ Year $+n s(S o a k D a y s, d f=4)+$ Month + Vessel + Captain + Area + Gear $+n s($ Depth, $d f=9)$ ),lower $=$ $\sim$ Year),family=negative.binomial(1.38),direction="forward",trace $=9, k=\log$ (nrow(datacore)) +1.0 )

Step 2:
glm.object $<-$ glm(Legals $\sim$ Year,family $=$ negative.binomial(1.38),data=datacore)
epotsampleout<-
stepCPUE(glm.object,scope $=$ list(upper $=\sim($ Year + Gear + Captain $+n s(S o a k D a y s, d f=4)+$
Month + Area),lower $=\sim$ Year),family=negative.binomial(1.38),direction="forward",trace=9,r 2. change $=0.01$ )

The final main effect models for EAG were:
Model 21.1a:
Initial selection by stepAIC:

```
\(\ln (\) CPUE \()=\) Year + Gear + Captain + ns \((\) Soak, 4\()+\) Month + Area
AIC=203,808
```

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + ns $($ Soak, 4$)$
for the $1995 / 96-2004 / 05$ period $\left[\theta=1.38, \mathrm{R}^{2}=0.1813\right.$, $\left.\mathrm{AIC}=133,925\right]$
Initial selection by stepAIC:
$\ln$ (CPUE) $=$ Year + Captain + Gear + ns (Soak, 3$)$
AIC=77,311
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + Gear + ns (Soak, 3)
for the 2005/06-2020/21 period $\left[\theta=2.32, \mathrm{R}^{2}=0.1027, \mathrm{AIC}=75,185\right]$.
Tables A. 2 and A. 3 list the summary statistics of the main effects GLM fits to 1995/96-2004/05 and 2005/06-2020/21 data series, respectively:

Table A.2. Summary statistics of the main effects GLM fit to EAG 1995/96-2004/05 data.

|  | Std. |  |  |  |
| :--- | ---: | ---: | ---: | :--- |
| Parameters | Estimate | Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| (Intercept) | 1.4867 | 0.0417 | 35.6522 | 0.000000 |
| Year1996 | 0.0195 | 0.0395 | 0.4955 | 0.620268 |
| Year1997 | 0.1179 | 0.0385 | 3.0593 | 0.002222 |
| Year1998 | 0.2907 | 0.0356 | 8.1664 | 0.000000 |
| Year1999 | 0.2792 | 0.0374 | 7.4707 | 0.000000 |
| Year2000 | 0.2199 | 0.0360 | 6.1150 | 0.000000 |
| Year2001 | 0.5136 | 0.0382 | 13.4416 | 0.000000 |
| Year2002 | 0.6344 | 0.0418 | 15.1926 | 0.000000 |
| Year2003 | 0.4926 | 0.0405 | 12.1673 | 0.000000 |
| Year2004 | 0.8981 | 0.0462 | 19.4240 | 0.000000 |
| Captain133 | 0.3427 | 0.0291 | 11.7956 | 0.000000 |
| Captain161 | -0.3400 | 0.0569 | -5.9729 | 0.000000 |
| Captain204 | -0.8126 | 0.1004 | -8.0978 | 0.000000 |
| Captain208 | 0.3052 | 0.1014 | 3.0108 | 0.002609 |
| Captain210 | -0.1137 | 0.0367 | -3.0998 | 0.001939 |
| Captain213 | 0.4689 | 0.0313 | 14.9932 | 0.000000 |
| Captain219 | 0.3028 | 0.0302 | 10.0428 | 0.000000 |
| Captain232 | -0.4570 | 0.0505 | -9.0579 | 0.000000 |
| Captain240 | -0.6116 | 0.1657 | -3.6917 | 0.000223 |
| Captain247 | 0.1122 | 0.0273 | 4.1037 | 0.000041 |
| Captain271 | -0.4641 | 0.0475 | -9.7680 | 0.000000 |
| Captain272 | 0.2183 | 0.0311 | 7.0210 | 0.000000 |
| Captain287 | -0.4317 | 0.0888 | -4.8621 | 0.000001 |


| Captain300 | -0.4125 | 0.0832 | -4.9559 | 0.000001 |
| :--- | ---: | ---: | ---: | ---: |
| Captain302 | 0.1780 | 0.0330 | 5.3943 | 0.000000 |
| Captain344 | -1.3531 | 0.1069 | -12.6579 | 0.000000 |
| Captain354 | 0.2946 | 0.0500 | 5.8979 | 0.000000 |
| Captain358 | -0.8345 | 0.2274 | -3.6699 | 0.000243 |
| Captain384 | 0.4867 | 0.0283 | 17.2067 | 0.000000 |
| Captain388 | -0.7940 | 0.1086 | -7.3140 | 0.000000 |
| Captain390 | 0.8001 | 0.0437 | 18.3102 | 0.000000 |
| ns(SoakDays, df $=4) 1$ | 0.3371 | 0.0323 | 10.4333 | 0.000000 |
| ns(SoakDays, df $=4) 2$ | 0.5101 | 0.0364 | 14.0224 | 0.000000 |
| ns(SoakDays, df $=4) 3$ | 0.7180 | 0.0621 | 11.5545 | 0.000000 |
| ns(SoakDays, df $=4) 4$ | 0.4982 | 0.0535 | 9.3145 | 0.000000 |

Table A.3. Summary statistics of the main effects GLM fit to EAG 2005/06-2020/21 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| (Intercept) | 3.0056 | 0.0402 | 74.6797 | 0.000000 |
| Year2006 | -0.1869 | 0.0321 | -5.8304 | 0.000000 |
| Year2007 | -0.0790 | 0.0342 | -2.3074 | 0.021054 |
| Year2008 | -0.1017 | 0.0374 | -2.7179 | 0.006583 |
| Year2009 | -0.2922 | 0.0428 | -6.8238 | 0.000000 |
| Year2010 | -0.2594 | 0.0418 | -6.2126 | 0.000000 |
| Year2011 | 0.1061 | 0.0430 | 2.4686 | 0.013585 |
| Year2012 | 0.0664 | 0.0406 | 1.6357 | 0.101936 |
| Year2013 | 0.0388 | 0.0393 | 0.9876 | 0.323359 |
| Year2014 | 0.3154 | 0.0432 | 7.3038 | 0.000000 |
| Year2015 | 0.2576 | 0.0406 | 6.3436 | 0.000000 |
| Year2016 | 0.0844 | 0.0381 | 2.2144 | 0.026828 |
| Year2017 | 0.0360 | 0.0395 | 0.9120 | 0.361803 |
| Year2018 | 0.2359 | 0.0446 | 5.2942 | 0.000000 |
| Year2019 | 0.1723 | 0.0400 | 4.3115 | 0.000016 |
| Year2020 | 0.0803 | 0.0411 | 1.9514 | 0.051039 |
| Captain112 | 0.3999 | 0.0742 | 5.3911 | 0.000000 |
| Captain133 | 0.1510 | 0.0271 | 5.5802 | 0.000000 |
| Captain155 | -0.5919 | 0.1472 | -4.0194 | 0.000059 |
| Captain160 | 0.1498 | 0.0320 | 4.6821 | 0.000003 |
| Captain215 | -0.1752 | 0.0592 | -2.9591 | 0.003094 |
| Captain219 | 0.1603 | 0.0244 | 6.5740 | 0.000000 |
| Captain353 | -0.4449 | 0.0594 | -7.4953 | 0.000000 |
| Captain384 | 0.1238 | 0.0448 | 2.7662 | 0.005683 |
| Captain404 | 0.0977 | 0.0401 | 2.4333 | 0.014983 |
| Captain405 | 0.9654 | 0.2730 | 3.5364 | 0.000408 |


| Captain406 | 0.3689 | 0.0865 | 4.2641 | 0.000020 |
| :--- | ---: | ---: | ---: | ---: |
| Captain408 | -0.2410 | 0.0912 | -2.6441 | 0.008206 |
| Gear6 | 0.1370 | 0.0153 | 8.9336 | 0.000000 |
| Gear7 | 0.2824 | 0.0268 | 10.5249 | 0.000000 |
| Gear8 | 0.4400 | 0.1133 | 3.8819 | 0.000104 |
| $\mathrm{~ns}($ SoakDays, df $=3) 1$ | 0.3039 | 0.0341 | 8.9014 | 0.000000 |
| $\mathrm{~ns}($ SoakDays, $\mathrm{df}=3) 2$ | 0.6289 | 0.0685 | 9.1750 | 0.000000 |
| $\mathrm{~ns}($ SoakDays, $\mathrm{df}=3) 3$ | 0.2108 | 0.0413 | 5.1063 | 0.000000 |

Figures A. 7 and A. 8 compare standardized and nominal CPUE indices for pre- and postrationalization periods for EAG:


Figure A.7. Trends in non-standardized (red line) vs. standardized CPUE (black line with confidence intervals) indices during the pre-rationalization period for EAG. The confidence intervals are $+/-2 \mathrm{SE}$.


Figure A.8. Trends in non-standardized (red line) vs. standardized CPUE (black line with confidence intervals) indices during the post-rationalization period for EAG. The confidence intervals are $+/-2$ SE.

The final main effect models for WAG were:

Model 21.1a:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Captain $+\mathrm{ns}($ Soak, 7$)+$ Gear + Area + Month $+\mathrm{ns}($ Depth, 5$)+$ Vessel
AIC=190,897
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain $+\mathrm{ns}($ Soak, 7$)$
for the 1995/96-2004/05 period $\left[\theta=0.97, \mathrm{R}^{2}=0.1425, \mathrm{AIC}=146,246\right]$
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Gear + Vessel + Month $+\mathrm{ns}($ Soak, 2$)+\mathrm{ns}($ Depth, 9$)$
AIC=117,799
Final selection by stepCPUE:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Year }+ \text { Gear }+\mathrm{ns}(\text { Soak, } 2) \tag{A.7}
\end{equation*}
$$

for the 2005/06-2020/21 period $\left[\theta=1.13, R^{2}=0.0482, \operatorname{AIC}=117,673\right.$, Soak forced in $]$.

Tables A. 4 and A. 5 list the summary statistics of the main effects GLM fits to 1995/96-2004/05 and 2005/06-2020/21 data series, respectively:

Table A.4. Summary statistics of the main effects GLM fit to WAG 1995/96-2004/05 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| (Intercept) | 1.7843 | 0.0586 | 30.4682 | 0.000000 |
| Year1996 | -0.0731 | 0.0385 | -1.9003 | 0.057406 |
| Year1997 | -0.0112 | 0.0431 | -0.2610 | 0.794117 |
| Year1998 | 0.0660 | 0.0501 | 1.3183 | 0.187422 |
| Year1999 | -0.0851 | 0.0434 | -1.9612 | 0.049864 |
| Year2000 | -0.2193 | 0.0434 | -5.0531 | 0.000000 |
| Year2001 | -0.1516 | 0.0468 | -3.2378 | 0.001206 |
| Year2002 | -0.0308 | 0.0488 | -0.6304 | 0.528410 |
| Year2003 | 0.2503 | 0.0496 | 5.0429 | 0.000000 |
| Year2004 | 0.2667 | 0.0498 | 5.3561 | 0.000000 |
| Captain105 | -0.8853 | 0.1828 | -4.8431 | 0.000001 |
| Captain108 | -0.6845 | 0.0791 | -8.6524 | 0.000000 |
| Captain112 | -1.0828 | 0.1644 | -6.5853 | 0.000000 |
| Captain114 | -0.5768 | 0.0815 | -7.0759 | 0.000000 |
| Captain128 | -0.6929 | 0.0510 | -13.5875 | 0.000000 |
| Captain130 | 0.2862 | 0.0450 | 6.3524 | 0.000000 |
| Captain131 | 0.2267 | 0.0658 | 3.4462 | 0.000569 |
| Captain133 | -0.6598 | 0.0632 | -10.4336 | 0.000000 |
| Captain145 | -0.2517 | 0.0440 | -5.7154 | 0.000000 |
| Captain156 | -1.2927 | 0.2236 | -5.7819 | 0.000000 |
| Captain157 | -0.4649 | 0.0640 | -7.2670 | 0.000000 |
| Captain159 | -0.5926 | 0.2764 | -2.1444 | 0.032010 |
| Captain160 | -0.4679 | 0.0458 | -10.2232 | 0.000000 |
| Captain182 | -0.7465 | 0.1053 | -7.0874 | 0.000000 |
| Captain188 | -0.4973 | 0.0849 | -5.8597 | 0.000000 |
| Captain201 | -0.6832 | 0.0674 | -10.1304 | 0.000000 |
| Captain210 | -0.6867 | 0.1217 | -5.6406 | 0.000000 |
| Captain219 | -0.3789 | 0.0441 | -8.6016 | 0.000000 |
| Captain230 | -1.0871 | 0.0747 | -14.5505 | 0.000000 |
| Captain232 | -1.7885 | 0.2207 | -8.1020 | 0.000000 |
| Captain235 | -1.1987 | 0.0754 | -15.8991 | 0.000000 |
| Captain244 | -0.3246 | 0.1004 | -3.2337 | 0.001224 |
| Captain271 | -0.8739 | 0.0834 | -10.4819 | 0.000000 |
| Captain272 | -0.6511 | 0.0702 | -9.2783 | 0.000000 |
| Captain277 | -0.4044 | 0.0792 | -5.1081 | 0.000000 |
| Captain287 | -0.7559 | 0.1049 | -7.2091 | 0.000000 |
|  |  |  |  |  |


| Captain302 | -0.6210 | 0.0827 | -7.5112 | 0.000000 |
| :--- | ---: | ---: | ---: | ---: |
| Captain304 | -0.6541 | 0.0731 | -8.9477 | 0.000000 |
| Captain315 | -0.8606 | 0.0863 | -9.9698 | 0.000000 |
| Captain318 | -0.9753 | 0.0780 | -12.5035 | 0.000000 |
| Captain322 | -1.4140 | 0.2408 | -5.8730 | 0.000000 |
| Captain326 | -0.4077 | 0.0497 | -8.1957 | 0.000000 |
| Captain328 | -0.7078 | 0.1657 | -4.2715 | 0.000019 |
| Captain332 | -0.8156 | 0.0761 | -10.7157 | 0.000000 |
| Captain335 | -0.4667 | 0.1825 | -2.5572 | 0.010557 |
| Captain345 | -0.7582 | 0.1773 | -4.2770 | 0.000019 |
| Captain359 | -0.2962 | 0.0809 | -3.6587 | 0.000254 |
| Captain363 | -0.6023 | 0.0995 | -6.0528 | 0.000000 |
| Captain369 | -0.8612 | 0.1106 | -7.7869 | 0.000000 |
| Captain384 | -0.3961 | 0.0792 | -5.0025 | 0.000001 |
| Captain387 | -1.1069 | 0.1218 | -9.0847 | 0.000000 |
| Captain389 | -0.6683 | 0.0747 | -8.9507 | 0.000000 |
| Captain390 | -0.3131 | 0.0974 | -3.2147 | 0.001308 |
| Captain392 | -1.4343 | 0.2788 | -5.1447 | 0.000000 |
| ns(SoakDays, $\mathrm{df}=7) 1$ | 0.1243 | 0.0495 | 2.5105 | 0.012063 |
| ns(SoakDays, $\mathrm{df}=7) 2$ | 0.3426 | 0.0606 | 5.6567 | 0.000000 |
| ns(SoakDays, $\mathrm{df}=7) 3$ | 0.4550 | 0.0549 | 8.2929 | 0.000000 |
| ns(SoakDays, $\mathrm{df}=7) 4$ | 0.7380 | 0.0493 | 14.9663 | 0.000000 |
| ns(SoakDays, df $=7) 5$ | 0.6957 | 0.0564 | 12.3329 | 0.000000 |
| ns(SoakDays, $\mathrm{df}=7) 6$ | 0.7966 | 0.0803 | 9.9170 | 0.000000 |
| ns(SoakDays, df $=7) 7$ | 0.7097 | 0.0683 | 10.3906 | 0.000000 |

Table A.5. Summary statistics of the main effects GLM fit to WAG 2005/06-2020/21 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| Intercept) | 2.7409 | 0.0398 | 68.9145 | 0.000000 |
| Year2006 | -0.0208 | 0.0402 | -0.5176 | 0.604752 |
| Year2007 | -0.1609 | 0.0394 | -4.0800 | 0.000045 |
| Year2008 | -0.0053 | 0.0403 | -0.1304 | 0.896249 |
| Year2009 | 0.0520 | 0.0419 | 1.2410 | 0.214632 |
| Year2010 | -0.0894 | 0.0422 | -2.1171 | 0.034271 |
| Year2011 | -0.0581 | 0.0438 | -1.3250 | 0.185180 |
| Year2012 | -0.0756 | 0.0398 | -1.8998 | 0.057474 |
| Year2013 | -0.3625 | 0.0392 | -9.2498 | 0.000000 |
| Year2014 | -0.4705 | 0.0396 | -11.8764 | 0.000000 |
| Year2015 | -0.4504 | 0.0384 | -11.7403 | 0.000000 |
| Year2016 | -0.3174 | 0.0411 | -7.7137 | 0.000000 |
| Year2017 | -0.1699 | 0.0439 | -3.8696 | 0.000109 |


| Year2018 | 0.0263 | 0.0459 | 0.5731 | 0.566572 |
| :--- | ---: | ---: | ---: | ---: |
| Year2019 | -0.1855 | 0.0425 | -4.3653 | 0.000013 |
| Year2020 | -0.3118 | 0.0412 | -7.5618 | 0.000000 |
| Gear6 | 0.3004 | 0.0232 | 12.9346 | 0.000000 |
| Gear7 | 0.3534 | 0.0272 | 13.0020 | 0.000000 |
| Gear8 | 0.5976 | 0.0341 | 17.5139 | 0.000000 |
| Gear13 | 1.0039 | 0.1840 | 5.4563 | 0.000000 |
| Gear25 | 1.4570 | 0.3378 | 4.3130 | 0.000016 |
| ns(SoakDays, df $=2) 1$ | 0.3401 | 0.0561 | 6.0654 | 0.000000 |
| ns(SoakDays, df $=2) 2$ | 0.1445 | 0.0351 | 4.1203 | 0.000038 |

Figures A. 9 and A. 10 compare standardized and nominal CPUE indices for pre- and postrationalization periods for WAG:


Figure A.9. Trends in non-standardized (red line) vs. standardized CPUE (black line with confidence intervals) indices during the pre-rationalization period for WAG. The confidence intervals are $+/-2 \mathrm{SE}$.


Figure A.10. Trends in non-standardized (red line) vs. standardized CPUE (black line with confidence intervals) indices during the post-rationalization period for WAG. The confidence intervals are $+/-2 \mathrm{SE}$.
b. Year:Area interaction effects GLM:

For year and area interaction analysis, we designed the areas in to $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids enmeshed in 10 larger blocks as follows. The number of blocks was restricted to a few to prevent GLM fitting problems (Figure A. 11 and Table A.6).


Figure A.11. The 1995/96-2020/21 observer pot samples enmeshed in 10 blocks for the Aleutian Islands golden king crab. The blocks were determined from visually exploring each year's pot distribution locations (available with the first author). The blocks contain observed patches of crab distribution during this period.

Table A.6. Number of $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids containing observer sample locations within each block by fishing year for the Aleutian Islands golden king crab, 1995/96-2020/21 data. Blocks 1-4 belong to EAG and 5-10 to WAG. Sum of ever fished number of grids for each block is listed at the bottom row.

| FMP Season | Block 1 | Block 2 | Block 3 | Block 4 | Block_5 | Block 6 | Block 7 | Block_8 | Block 9 | Block_10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | 125 | 529 | 748 | 379 | 218 | 373 | 112 | 722 | 166 | 122 |
| 1996 | 149 | 814 | 761 | 372 | 89 | 473 | 359 | 799 | 200 | 35 |
| 1997 | 116 | 530 | 755 | 257 | 202 | 443 | 104 | 568 | 274 | 0 |
| 1998 | 78 | 581 | 453 | 236 | 18 | 318 | 157 | 251 | 132 | 0 |
| 1999 | 123 | 593 | 454 | 231 | 163 | 476 | 182 | 627 | 193 | 145 |
| 2000 | 72 | 540 | 754 | 301 | 187 | 440 | 195 | 555 | 547 | 47 |
| 2001 | 123 | 507 | 507 | 329 | 45 | 369 | 288 | 634 | 256 | 9 |
| 2002 | 97 | 387 | 584 | 271 | 71 | 341 | 205 | 335 | 242 | 37 |
| 2003 | 43 | 492 | 530 | 299 | 111 | 347 | 212 | 465 | 150 | 61 |
| 2004 | 81 | 289 | 377 | 216 | 77 | 319 | 150 | 359 | 172 | 116 |
| 2005 | 0 | 205 | 221 | 118 | 8 | 220 | 83 | 261 | 54 | 0 |
| 2006 | 0 | 154 | 248 | 122 | 15 | 191 | 58 | 220 | 39 | 0 |
| 2007 | 0 | 111 | 177 | 110 | 24 | 228 | 78 | 173 | 20 | 0 |
| 2008 | 0 | 111 | 203 | 93 | 12 | 181 | 67 | 196 | 0 | 0 |
| 2009 | 0 | 59 | 146 | 60 | 6 | 137 | 95 | 220 | 25 | 0 |
| 2010 | 0 | 81 | 141 | 85 | 1 | 115 | 73 | 260 | 39 | 0 |
| 2011 | 0 | 126 | 117 | 33 | 3 | 83 | 73 | 266 | 9 | 0 |
| 2012 | 0 | 146 | 110 | 56 | 7 | 91 | 85 | 312 | 53 | 0 |
| 2013 | 2 | 149 | 129 | 51 | 12 | 144 | 105 | 293 | 86 | 0 |
| 2014 | 1 | 138 | 96 | 41 | 39 | 120 | 114 | 319 | 37 | 0 |
| 2015 | 0 | 135 | 147 | 61 | 46 | 163 | 106 | 280 | 16 | 48 |
| 2016 | 0 | 145 | 231 | 63 | 26 | 134 | 89 | 210 | 106 | 0 |
| 2017 | 0 | 97 | 170 | 110 | 11 | 87 | 79 | 198 | 118 | 0 |
| 2018 | 0 | 91 | 158 | 95 | 7 | 69 | 82 | 204 | 121 | 0 |
| 2019 | 1 | 112 | 171 | 101 | 0 | 0 | 89 | 316 | 138 | 0 |
| 2020 | 4 | 109 | 193 | 95 | 0 | 0 | 76 | 287 | 91 | 36 |


| Ever Fished: |  |  |  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AIGKC All Seasons | Block_1 | Block_2 | Block_3 | Block_4 | Block_5 | Block_6 | Block_7 | Block_8 | Block_9 | Block_10 |
| 1995-2020 - Sum of 1x1 cells | 381 | 1402 | 1792 | 917 | 459 | 1028 | 796 | 2012 | 1021 | 334 |

We assumed the null model to be

$$
\begin{equation*}
\ln \left(\mathrm{CPUE}_{\mathrm{i}}\right)=\text { Year }_{\mathrm{y}_{\mathrm{i}}}: \text { Area }_{\mathrm{ai}} \tag{A.8}
\end{equation*}
$$

The maximum set of model terms offered to the stepwise selection procedure was:
$\ln \left(\right.$ CPUE $\left._{\mathrm{I}}\right)=$ Year $_{\mathrm{yi}_{\mathrm{i}}}:$ Area $_{a i}+\mathrm{ns}\left(\right.$ Soak $_{\text {si }}$, df $)+$ Month $_{\mathrm{m}_{\mathrm{i}}}+$ Vessel $_{\mathrm{vi}}+$ Captain $_{\text {ci }}+$ $\mathrm{Area}_{\mathrm{ai}}+$ Gear $_{\mathrm{gi}}+\mathrm{ns}\left(\right.$ Depth $\left._{\mathrm{di}}, \mathrm{df}\right)$.

Example R codes used for interaction effects GLM fitting are as follows:
For WAG 1995_04 CPUE indices:
library(MASS)
library(splines)
Step 1:
glm.object<- glm(Legals $\sim$ Year:Area,family $=$ negative.binomial(0.97),data=datacore)
wpotsampleoutAIC<-stepAIC(glm.object,scope=list(upper=
$\sim($ Year:Area + ns(SoakDays, $\mathrm{df}=7$ ) + Month+Vessel+Captain+Area+Gear +
ns(Depth,df=5)), lower=~Year:Area),family=
negative.binomial(0.97),direction="forward",trace $=9, \mathrm{k}=\log ($ nrow(datacore) $)+1.0)$
Step 2:
glm.object<- glm(Legals~Year:Area,family = negative.binomial(0.97),data=datacore)
wpotsampleout<-stepCPUE(glm.object,scope=list(upper=
$\sim($ Vessel + ns(SoakDays,df=7)+Gear+Month + ns(Depth,df=5)+Year:Area),lower=
$\sim$ Year:Area),family= negative.binomial(0.97),direction="forward",trace $=9$, ,2.change $=0.01$ )
The final interaction effects models for EAG were:
Model 21.1b:

Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Gear + Captain $+\mathrm{ns}($ Soak, 4$)+$ Month + Year: Area
AIC=203,851
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Captain + ns(Soak, 4$)+$ Year: Area
for the 1995/96-2004/05 period $\left[\theta=1.38, \mathrm{R}^{2}=0.2060\right.$, $\left.\mathrm{AIC}=170,920\right]$,

Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Vessel + Gear $+\mathrm{ns}($ Soak, 3$)+$ Month + Year: Area $\mathrm{AIC}=77,473$

Final selection by stepCPUE:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Vessel }+ \text { Year: Area }+\mathrm{ns}(\text { Soak, } 3) \tag{A.11}
\end{equation*}
$$

for the 2005/06-2020/21 period $\left[\theta=2.32, \mathrm{R}^{2}=0.1047\right.$, $\mathrm{AIC}=46,455$, Soak forced in $]$.

Tables A. 6 and A. 7 list the summary statistics of the interaction effects GLM fits to 1995/962004/05 and 2005/06-2020/21 data series, respectively:

Table A.6. Summary statistics of the interaction effects GLM fit to EAG 1995/96-2004/05 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}( \rangle\|\mathrm{t}\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| (Intercept) | 2.5926 | 0.0938 | 27.6541 | 0.000000 |
| Captain105 | 0.1763 | 0.0302 | 5.8416 | 0.000000 |
| Captain108 | 0.2694 | 0.0770 | 3.4977 | 0.000470 |
| Captain128 | 0.1930 | 0.0383 | 5.0374 | 0.000000 |
| Captain133 | 0.4112 | 0.0323 | 12.7236 | 0.000000 |
| Captain145 | 0.2995 | 0.1119 | 2.6774 | 0.007425 |
| Captain160 | 0.1895 | 0.0452 | 4.1907 | 0.000028 |
| Captain161 | 0.1706 | 0.0677 | 2.5194 | 0.011761 |
| Captain204 | -0.4792 | 0.0758 | -6.3175 | 0.000000 |
| Captain208 | 0.2529 | 0.1007 | 2.5115 | 0.012029 |
| Captain210 | -0.0976 | 0.0367 | -2.6588 | 0.007848 |
| Captain213 | 0.5125 | 0.0312 | 16.4285 | 0.000000 |
| Captain219 | 0.3765 | 0.0352 | 10.7078 | 0.000000 |
| Captain232 | -0.3570 | 0.0570 | -6.2662 | 0.000000 |
| Captain233 | -0.1997 | 0.0860 | -2.3231 | 0.020184 |
| Captain240 | -0.4844 | 0.1687 | -2.8720 | 0.004083 |
| Captain247 | 0.1478 | 0.0274 | 5.3913 | 0.000000 |
| Captain271 | -0.4623 | 0.0482 | -9.5888 | 0.000000 |
| Captain272 | 0.2739 | 0.0320 | 8.5506 | 0.000000 |
| Captain276 | 0.2376 | 0.0572 | 4.1515 | 0.000033 |
| Captain287 | -0.2346 | 0.0834 | -2.8116 | 0.004934 |
| Captain300 | -0.2967 | 0.0867 | -3.4227 | 0.000621 |
| Captain332 | 0.5175 | 0.0748 | 6.9136 | 0.000000 |
| Captain344 | -1.1827 | 0.2230 | -5.3026 | 0.000000 |
| Captain353 | 0.3019 | 0.0964 | 3.1310 | 0.001744 |
| Captain354 | 0.3680 | 0.0558 | 6.5903 | 0.000000 |


| Captain358 | -0.8059 | 0.2254 | -3.5750 | 0.000351 |
| :--- | ---: | ---: | ---: | ---: |
| Captain384 | 0.4971 | 0.0284 | 17.4873 | 0.000000 |
| Captain388 | -0.6976 | 0.1066 | -6.5436 | 0.000000 |
| Captain390 | 1.0182 | 0.0623 | 16.3508 | 0.000000 |
| Captain392 | 0.1242 | 0.0544 | 2.2846 | 0.022343 |
| ns(SoakDays, df $=4) 1$ | 0.3006 | 0.0300 | 10.0227 | 0.000000 |
| ns(SoakDays, df = 4)2 | 0.4499 | 0.0322 | 13.9788 | 0.000000 |
| ns(SoakDays, df = 4)3 | 0.7053 | 0.0528 | 13.3668 | 0.000000 |
| ns(SoakDays, df = 4)4 | 0.4991 | 0.0427 | 11.6981 | 0.000000 |
| Year1995:Block1 | -1.6666 | 0.1879 | -8.8720 | 0.000000 |
| Year1996:Block1 | -0.7086 | 0.2230 | -3.1772 | 0.001489 |
| Year1997:Block1 | -1.6974 | 0.1448 | -11.7221 | 0.000000 |
| Year1998:Block1 | -1.7624 | 0.1669 | -10.5581 | 0.000000 |
| Year1999:Block1 | -1.7888 | 0.1022 | -17.4965 | 0.000000 |
| Year2000:Block1 | -2.5934 | 0.4414 | -5.8753 | 0.000000 |
| Year2002:Block1 | -0.6584 | 0.2636 | -2.4977 | 0.012506 |
| Year1995:Block2 | -0.9092 | 0.0993 | -9.1555 | 0.000000 |
| Year1996:Block2 | -1.0318 | 0.0899 | -11.4719 | 0.000000 |
| Year1997:Block2 | -0.7332 | 0.0913 | -8.0322 | 0.000000 |
| Year1998:Block2 | -0.8568 | 0.0924 | -9.2727 | 0.000000 |
| Year1999:Block2 | -0.7493 | 0.0913 | -8.2034 | 0.000000 |
| Year2000:Block2 | -0.9120 | 0.0916 | -9.9538 | 0.000000 |
| Year2001:Block2 | -0.5395 | 0.0932 | -5.7908 | 0.000000 |
| Year2002:Block2 | -0.4806 | 0.0943 | -5.0976 | 0.000000 |
| Year2003:Block2 | -0.6272 | 0.0927 | -6.7669 | 0.000000 |
| Year2004:Block2 | -0.2343 | 0.0978 | -2.3952 | 0.016618 |
| Year1995:Block3 | -1.3184 | 0.0951 | -13.8670 | 0.000000 |
| Year1996:Block3 | -1.2736 | 0.0938 | -13.5812 | 0.000000 |
| Year1997:Block3 | -1.0044 | 0.0973 | -10.3213 | 0.000000 |
| Year1998:Block3 | -0.8364 | 0.0949 | -8.8113 | 0.000000 |
| Year1999:Block3 | -0.8690 | 0.0960 | -9.0495 | 0.000000 |
| Year2000:Block3 | -0.8673 | 0.0917 | -9.4581 | 0.000000 |
| Year2001:Block3 | -0.6011 | 0.0929 | -6.4715 | 0.000000 |
| Year2002:Block3 | -0.5602 | 0.0961 | -5.8317 | 0.000000 |
| Year2003:Block3 | -0.6735 | 0.0972 | -6.9293 | 0.000000 |
| Year2004:Block3 | -0.2005 | 0.1057 | -1.8962 | 0.057932 |
| Year1995:Block4 | -1.2762 | 0.0978 | -13.0518 | 0.000000 |
| Year1996:Block4 | -1.7048 | 0.1020 | -16.7190 | 0.000000 |
| Year1997:Block4 | -1.5322 | 0.0994 | -15.4188 | 0.000000 |
| Year1998:Block4 | -0.7937 | 0.1015 | -7.8216 | 0.000000 |
| Year1999:Block4 | -0.8464 | 0.1089 | -7.7675 | 0.000000 |
| Year2000:Block4 | -0.5678 | 0.0904 | -6.2796 | 0.000000 |
| Ye |  |  |  |  |


| Year2001:Block4 | -0.4441 | 0.0937 | -4.7409 | 0.000002 |
| :--- | :--- | :--- | :--- | :--- |
| Year2002:Block4 | -0.6174 | 0.1029 | -6.0017 | 0.000000 |
| Year2003:Block4 | -0.6951 | 0.0978 | -7.1065 | 0.000000 |

Table A.7. Summary statistics of the interaction effects GLM fit to EAG 2005/06-2020/21 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |
| :---: | :---: | :---: | :---: | :---: |
| (Intercept) | 3.2208 | 0.4366 | 7.3766 | 0.000000 |
| Vessel3645 | -0.4678 | 0.0654 | -7.1557 | 0.000000 |
| Vessel8653 | -0.2262 | 0.0614 | -3.6833 | 0.000233 |
| Vessel20556 | 0.0410 | 0.0220 | 1.8601 | 0.062927 |
| Vessel62436 | -0.6778 | 0.0643 | -10.5349 | 0.000000 |
| $\mathrm{ns}($ SoakDays, $\mathrm{df}=3) 1$ | 0.2237 | 0.0449 | 4.9834 | 0.000001 |
| $\mathrm{ns}($ SoakDays, $\mathrm{df}=3) 2$ | 0.6423 | 0.0818 | 7.8537 | 0.000000 |
| $\mathrm{ns}($ SoakDays, $\mathrm{df}=3) 3$ | 0.2361 | 0.0530 | 4.4571 | 0.000008 |
| Year2013:Block1 | -0.1279 | 0.5334 | -0.2398 | 0.810470 |
| Year2019:Block1 | -1.7949 | 0.8255 | -2.1743 | 0.029725 |
| Year2020:Block1 | -0.5415 | 0.5054 | -1.0714 | 0.284041 |
| Year2005:Block2 | -0.0134 | 0.4379 | -0.0305 | 0.975666 |
| Year2006:Block2 | -0.2225 | 0.4388 | -0.5070 | 0.612141 |
| Year2007:Block2 | -0.2192 | 0.4384 | -0.5001 | 0.617031 |
| Year2008: Block2 | -0.3664 | 0.4401 | -0.8327 | 0.405068 |
| Year2009:Block2 | -0.1183 | 0.4605 | -0.2569 | 0.797284 |
| Year2010:Block2 | -0.1414 | 0.4446 | -0.3181 | 0.750422 |
| Year2011:Block2 | 0.1617 | 0.4400 | 0.3676 | 0.713194 |
| Year2012:Block2 | 0.1277 | 0.4391 | 0.2907 | 0.771278 |
| Year2013:Block2 | 0.0668 | 0.4392 | 0.1521 | 0.879141 |
| Year2014:Block2 | 0.1903 | 0.4397 | 0.4327 | 0.665255 |
| Year2015:Block2 | 0.0900 | 0.4390 | 0.2049 | 0.837660 |
| Year2016:Block2 | -0.0277 | 0.4403 | -0.0628 | 0.949890 |
| Year2017:Block2 | -0.1071 | 0.4396 | -0.2435 | 0.807595 |
| Year2018:Block2 | 0.2189 | 0.4395 | 0.4981 | 0.618469 |
| Year2019:Block2 | 0.1784 | 0.4383 | 0.4071 | 0.683970 |
| Year2020:Block2 | -0.0306 | 0.4380 | -0.0699 | 0.944300 |
| Year2005:Block3 | -0.0136 | 0.4393 | -0.0309 | 0.975336 |
| Year2006:Block3 | -0.0854 | 0.4380 | -0.1949 | 0.845489 |
| Year2007:Block3 | -0.1782 | 0.4364 | -0.4083 | 0.683052 |
| Year2008:Block3 | -0.2085 | 0.4378 | -0.4762 | 0.633955 |
| Year2009:Block3 | -0.3800 | 0.4414 | -0.8609 | 0.389329 |
| Year2010:Block3 | -0.1719 | 0.4434 | -0.3876 | 0.698323 |
| Year2011:Block3 | 0.0206 | 0.4428 | 0.0464 | 0.962961 |
| Year2012:Block3 | -0.1577 | 0.4425 | -0.3564 | 0.721524 |


| Year2013:Block3 | 0.0818 | 0.4402 | 0.1859 | 0.852513 |
| :--- | ---: | ---: | ---: | ---: |
| Year2014:Block3 | 0.2426 | 0.4421 | 0.5488 | 0.583132 |
| Year2015:Block3 | 0.0734 | 0.4420 | 0.1662 | 0.868026 |
| Year2016:Block3 | 0.1711 | 0.4395 | 0.3893 | 0.697098 |
| Year2017:Block3 | -0.1834 | 0.4411 | -0.4159 | 0.677517 |
| Year2018:Block3 | 0.0341 | 0.4413 | 0.0773 | 0.938363 |
| Year2019:Block3 | 0.2339 | 0.4390 | 0.5328 | 0.594197 |
| Year2020:Block3 | -0.0439 | 0.4407 | -0.0995 | 0.920732 |
| Year2005:Block4 | 0.2012 | 0.4379 | 0.4595 | 0.645879 |
| Year2006:Block4 | -0.1929 | 0.4372 | -0.4413 | 0.659011 |
| Year2007:Block4 | -0.0009 | 0.4367 | -0.0020 | 0.998387 |
| Year2008:Block4 | -0.2204 | 0.4389 | -0.5021 | 0.615607 |
| Year2009:Block4 | -0.3157 | 0.4407 | -0.7164 | 0.473752 |
| Year2010:Block4 | -0.4546 | 0.4413 | -1.0301 | 0.302999 |
| Year2011:Block4 | -0.0189 | 0.4455 | -0.0423 | 0.966238 |
| Year2012:Block4 | -0.0290 | 0.4415 | -0.0657 | 0.947588 |
| Year2013:Block4 | -0.0547 | 0.4404 | -0.1242 | 0.901158 |
| Year2014:Block4 | 0.1934 | 0.4426 | 0.4369 | 0.662205 |
| Year2015:Block4 | 0.0861 | 0.4405 | 0.1955 | 0.845004 |
| Year2016:Block4 | 0.2352 | 0.4412 | 0.5332 | 0.593951 |
| Year2017:Block4 | -0.0031 | 0.4506 | -0.0068 | 0.994545 |

The final interaction effects models for WAG were:
Model 21.1b:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Vessel + ns(Soak, 7$)+$ Gear + Month $+\mathrm{ns}($ Depth, 5$)+$ Year: Area
AIC=191,018
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Vessel + ns(Soak, 7) + Year: Area
for the 1995/96-2004/05 period $\left[\theta=0.97, \mathrm{R}^{2}=0.1657\right.$, $\left.\mathrm{AIC}=147,887\right]$
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Gear + Vessel + Month $+n s($ Soak, 2$)+n s($ Depth,$d f=9)+$ Year: Area AIC=120,656

Final selection by stepCPUE:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Gear }+ \text { Year }: \text { Area }+n s(\text { Soak, } 2) \tag{A.13}
\end{equation*}
$$

for the $2005 / 06-2020 / 21$ period $\left[\theta=1.13, R^{2}=0.0862, \operatorname{AIC}=76,797\right.$, Soak forced in $]$.
Tables A. 8 and A. 9 list the summary statistics of the interaction effects GLM fits to 1995/962004/05 and 2005/06-2020/21 data series, respectively:

Table A.8. Summary statistics of the interaction effects GLM fit to WAG 1995/96-2004/05 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| (Intercept) | 0.9199 | 0.1981 | 4.6434 | 0.000003 |
| Vessel5992 | -0.3046 | 0.0875 | -3.4794 | 0.000503 |
| Vessel6205 | 0.1216 | 0.0514 | 2.3682 | 0.017885 |
| Vessel8653 | 0.2662 | 0.0470 | 5.6674 | 0.000000 |
| Vessel9069 | -0.4239 | 0.0568 | -7.4616 | 0.000000 |
| Vessel21436 | -3.1442 | 1.1084 | -2.8368 | 0.004561 |
| Vessel35767 | 0.3836 | 0.0417 | 9.1944 | 0.000000 |
| Vessel37887 | 0.7365 | 0.0412 | 17.8893 | 0.000000 |
| Vessel39002 | 0.2081 | 0.0441 | 4.7224 | 0.000002 |
| Vessel55124 | -0.2914 | 0.0786 | -3.7074 | 0.000210 |
| Vessel59521 | -0.5079 | 0.2203 | -2.3057 | 0.021138 |
| Vessel62436 | -0.6407 | 0.0600 | -10.6766 | 0.000000 |
| ns(SoakDays, df $=7) 1$ | 0.2492 | 0.0561 | 4.4416 | 0.000009 |
| ns(SoakDays, df $=7) 2$ | 0.5629 | 0.0624 | 9.0225 | 0.000000 |
| ns(SoakDays, df $=7) 3$ | 0.7102 | 0.0514 | 13.8103 | 0.000000 |
| ns(SoakDays, df $=7) 4$ | 0.7943 | 0.0512 | 15.5159 | 0.000000 |
| ns(SoakDays, df $=7) 5$ | 0.8458 | 0.0568 | 14.8795 | 0.000000 |
| ns(SoakDays, df $=7) 6$ | 1.0300 | 0.0872 | 11.8133 | 0.000000 |
| ns(SoakDays, df=7)7 | 0.9588 | 0.0661 | 14.5013 | 0.000000 |
| Year1995:Block5 | 0.2784 | 0.2113 | 1.3177 | 0.187616 |
| Year1996:Block5 | 0.4299 | 0.2173 | 1.9782 | 0.047917 |
| Year1997:Block5 | -0.0802 | 0.2022 | -0.3966 | 0.691639 |
| Year1998:Block5 | -0.0851 | 0.3111 | -0.2735 | 0.784457 |
| Year1999:Block5 | -0.3666 | 0.2117 | -1.7321 | 0.083268 |
| Year2000:Block5 | -0.0737 | 0.2067 | -0.3566 | 0.721419 |
| Year2001:Block5 | 0.0936 | 0.2488 | 0.3761 | 0.706842 |
| Year2002:Block5 | 0.0640 | 0.2599 | 0.2461 | 0.805606 |
| Year2003:Block5 | 0.8138 | 0.4471 | 1.8200 | 0.068769 |
| Year2004:Block5 | -0.9071 | 0.3897 | -2.3279 | 0.019925 |
| Year1995:Block6 | 0.4224 | 0.1975 | 2.1391 | 0.032438 |
| Year1996:Block6 | 0.4019 | 0.1960 | 2.0509 | 0.040290 |
| Year1997:Block6 | 0.2397 | 0.1960 | 1.2230 | 0.221324 |
| Year1998:Block6 | 0.2061 | 0.1974 | 1.0440 | 0.296496 |
| Year1999:Block6 | 0.0811 | 0.1950 | 0.4157 | 0.677618 |
|  |  |  |  |  |


| Year2000:Block6 | 0.0073 | 0.1952 | 0.0376 | 0.969968 |
| :--- | ---: | ---: | ---: | ---: |
| Year2001:Block6 | -0.1651 | 0.1971 | -0.8378 | 0.402143 |
| Year2002:Block6 | -0.0752 | 0.1988 | -0.3782 | 0.705263 |
| Year2003:Block6 | 0.0387 | 0.2006 | 0.1928 | 0.847092 |
| Year2004:Block6 | 0.3405 | 0.1981 | 1.7187 | 0.085681 |
| Year1995:Block7 | 0.7610 | 0.2035 | 3.7401 | 0.000184 |
| Year1996:Block7 | 0.4769 | 0.1963 | 2.4288 | 0.015156 |
| Year1997:Block7 | 0.2156 | 0.2033 | 1.0607 | 0.288835 |
| Year1998:Block7 | 0.5827 | 0.2007 | 2.9026 | 0.003704 |
| Year1999:Block7 | 0.2610 | 0.2004 | 1.3022 | 0.192849 |
| Year2000:Block7 | -0.0515 | 0.2116 | -0.2435 | 0.807627 |
| Year2001:Block7 | 0.0720 | 0.1983 | 0.3631 | 0.716530 |
| Year2002:Block7 | 0.0268 | 0.1986 | 0.1350 | 0.892639 |
| Year2003:Block7 | -0.2107 | 0.2019 | -1.0436 | 0.296702 |
| Year2004:Block7 | -0.1077 | 0.2070 | -0.5204 | 0.602790 |
| Year1995:Block8 | 0.4407 | 0.1974 | 2.2324 | 0.025596 |
| Year1996:Block8 | 0.1612 | 0.1950 | 0.8268 | 0.408348 |
| Year1997:Block8 | 0.2143 | 0.1952 | 1.0979 | 0.272247 |
| Year1998:Block8 | 0.2783 | 0.1994 | 1.3961 | 0.162689 |
| Year1999:Block8 | 0.1986 | 0.1989 | 0.9989 | 0.317841 |
| Year2000:Block8 | -0.1799 | 0.2047 | -0.8788 | 0.379534 |
| Year2001:Block8 | -0.1757 | 0.1978 | -0.8884 | 0.374344 |
| Year2002:Block8 | 0.2407 | 0.1984 | 1.2136 | 0.224924 |
| Year2003:Block8 | 0.3854 | 0.1957 | 1.9691 | 0.048954 |
| Year2004:Block8 | 0.5986 | 0.1971 | 3.0376 | 0.002387 |
| Year1996:Block9 | 0.2707 | 0.2259 | 1.1984 | 0.230789 |
| Year1997:Block9 | 0.6695 | 0.2109 | 3.1752 | 0.001499 |
| Year1998:Block9 | 0.3100 | 0.2117 | 1.4641 | 0.143172 |
| Year1999:Block9 | 0.4255 | 0.2248 | 1.8931 | 0.058351 |
| Year2000:Block9 | 0.6374 | 0.2059 | 3.0957 | 0.001966 |
| Year2001:Block9 | -0.0225 | 0.2331 | -0.0965 | 0.923099 |
| Year2002:Block9 | 0.3032 | 0.2327 | 1.3029 | 0.192620 |
| Year2003:Block9 | 0.7034 | 0.2036 | 3.4540 | 0.000553 |
| Year2004:Block9 | 0.5940 | 0.2037 | 2.9166 | 0.003542 |
|  |  |  |  |  |

Table A.9. Summary statistics of the interaction effects GLM fit to WAG 2005/06-2020/21 data.

| Parameters | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| :--- | ---: | ---: | :--- | ---: |
| Intercept) | 1.8164 | 0.1584 | 11.4709 | 0.000000 |
| Gear6 | 0.3374 | 0.0318 | 10.6048 | 0.000000 |


| Gear7 | 0.3843 | 0.0355 | 10.8193 | 0.000000 |
| :---: | :---: | :---: | :---: | :---: |
| Gear8 | 0.6964 | 0.0455 | 15.3021 | 0.000000 |
| Gear13 | 1.1981 | 0.2382 | 5.0294 | 0.000001 |
| Gear25 | 1.3304 | 0.3332 | 3.9928 | 0.000066 |
| $\mathrm{ns}($ SoakDays, $\mathrm{df}=2) 1$ | 0.3087 | 0.0766 | 4.0289 | 0.000057 |
| $\mathrm{ns}($ SoakDays, $\mathrm{df}=2) 2$ | 0.0715 | 0.0435 | 1.6425 | 0.100526 |
| Year2006:Block5 | -0.7925 | 0.4248 | -1.8657 | 0.062120 |
| Year2007:Block5 | -0.8394 | 0.3129 | -2.6827 | 0.007315 |
| Year2009:Block5 | 1.9197 | 0.8850 | 2.1691 | 0.030103 |
| Year2014:Block5 | 1.1934 | 0.2715 | 4.3949 | 0.000011 |
| Year2015:Block5 | -0.5197 | 0.2508 | -2.0724 | 0.038257 |
| Year2016:Block5 | 0.1773 | 0.3785 | 0.4683 | 0.639596 |
| Year2017:Block5 | 1.0826 | 0.3501 | 3.0920 | 0.001994 |
| Year2018:Block5 | 1.0906 | 0.4253 | 2.5643 | 0.010355 |
| Year2005:Block6 | 1.0915 | 0.1674 | 6.5213 | 0.000000 |
| Year2006:Block6 | 0.8692 | 0.1666 | 5.2171 | 0.000000 |
| Year2007:Block6 | 0.6661 | 0.1667 | 3.9955 | 0.000065 |
| Year2008:Block6 | 0.7828 | 0.1699 | 4.6073 | 0.000004 |
| Year2009:Block6 | 0.5742 | 0.1761 | 3.2617 | 0.001111 |
| Year2010:Block6 | 0.4267 | 0.1757 | 2.4286 | 0.015175 |
| Year2011:Block6 | 0.9623 | 0.1776 | 5.4190 | 0.000000 |
| Year2012:Block6 | 0.9363 | 0.1755 | 5.3345 | 0.000000 |
| Year2013:Block6 | 0.7252 | 0.1751 | 4.1405 | 0.000035 |
| Year2014:Block6 | 0.5758 | 0.1778 | 3.2377 | 0.001209 |
| Year2015:Block6 | 0.5593 | 0.1735 | 3.2237 | 0.001270 |
| Year2016:Block6 | 0.4234 | 0.1787 | 2.3690 | 0.017856 |
| Year2017:Block6 | 0.9864 | 0.1807 | 5.4574 | 0.000000 |
| Year2018:Block6 | 1.2769 | 0.1817 | 7.0266 | 0.000000 |
| Year2005:Block7 | 0.8233 | 0.1855 | 4.4388 | 0.000009 |
| Year2006:Block7 | 0.7355 | 0.1988 | 3.6996 | 0.000217 |
| Year2007:Block7 | 0.8960 | 0.1817 | 4.9324 | 0.000001 |
| Year2008:Block7 | 0.9549 | 0.1831 | 5.2159 | 0.000000 |
| Year2009:Block7 | 1.1074 | 0.1805 | 6.1338 | 0.000000 |
| Year2010:Block7 | 0.8976 | 0.1824 | 4.9223 | 0.000001 |
| Year2011:Block7 | 0.8602 | 0.1862 | 4.6196 | 0.000004 |
| Year2012:Block7 | 1.1296 | 0.1788 | 6.3161 | 0.000000 |
| Year2013:Block7 | 0.7104 | 0.1845 | 3.8497 | 0.000119 |
| Year2014:Block7 | 0.6683 | 0.1808 | 3.6963 | 0.000220 |
| Year2015:Block7 | 0.3638 | 0.1774 | 2.0503 | 0.040360 |
| Year2016:Block7 | 0.5057 | 0.1793 | 2.8207 | 0.004802 |
| Year2017:Block7 | 0.5153 | 0.1862 | 2.7677 | 0.005656 |


| Year2018:Block7 | 0.9848 | 0.1838 | 5.3582 | 0.000000 |
| :--- | ---: | :--- | :--- | :--- |
| Year2019:Block7 | 0.7804 | 0.1796 | 4.3449 | 0.000014 |
| Year2020:Block7 | 0.7788 | 0.1792 | 4.3464 | 0.000014 |
| Year2005:Block8 | 0.8919 | 0.1675 | 5.3254 | 0.000000 |
| Year2006:Block8 | 0.8626 | 0.1712 | 5.0388 | 0.000000 |
| Year2007:Block8 | 0.9372 | 0.1744 | 5.3746 | 0.000000 |
| Year2008:Block8 | 1.0895 | 0.1726 | 6.3128 | 0.000000 |
| Year2009:Block8 | 1.1043 | 0.1712 | 6.4488 | 0.000000 |
| Year2010:Block8 | 0.9542 | 0.1684 | 5.6674 | 0.000000 |
| Year2011:Block8 | 0.9308 | 0.1702 | 5.4698 | 0.000000 |
| Year2012:Block8 | 0.7119 | 0.1677 | 4.2446 | 0.000022 |
| Year2013:Block8 | 0.4717 | 0.1682 | 2.8045 | 0.005049 |
| Year2014:Block8 | 0.4053 | 0.1673 | 2.4223 | 0.015443 |
| Year2015:Block8 | 0.5327 | 0.1688 | 3.1556 | 0.001607 |
| Year2016:Block8 | 0.6549 | 0.1724 | 3.7980 | 0.000147 |
| Year2017:Block8 | 0.5736 | 0.1712 | 3.3511 | 0.000808 |
| Year2018:Block8 | 0.7707 | 0.1757 | 4.3852 | 0.000012 |
| Year2019:Block8 | 0.7001 | 0.1662 | 4.2118 | 0.000026 |
| Year2020:Block8 | 0.7245 | 0.1656 | 4.3749 | 0.000012 |
| Year2005:Block9 | 0.7498 | 0.2241 | 3.3462 | 0.000822 |
| Year2006:Block9 | 1.1511 | 0.2109 | 5.4587 | 0.000000 |
| Year2007:Block9 | 0.8563 | 0.2683 | 3.1916 | 0.001419 |
| Year2010:Block9 | 1.0404 | 0.5366 | 1.9390 | 0.052534 |
| Year2011:Block9 | 0.9847 | 0.6434 | 1.5304 | 0.125962 |
| Year2012:Block9 | 0.9403 | 0.2122 | 4.4305 | 0.000010 |
| Year2013:Block9 | 0.4727 | 0.2232 | 2.1177 | 0.034228 |
| Year2014:Block9 | 0.3945 | 0.2208 | 1.7866 | 0.074036 |
| Year2015:Block9 | 0.1396 | 0.3096 | 0.4510 | 0.651994 |
| Year2016:Block9 | 0.5725 | 0.1914 | 2.9911 | 0.002787 |
| Year2017:Block9 | 0.9833 | 0.1759 | 5.5915 | 0.000000 |
| Year2018:Block9 | 0.9785 | 0.1730 | 5.6559 | 0.000000 |
| Year2019:Block9 | 0.4590 | 0.1666 | 2.7552 | 0.005877 |
| Year2020:Block9 | -0.1047 | 0.1769 | -0.5918 | 0.553990 |
| Year2015:Block10 | 0.4726 | 0.2136 | 2.2121 | 0.026986 |
|  |  |  |  |  |

Steps:

1. Block-scale analysis:

The bias corrected estimate of CPUE index for each Year-Area (Area=Block) interaction was first obtained as:
$C P U E_{i j}=e^{Y B_{i j}+\sigma_{i j}^{2} / 2}$
where $C P U E_{i j}$ is the CPUE index in the ith year and jth block, $Y B_{i j}$ is the coefficient of the $i$ th year and $j$ th block interaction, and $\sigma_{i j}$ is the biased correction standard error for expected CPUE value.

The number of $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids in each block can change from year to year; so, we considered using the number of grids ever fished in a block, $N_{\text {ever } j}$ [this is equivalent to assuming that the grids fished in any year randomly sample the stock in that block (Campbell, 2004)].

The abundance index for $j$ th block in $i$ th year is

$$
\begin{equation*}
B_{i j}=N_{\text {ever }_{j}} C P U E_{i j} \tag{A.15}
\end{equation*}
$$

Notice in Table A. 6 that none or very few observer samplings occurred in certain years for a whole block. We filled the $B_{i j}$ index gaps resulting from Year:Area CPUE standardization model fit as follows:

$$
\begin{equation*}
\widehat{B_{l, j}}=e^{A_{i}+C_{j}} \tag{A.16}
\end{equation*}
$$

fitted by GLM [i.e., fitting a log-linear model, $\ln \left(\widehat{B}_{i, j}\right)=A_{i}+C_{j}$ ], where $B_{i, j}$ is the available index of biomass for year i and block $j, A_{i}$ is a year factor, and $C_{j}$ is a block factor, and used this model to predict the unavailable biomass index for blocks x years with no (or very limited) data.

An example set of R codes used to predict the missing biomass index is as follows:

## library(MASS)

To fit the log-linear model (Equation A.16):
glm.fit<- $\operatorname{glm}\left(\log \left(\mathrm{B}_{\mathrm{ij}}\right) \sim\right.$ Year $_{\mathrm{i}}+$ Block $_{\mathbf{j}}$, data=Bindex $)$
where the data frame "Bindex" contains available $\mathrm{B}_{\mathrm{ij}}$, Year $_{\mathrm{i}}$, and Block $\mathrm{k}_{\mathrm{j}}$ column values.
To predict the missing biomass index Y :

## Y<- predict.glm (glm.fit, BindexFillpredict, se.fit=TRUE)

where the new data frame "BindexFillpredict" contains Year $_{i}$ and Block $_{j}$ column values for which $\mathrm{B}_{\mathrm{ij}}$ indices are needed and contains an empty $\mathrm{B}_{\mathrm{ij}}$ column for fill in.

By setting se.fit=TRUE, the standard errors, $\sigma_{i j}$, of predictions are also estimated.

Bias correction was made to each predicted biomass index by $B_{i, j}=e^{Y_{i, j}+\sigma_{i j}^{2} / 2}$ where $\sigma_{i j}$ is the standard error of predicted $Y_{i, j}$ value, which is on the scale of the linear predictor (i.e., $\log$ transformed $\mathrm{B}_{\mathrm{ij}}$. The standard error for each year and area combination is estimated as follows.

If we denote the covariance matrix of the fitted "glm.fit" as $\Sigma$ and write the coefficients for linear combination of a set of predictors in a vector form as $C$, then the standard error of prediction for that combination is $\sqrt{\mathrm{C}^{\prime} \Sigma \mathrm{C}}$, where $C^{\prime}$ is the transpose of vector $C$.

Annual biomass index, $B_{i}$, was estimated as,
$B_{i}=\sum_{j} B_{i j}$

The variance of the total biomass index was computed as:
$\operatorname{Var}\left(B_{i}\right)=\sum_{j} N_{\text {ever }, j}{ }^{2} \operatorname{var}\left(\right.$ CPUE $\left._{i, j}\right)$
where $\boldsymbol{N}_{\text {ever }, j}$ is the total number of 1 mnix 1 mni cells ever fished in block $j$, and $C P U E_{i, j}$ is the CPUE index for year $i$ and block $j$.

To use in the assessment model 21.1b, we rescaled the $B_{i}$ indices by the geometric mean of estimated $B_{i}$ values (Equation A.17) separately for the pre- and post-rationalization periods. The corresponding standard error $(\sim \mathrm{CV})$ of $B_{i}$ was estimated by

$$
\begin{equation*}
\sqrt{\frac{\operatorname{Var}\left(B_{i}\right)}{\left(B_{i}\right)^{2}}} \tag{A.19}
\end{equation*}
$$

The rescaled biomass indices with standard errors are listed in Table A. 10 for EAG and Table A. 11 for WAG.

Table A.10. Steps to estimate biomass-based abundance indices with standard errors for 1995/96-2020/21 in EAG. GMScaled B_index and B_Index SE were used as CPUE index and its standard error.

| Year | B_Index | GMScaled B_Index | Var(B_index) | Var(B_Index)/(B_Index) ${ }^{2}$ | B_Index SE |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1995 | 1379.633 | 0.691 | 61573.282 | 0.032 | 0.180 |
| 1996 | 1365.166 | 0.684 | 60106.142 | 0.032 | 0.180 |
| 1997 | 1605.361 | 0.804 | 58137.446 | 0.023 | 0.150 |
| 1998 | 1860.742 | 0.932 | 58418.764 | 0.017 | 0.130 |
| 1999 | 1880.171 | 0.941 | 57513.803 | 0.016 | 0.128 |
| 2000 | 1874.802 | 0.939 | 78660.880 | 0.022 | 0.150 |
| 2001 | 2528.159 | 1.266 | 62464.339 | 0.010 | 0.099 |
| 2002 | 2600.496 | 1.302 | 66082.978 | 0.010 | 0.099 |
| 2003 | 2243.100 | 1.123 | 65590.905 | 0.013 | 0.114 |
| 2004 | 3452.219 | 1.729 | 144799.925 | 0.012 | 0.110 |


| 2005 | 4917.372 | 1.062 | 1163624.359 | 0.048 | 0.219 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2006 | 4049.173 | 0.875 | 1160862.952 | 0.071 | 0.266 |
| 2007 | 4073.086 | 0.880 | 1155171.590 | 0.070 | 0.264 |
| 2008 | 3636.378 | 0.786 | 1163779.482 | 0.088 | 0.297 |
| 2009 | 3626.917 | 0.784 | 1211346.496 | 0.092 | 0.303 |
| 2010 | 3806.297 | 0.822 | 1189199.573 | 0.082 | 0.287 |
| 2011 | 5040.150 | 1.089 | 1182514.876 | 0.047 | 0.216 |
| 2012 | 4621.799 | 0.998 | 1177368.913 | 0.055 | 0.235 |
| 2013 | 5136.348 | 1.110 | 1205726.963 | 0.046 | 0.214 |
| 2014 | 5862.202 | 1.266 | 1177918.301 | 0.034 | 0.185 |
| 2015 | 5136.051 | 1.110 | 1174885.547 | 0.045 | 0.211 |
| 2016 | 5352.054 | 1.156 | 1170521.737 | 0.041 | 0.202 |
| 2017 | 4226.525 | 0.913 | 1180886.859 | 0.066 | 0.257 |
| 2018 | 5346.671 | 1.155 | 1035892.222 | 0.036 | 0.190 |
| 2019 | 5317.395 | 1.149 | 1114827.923 | 0.039 | 0.199 |
| 2020 | 4709.373 | 1.017 | 1057118.666 | 0.048 | 0.218 |

Table A.11. Steps to estimate biomass-based abundance indices with standard errors for 1995/96-2020/21 in WAG. GMScaled B_index and B_Index SE were used as CPUE index and its standard error.

| Year | B_Index | GMScaled B_Index | Var(B_index) | Var(B_Index)/(B_Index) ${ }^{2}$ | B_Index SE |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1995 | 9496.922 | 1.346 | 269653.315 | 0.003 | 0.055 |
| 1996 | 7378.444 | 1.046 | 282106.082 | 0.005 | 0.072 |
| 1997 | 7352.241 | 1.042 | 275944.700 | 0.005 | 0.071 |
| 1998 | 7319.828 | 1.038 | 294756.661 | 0.006 | 0.074 |
| 1999 | 6621.496 | 0.939 | 287827.355 | 0.007 | 0.081 |
| 2000 | 5953.261 | 0.844 | 291509.455 | 0.008 | 0.091 |
| 2001 | 5027.783 | 0.713 | 294052.109 | 0.012 | 0.108 |
| 2002 | 6343.625 | 0.899 | 296725.139 | 0.007 | 0.086 |
| 2003 | 8014.390 | 1.136 | 308791.673 | 0.005 | 0.069 |
| 2004 | 8024.629 | 1.138 | 301091.007 | 0.005 | 0.068 |
| 2005 | 14217.841 | 1.109 | 261342.165 | 0.001 | 0.036 |
| 2006 | 13432.365 | 1.048 | 285073.428 | 0.002 | 0.040 |
| 2007 | 12755.536 | 0.995 | 296740.441 | 0.002 | 0.043 |
| 2008 | 15418.982 | 1.203 | 300907.002 | 0.001 | 0.036 |
| 2009 | 20944.373 | 1.634 | 429892.086 | 0.001 | 0.031 |
| 2010 | 14234.971 | 1.110 | 512600.828 | 0.003 | 0.050 |
| 2011 | 15460.033 | 1.206 | 648156.872 | 0.003 | 0.052 |
| 2012 | 14155.450 | 1.104 | 257682.434 | 0.001 | 0.036 |
| 2013 | 10298.831 | 0.803 | 264483.182 | 0.002 | 0.050 |
| 2014 | 10571.862 | 0.825 | 261535.472 | 0.002 | 0.048 |
| 2015 | 8527.265 | 0.665 | 285388.593 | 0.004 | 0.063 |


| 2016 | 10099.723 | 0.788 | 270564.352 | 0.003 | 0.052 |
| ---: | ---: | ---: | ---: | :--- | :--- |
| 2017 | 13139.793 | 1.025 | 260852.220 | 0.002 | 0.039 |
| 2018 | 15980.111 | 1.246 | 278353.271 | 0.001 | 0.033 |
| 2019 | 11131.903 | 0.868 | 291060.530 | 0.002 | 0.048 |
| 2020 | 9929.760 | 0.774 | 293822.916 | 0.003 | 0.055 |

Figures A. 12 and A. 13 compare the non-interaction and interaction effects GLM derived CPUE indices for EAG and WAG, respectively. The estimated indices by the two effects are similar but the confidence intervals for interaction effects are wider than that for main effects.


Figure A.12. Main effects (black) vs. interaction effects (red) CPUE indices during pre- (left panel) and post (right panel)-rationalization periods for EAG. The confidence intervals are $+/-2$ SE.


Figure A.13. Main effects (black) vs. interaction effects (red) CPUE indices during pre- (left panel) and post (right panel)-rationalization periods for WAG. The confidence intervals are $+/-2$ SE.

## c. Non-interaction GLM model without one vessel's data

As per industry request, we calculated non-interaction CPUE indices for pre- and postrationalization periods after removing a vessel's data (confidential vessel identity). This analysis was done for WAG.

The final main effect models for WAG were:
Model 21.1a:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Captain $+\mathrm{ns}($ Soak, 5$)+$ Gear + Area + Month $+\mathrm{ns}($ Depth, 5$)+$
Vessel
AIC=188,469
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + ns $($ Soak, 5$)+$ Area
for the 1995/96-2004/05 period $\left[\theta=0.97, \mathrm{R}^{2}=0.1478\right.$, $\left.\mathrm{AIC}=17,432\right]$
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Gear + Captain + Month + ns $($ Depth, 6$)+\mathrm{ns}($ Soak, 2$)$
AIC=120,331
Final selection by stepCPUE:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Year }+ \text { Captain }+ \text { Gear }+ \text { ns }(\text { Soak, } 2) \tag{A.21}
\end{equation*}
$$

for the 2005/06-2020/21 period $\left[\theta=1.13, R^{2}=0.0538\right.$, AIC $=117,597$, Soak forced in $]$.
Tables A. 12 and A. 13 list the summary statistics of the main effects GLM fits to 1995/96-2004/05 and 2005/06-2020/21 data series, respectively:

Table A.12. Summary statistics of the main effects GLM fit to WAG 1995/96-2004/05 data.

|  | Std. |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
|  | Estimate | Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| (Intercept) | -1.9459 | 1.1432 | -1.7021 | 0.088821 |
| Year1996 | 2.1546 | 1.1490 | 1.8753 | 0.060839 |
| Year1997 | 2.7023 | 1.1472 | 2.3555 | 0.018551 |
| Year1998 | 2.5409 | 1.1549 | 2.2001 | 0.027869 |
| Year1999 | 2.0416 | 1.1481 | 1.7782 | 0.075455 |
| Year2000 | 2.7003 | 1.1475 | 2.3532 | 0.018670 |
| Year2001 | 2.8701 | 1.1505 | 2.4946 | 0.012655 |
| Year2002 | 2.6392 | 1.1497 | 2.2956 | 0.021756 |
| Year2003 | 2.7290 | 1.1518 | 2.3694 | 0.017870 |
| Year2004 | 2.7558 | 1.1489 | 2.3987 | 0.016507 |
| ns(SoakDays, df $=5) 1$ | 0.5816 | 0.1036 | 5.6145 | 0.000000 |
| ns(SoakDays, df $=5) 2$ | 0.2478 | 0.1218 | 2.0342 | 0.042006 |
| ns(SoakDays, $\mathrm{df}=5) 3$ | 1.1716 | 0.1436 | 8.1595 | 0.000000 |
| ns(SoakDays, df $=5) 4$ | 0.7868 | 0.1469 | 5.3567 | 0.000000 |
| ns(SoakDays, $\mathrm{df}=5) 5$ | 0.6941 | 0.1427 | 4.8636 | 0.000001 |
| Block8 | 0.2165 | 0.0962 | 2.2510 | 0.024448 |
| Block9 | 0.4628 | 0.0820 | 5.6456 | 0.000000 |
| Block10 | 0.9291 | 0.0843 | 11.0261 | 0.000000 |

Table A.13. Summary statistics of the main effects GLM fit to WAG 2005/06-2020/21 data.

|  | Std. |  |  |  |
| :--- | ---: | :---: | ---: | ---: |
|  | Estimate | Error | t value | $\operatorname{Pr}( \rangle\|t\|)$ |
| (Intercept) | 2.2212 | 0.1770 | 12.5495 | 0.000000 |
| Year2006 | -0.0569 | 0.0470 | -1.2110 | 0.225915 |
| Year2007 | -0.1741 | 0.0523 | -3.3287 | 0.000875 |
| Year2008 | -0.0388 | 0.0432 | -0.8986 | 0.368856 |
| Year2009 | 0.0107 | 0.0435 | 0.2460 | 0.805714 |
| Year2010 | -0.1360 | 0.0451 | -3.0174 | 0.002554 |
| Year2011 | -0.0778 | 0.0468 | -1.6632 | 0.096297 |
| Year2012 | -0.0472 | 0.0458 | -1.0311 | 0.302492 |
| Year2013 | -0.3871 | 0.0474 | -8.1736 | 0.000000 |
| Year2014 | -0.4423 | 0.0515 | -8.5808 | 0.000000 |
| Year2015 | -0.4587 | 0.0504 | -9.1008 | 0.000000 |


| Year2016 | -0.3312 | 0.0515 | -6.4311 | 0.000000 |
| :--- | ---: | ---: | ---: | ---: |
| Year2017 | -0.1799 | 0.0557 | -3.2313 | 0.001235 |
| Year2018 | 0.0389 | 0.0569 | 0.6830 | 0.494603 |
| Year2019 | -0.1950 | 0.0519 | -3.7556 | 0.000174 |
| Year2020 | -0.3312 | 0.0510 | -6.4902 | 0.000000 |
| Captain104 | 0.4572 | 0.1783 | 2.5649 | 0.010330 |
| Captain131 | 0.4789 | 0.1795 | 2.6678 | 0.007643 |
| Captain133 | 0.6408 | 0.1813 | 3.5342 | 0.000410 |
| Captain145 | 0.6813 | 0.1761 | 3.8696 | 0.000109 |
| Captain166 | 0.5354 | 0.1770 | 3.0245 | 0.002495 |
| Captain215 | 0.4562 | 0.1961 | 2.3267 | 0.019994 |
| Captain257 | 0.4944 | 0.1828 | 2.7044 | 0.006851 |
| Captain336 | 0.4551 | 0.1855 | 2.4541 | 0.014134 |
| Captain384 | 0.4465 | 0.1861 | 2.3996 | 0.016423 |
| Captain403 | 0.5965 | 0.1793 | 3.3262 | 0.000882 |
| Captain404 | 0.5224 | 0.1837 | 2.8437 | 0.004466 |
| Captain405 | 0.5968 | 0.1806 | 3.3044 | 0.000954 |
| Gear6 | 0.2772 | 0.0260 | 10.6599 | 0.000000 |
| Gear7 | 0.3227 | 0.0289 | 11.1472 | 0.000000 |
| Gear8 | 0.5816 | 0.0355 | 16.3646 | 0.000000 |
| Gear13 | 0.8796 | 0.1842 | 4.7758 | 0.000002 |
| Gear25 | 1.3950 | 0.3369 | 4.1409 | 0.000035 |
| ns(SoakDays, df =2)1 | 0.3139 | 0.0604 | 5.1988 | 0.000000 |
| ns(SoakDays, df =2)2 | 0.0603 | 0.0376 | 1.6035 | 0.108845 |

Figure 14 shows the comparison of CPUE indices between the full and reduced data sets for WAG. Removal of one vessel's data has significantly affected the CPUE indices during the prerationalization period but not the post-rationalization period.


Figure A.14. Comparison of the trends in standardized CPUE indices for full data (black) and reduced data (red) for WAG. The confidence intervals are $+/-2$ SE.

## Commercial fishery CPUE index by non-interaction model

We fitted the negative binomial GLM model for fish ticket retained CPUE time series 1985/86 1998/99 offering Year, Month, Vessel, Captain, and Area as explanatory variables and applying the hybrid selection method. Reduced area resolution (grouped ADF\&G codes to AreaGP) was used for model fitting.

The final model for EAG was:
Initial selection by stepAIC:
$\ln$ (CPUE) $=$ Year + Vessel + Month
AIC=16,996
Final selection by stepCPUE:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Year }+ \text { Vessel }+ \text { Month } \tag{A.22}
\end{equation*}
$$

for the $1985 / 86-1998 / 99$ period $\left[\theta=10.40, \mathrm{R}^{2}=0.3327, \mathrm{AIC}=16,535\right]$
and that for WAG was:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Vessel + Area
AIC=31,701

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Vessel + Area
for the 1985/86-1998/99 period $\left[\theta=6.67, \mathrm{R}^{2}=0.3569\right.$, AIC $\left.=31,215\right]$
We did not fine tune the fishery CPUE fits for nonsignificant parameter estimates because this drastically reduced the number of data points for the fit, especially EAG data.

Figures A. 15 and A. 16 compare standardized and nominal CPUE indices for 1985/86-1998/99 fishery data for EAG and WAG, respectively.


Figure A.15. Trends in non-standardized (red line) vs. standardized CPUE (black line with confidence intervals) indices during 1985/86-1998/99 period for EAG. The confidence intervals are +/- 2 SE.


Figure A.16. Trends in non-standardized (red line) vs. standardized CPUE (black line with confidence intervals) indices during 1985/86-1998/99 period for WAG. The confidence intervals are +/- 2 SE.

## Appendix B: Male Maturity

## Introduction

Sexual maturity is associated with alterations in both external morphology, internal physiology, and incidence of copulation on which bases different types of maturity can be defined:
physiological, morphometric, and functional maturity. Although functional maturity is the true way of determining maturity, it requires elaborate lab or field experiments. Hence, crab researchers often adapt an indirect detection technique via morphometric measurement for maturity determination. Chelae allometry has been used to determine morphometric male size-atmaturity among several king crab (Lithodidae) stocks. Male golden king crab provide a better discrimination of chelae height against size at onset of maturity than other king crab stocks (Somerton and Otto 1986). Table B. 1 lists the literature reported estimates of size-at-maturity of male golden king crab (Lithodes aequispins) stocks in Alaska. Breakpoint analysis has been used to estimate maturity in majority of cases.

Table B.1. Review of estimates of male size-at-maturity of golden (Lithodes aequispins) king crab by regions in Alaska. Numbers in parentheses are standard errors (SE).

| Species | Sex | Size-at- <br> Maturity <br> $(\mathbf{m m ~ C L})$ | Method | Area | Sources |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Lithodes <br> aequispins | Male | $114(11.4)$ | Breakpoint analysis on <br> ln chela height vs. In <br> carapace length <br> Breakpoint analysis on <br> ln chela height vs. ln <br> carapace length | British Columbia, <br> Canada | St. Matthew Is. District <br> Pribilof Is. District <br> Eastern Aleutian Is. | | Somerton and |
| :--- |
|  |

[^0]Table B.2. Golden king crab male carapace length and chela height data collected during 2018/19 - 2020/21 fishing seasons in the Aleutian Islands.

| Measurement type | Source and season of <br> data collection | Aleutian Islands <br> $(\mathrm{AI}) \quad 2018 / 19-$ <br> $2020 / 21$ | EAG <br> $2018 / 19-$ <br> $2019 / 20$ | WAG <br> 2018/19- <br> $2020 / 21$ |
| :--- | :--- | :--- | :--- | :--- |
|  | Co-operative survey <br> $(2018 / 19,2019 / 20)$ |  |  |  |
|  | Observer sampling <br> $(2018 / 19,2019 / 20)$ |  |  |  |
|  | Retained catch <br> sampling (2018/19, <br> 2019/20, 2020/21) |  |  |  |
|  | Special sampling |  | 5351 |  |
| Carapace length <br> and chela height <br> records (all sizes) | WAG (2020/21) |  |  |  |

## Method

The male size-at-maturity is determined as the breakpoint in the following model:
$C H=\beta_{0}+\beta_{1} C L+\beta_{2}[C L-c]^{+}+\varepsilon$
where $\beta_{0}$ is the intercept, $\beta_{1}$ is the left slope, $\beta_{2}$ is the difference in slopes when $C L \geq c$, and $c$ is the breakpoint and $\varepsilon$ is the random error.

The term $[C L-c]^{+}$reduces to zero if $C L<c$, otherwise takes the value of the argument in the following form of the model:
$C H=\beta_{0}+\beta_{1} C L+\beta_{2}[C L-c]$

The "segmented regression" package by Muggeo (2003, 2008), available in R (ver 4.1, R Core Team 2021), is used to determine breakpoints and corresponding two segmented lines for different groups of data outside the assessment model. Muggeo's method first fits a single line to CH vs CL data and then proceeds to estimate an optimum break point iteratively from an initial guess value over the CL range. In the process, it estimates the parameters of equation B. 2 including the breakpoint. Olson et al. (2018) followed a similar approach to analyze CH vs CL data in the southeast Alaska but used a different R package to that of Muggeo.

The estimates are further refined by bootstrapping each data set (CH, CL pairs) 1000 times and applying 'segmented regression' to each boot strapped sample. The bootstrap median breakpoint (i.e., size-at-maturity), standard error, and confidence intervals are also estimated.

## Results

The original sample data produced the breakpoints (size-at-maturity) of 117.865 mm CL, 104.295 mm CL, and 120.199 mm CL for AI, EAG, and WAG, respectively. The EAG data produced the lowest estimate but unreliable. This is likely due to existence of outliers in the EAG data. Hence, the EAG data are restricted to a plausible size range $85-142 \mathrm{~mm}$ CL within which the breakpoint is likely to fall, and re-estimated the breakpoint to be 128.72 mm CL. On the other hand, breakpoint estimates from whole data of WAG and AI are reliable (see Figures B.1, B.2, and B.3).

Because of uncertainty in EAG breakpoint estimate, the AI estimate is considered as reliable for applying to both EAG and WAG regions. The bootstrap analysis is also done only on AI data. The bootstrap statistics are listed in Table B.3:

Table B.3. Bootstrap estimate of breakpoint with standard error and confidence bounds for AI 2018/19-2020/21 data.

| Parameter | Mean | Median | SE | Upper <br> Bound | Lower <br> Bound |
| :---: | :---: | :---: | :---: | :---: | :---: |
| breakpoint | 116.575 | 117.996 | 0.159 | 122.562 | 105.212 |

The breakpoint (mean/median) values are approximately one 5 mm CL bin higher than the currently used 111 mm CL. Two options for MMB estimation are suggested: $\geq \mathbf{1 1 1} \mathbf{~ m m} \mathbf{C L}$ (lower limit of the 111-115 mm CL bin, status quo knife edge maturity) and $\geq 116 \mathbf{~ m m ~ C L}$ (lower limit of the 116-120 mm CL bin, based on present analysis).

Note that the mean and median estimates are 117 and 118 mm CL, respectively, falling within the 116-120 mm CL bin.


Figure B.1. Segmented linear regression fit to CH vs. CL data (restricted to $85-142 \mathrm{~mm}$ CL size range) of male golden king crab for 2018/19-2020/21 in EAG.


Figure B.2. Segmented linear regression fit to CH vs. CL data of male golden king crab for 2018/19-2020/21 in WAG.


Figure B.3. Segmented linear regression fit to CH vs. CL data of male golden king crab for 2018-2020 in AI.

## Appendix C: Jittering

## Jittering of model 21.1 a parameter estimates:

We followed the Stock Synthesis approach to do 100 jitter runs of model 21.1a parameter estimates to use as initial parameter values (as .PIN file in ADMB) to assess model stability and to determine whether a global as opposed to local minima has been reached by the search algorithm:

The Jitter factor of 0.3 was multiplied by a random normal deviation $r \operatorname{dev}=N(0,1)$, to a transformed parameter value based upon the predefined parameter:

$$
\begin{equation*}
\text { temp }=0.5 * r d e v^{*} \text { Jitterfactor } * \ln \left(\frac{P_{\max }-P_{\min }+0.0000002}{P_{v a l}-P_{\min }+0.0000001}-1\right), \tag{C.1}
\end{equation*}
$$

with the final jittered initial parameter value back transformed as:

$$
\begin{equation*}
P_{\text {new }}=P_{\min }+\frac{P_{\max }-P_{\min }}{1.0+\exp (-2.0 \text { temp })} \tag{C.2}
\end{equation*}
$$

where $P_{\max }$ and $P_{\min }$ are upper and lower bounds of parameter search space and $P_{\text {val }}$ is the estimated parameter value before the jittering.

The jitter results are summarized for model 21.1a in Table C. 1 for EAG and in Table C. 2 for WAG. Original fits produced the highest log likelihood values (global minimum) for EAG and WAG.

Table C.1. Results from 100 jitter runs for model 21.1a for EAG. Jitter run 0 corresponds to the original optimized estimates. NA: not converged.

$\left.\begin{array}{ccccccrr} & 13 & & -890.8549 & & 0.00012718 & 9,298 & 3,795 \\ \\ & 14 & & -890.8549 & & 0.00004932 & 9,298 & 3,795\end{array}\right)$


Table C.2. Results from 100 jitter runs for model 21.1a for WAG. Jitter run 0 corresponds to the original optimized estimates. NA: not converged.

| Jitter <br> Run | Negative Log Likelihood | Maximum Gradient | $\mathrm{B}_{35 \%}(\mathrm{t})$ | OFL (t) | Current MMB (t) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| NA | -940.3565 | 0.00012460 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00018420 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00009135 | 7,370 | 1,669 | 6,702 |
|  | NA | NA | NA | NA | NA |
|  | -940.3565 | 0.00013242 | 7,370 | 1,669 | 6,702 |
| 5 | -940.3565 | 0.00024676 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00000718 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00031064 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00004772 | 7,370 | 1,669 | 6,702 |
|  | 1016.2650 | 12490.57000000 | 0 | 4 | 24,262 |
| 10 | -940.3565 | 0.00008327 | 7,370 | 1,669 | 6,702 |
| 11 | -940.3565 | 0.00005325 | 7,370 | 1,669 | 6,702 |
| 12 | -940.3565 | 0.00016437 | 7,370 | 1,669 | 6,702 |
| 13 | 6409.1580 | 3123883.00000000 | 82,616 | 2,491 | 33,440 |
| 14 | -940.3565 | 0.00005322 | 7,370 | 1,669 | 6,702 |
| 15 | -940.3565 | 0.00001754 | 7,370 | 1,669 | 6,702 |
| 16 | -940.3565 | 0.00005742 | 7,370 | 1,669 | 6,702 |
| 17 | -531.8893 | 1452.33500000 | 9,208 | 2,710 | 8,619 |
| 18 | -940.3565 | 0.00015244 | 7,370 | 1,669 | 6,702 |
| NA | NA | NA | NA | NA | NA |
|  | -940.3565 | 0.00006466 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00000543 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00033342 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00003234 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00031754 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00000897 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00007067 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00006223 | 7,370 | 1,669 | 6,702 |
|  | 16324.7100 | 738.99210000 | 1,974,190 | 1,145,060 | 2,986,830 |
|  | -940.3565 | 0.00218114 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00016393 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00004037 | 7,370 | 1,669 | 6,702 |
|  | -494.3506 | 1517.71400000 | 8,991 | 2,597 | 8,674 |
|  | -940.3565 | 0.00008052 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00020356 | 7,370 | 1,669 | 6,702 |
|  | -940.3565 | 0.00008063 | 7,370 | 1,669 | 6,702 |
|  | 1014.9310 | 1597.97000000 | 195,483 | 165,717 | 442,185 |


| 37 | -940.3565 | 0.00012173 | 7,370 | 1,669 | 6,702 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| 38 | -940.3565 | 0.00007363 | 7,370 | 1,669 | 6,702 |
| 39 | -940.3565 | 0.00013596 | 7,370 | 1,669 | 6,702 |
| 40 | -940.3565 | 0.00011785 | 7,370 | 1,669 | 6,702 |
| 41 | 1955.0850 | 2440.73800000 | 11,428 | 2,785 | 976 |
| 42 | -940.3565 | 0.00001743 | 7,370 | 1,669 | 6,702 |
| 43 | -940.3565 | 0.00014257 | 7,370 | 1,669 | 6,702 |
| 44 | -940.3565 | 0.00006468 | 7,370 | 1,669 | 6,702 |
| 45 | -940.3565 | 0.00027955 | 7,370 | 1,669 | 6,702 |
| 46 | -940.3565 | 0.00004034 | 7,370 | 1,669 | 6,702 |
| 47 | -940.3565 | 0.00007934 | 7,370 | 1,669 | 6,702 |
| 48 | 16057.2500 | 759.14390000 | 53,287 | 45,551 | 118,591 |
| 49 | -940.3565 | 0.00008810 | 7,370 | 1,669 | 6,702 |
| 50 | -940.3565 | 0.00002177 | 7,370 | 1,669 | 6,702 |
| 51 | -940.3565 | 0.00005384 | 7,370 | 1,669 | 6,702 |
| 52 | -940.3565 | 0.00009783 | 7,370 | 1,669 | 6,702 |
| 53 | -940.3565 | 0.00032531 | 7,370 | 1,669 | 6,702 |
| 54 | -940.3565 | 0.00005679 | 7,370 | 1,669 | 6,702 |
| 55 | -940.3565 | 0.00004213 | 7,370 | 1,669 | 6,702 |
| 56 | -940.3565 | 0.00010423 | 7,370 | 1,669 | 6,702 |
| 57 | 922.1206 | 3510.20200000 | 6,339 | 1,777 | 5,545 |
| 58 | -940.3565 | 0.00003724 | 7,370 | 1,669 | 6,702 |
| 59 | -940.3565 | 0.00002954 | 7,370 | 1,669 | 6,702 |
| 60 | -940.3565 | 0.00006124 | 7,370 | 1,669 | 6,702 |
| 61 | -940.3565 | 0.00004589 | 7,370 | 1,669 | 6,702 |
| 62 | -940.3565 | 0.00001069 | 7,370 | 1,669 | 6,702 |
| 63 | -940.3565 | 0.00000625 | 7,370 | 1,669 | 6,702 |
| 64 | -940.3565 | 0.00039464 | 7,370 | 1,669 | 6,702 |
| 65 | -940.3565 | 0.00005705 | 7,370 | 1,669 | 6,702 |
| 66 | 1396.0840 | 142015.20000000 | 374,199 | 261,351 | 972,679 |
| 67 | -940.3565 | 0.00001739 | 7,370 | 1,669 | 6,702 |
| 68 | 3966.7420 | 40652.93000000 | 0 | 3 | 27,940 |
| 69 | -940.3565 | 0.00015450 | 7,370 | 1,669 | 6,702 |
| 70 | -940.3565 | 0.00006020 | 7,370 | 1,669 | 6,702 |
| 71 | -940.3565 | 0.00012604 | 7,370 | 1,669 | 6,702 |
| 72 | -940.3565 | 0.00004411 | 7,370 | 1,669 | 6,702 |
| 73 | -940.3565 | 0.00012948 | 7,370 | 1,669 | 6,702 |
| 74 | 723.5230 | 32543.63000000 | 43,405 | 39,397 | 72,414 |
| 75 | -589.7174 | 0.00080621 | 7,780 | 2,211 | 7,481 |
| 76 | -940.3565 | 0.00005599 | 7,370 | 1,669 | 6,702 |
| 77 | -940.3565 | 0.00033817 | 7,370 | 1,669 | 6,702 |
| 78 | 703.5092 | 4897.44500000 | 242,756 | 210,920 | 1,669 |
| 79 | -940.3565 | 0.00025400 | 7,370 | 1,669 |  |
| 80 | 1211.0410 | 79437.21000000 | 213,118 | 184,110 | 415,783 |
|  |  |  |  |  | 7 |


| 81 | -842.6195 | 0.00002915 | 7,088 | 1,468 | 6,272 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 82 | -940.3565 | 0.00004797 | 7,370 | 1,669 | 6,702 |
| 83 | -940.3565 | 0.00002080 | 7,370 | 1,669 | 6,702 |
| 84 | -940.3565 | 0.00004050 | 7,370 | 1,669 | 6,702 |
| 85 | -940.3565 | 0.00009084 | 7,370 | 1,669 | 6,702 |
| 86 | 819.5523 | 8582.10500000 | 89,624 | 84,054 | 186,557 |
| 87 | -940.3565 | 0.00010075 | 7,370 | 1,669 | 6,702 |
| 88 | -940.3565 | 0.00005397 | 7,370 | 1,669 | 6,702 |
| 89 | 408.8454 | 100758.20000000 | 24,013 | 19,444 | 41,721 |
| 90 | -940.3565 | 0.00000698 | 7,370 | 1,669 | 6,702 |
| 91 | -940.3565 | 0.00013725 | 7,370 | 1,669 | 6,702 |
| 92 | 1170.6040 | 257733.60000000 | 28,695 | 4,546 | 15,019 |
| 93 | -489.5243 | 0.00025367 | 8,991 | 2,597 | 8,674 |
| 94 | -940.3565 | 0.00031713 | 7,370 | 1,669 | 6,702 |
| 95 | -940.3565 | 0.00005909 | 7,370 | 1,669 | 6,702 |
| 96 | 382.8459 | 0.00316184 | 5,071 | 1,418 | 4,999 |
| 97 | -940.3565 | 0.00008396 | 7,370 | 1,669 | 6,702 |
| 98 | -940.3565 | 0.00004028 | 7,370 | 1,669 | 6,702 |
| 99 | -842.6195 | 0.00040310 | 7,088 | 1,468 | 6,272 |
| 100 | -940.3565 | 0.00006259 | 7,370 | 1,669 | 6,702 |

# Appendix D: RACE AIGKC Slope Survey 

L. Lee, M.S.M. Siddeek, and C. Chris

Alaska Department of Fish and Game, Division of Commercial Fisheries, P.O. Box 115526, Juneau, Alaska 99811

## Introduction

The SSC in their June 2021 meeting requested Aleutian Islands golden king crab (AIGKC) assessment authors to consider including Resource Assessment and Conservation Engineering (RACE) Division biennial slope survey data in the assessment. This appendix provides some introductory tables and figures on AIGKC slope (trawl) surveys' data with the anticipation of some guidance from CPT/SSC on appropriate ways to incorporate these data into AIGKC assessment model.

## Method

Comparison of RACE AIGKC slope survey data with observer sample data
The RACE conducted biennial (trawl) slope surveys in the Aleutian Islands starting in 1980. The Poly'Noreatern (PNE) net was used in the trawl since 1991. Due to logistic problems, there were some gaps in survey periodicity during the 1980-2018 period (Table D.1). For data exploration and comparison with observer pot sample CPUE, the 1997-2018 slope survey CPUE (standardized for 15 minutes tow for the trawl configuration in 1991) were used. The trawl survey data comprised of all sizes and sexes of golden king crab. Unfortunately, no size measurements were recorded in trawl samples. To obtain a comparable observer CPUE to that of trawl survey, a new "Total Crab" column was created in the observer database with females, sublegal, legal retained, and legal non retained crab numbers pooled. Observer CPUE data by lat//long. locations were summed up within a given $1 \times 1 \mathrm{nmi}$ cell. Since trawl CPUE data were reported by trawl start locations, they were not summed up within the $1 \times 1 \mathrm{nmi}$ cells but used as they were.

Furthermore, because the magnitude of survey CPUE and observer CPUE were different, they were scaled by the respective maxima for the Fisheries Management (FMP) seasons. The scaled (i.e., proportion) CPUE values for 1997, 2000, 2002, 2004, 2006, 2008, 2010, 2012, 2014, 2016, and 2018 of observer and trawl survey samples were plotted in several figures (Figures D. 1 to D.7). The total crab counts for observer samples summed up within the $1 \times 1 \mathrm{nmi}$ cells ranged from 467 to 1433 whereas the crab counts per trawl tow ranged from 10 to 289 .

## Results

Table D. 1 lists the trawl survey nonzero CPUE, and abundance estimates made for EAG, WAG, and the other nonspecific Aleutian Islands areas by RACE for 1980-2018. Unusually high catches in certain year's trawl tow (for example, 2018) yielded high abundance variance. A
detailed trawl survey database with zero golden king crab catch tows is available but not summarized in this note.

Figure D. 1 shows the Observer-Trawl CPUE proportion summary comparison bars in 10 blocks, which covers the entire Aleutian Islands and for the selected 1997 to 2018 range. These blocks are currently used in AIGKC CPUE standardization with 1 to 4 numbered blocks for EAG and 5 to 10 blocks for WAG. Figures D. 2 to D. 7 show the detail observer CPUE proportion bars by 1x1 nmi cell and trawl CPUE proportion bars by location for the entire Aleutian Islands, EAG, and WAG, respectively for arbitrarily selected years, 1997 and 2018. These figures indicate that trawl survey encountered golden king crab density outside the observer sampled areas. Whether any fishing vessel frequented these locations or not during the fishing season was not investigated.

## Acknowledgement

Wayne Palsson from the RACE division of AFSC provided the AIGKC slope survey data and it is greatly appreciated.

Table D.1. Summary CPUE and biomass estimates of Aleutian Islands golden king crab from RACE slope survey during 1980-2018.

| $\begin{aligned} & \text { FMP } \\ & \text { Area } \end{aligned}$ | Year | Area Code | Haul Count | Catch Count | $\begin{aligned} & \text { Mean CPUE } \\ & (\mathrm{Kg}) \end{aligned}$ | Var CPUE | $\begin{aligned} & \text { Mean CPUE } \\ & \text { (no.) } \end{aligned}$ | Var CPUE | Area Biomass(t) | Var <br> Biomass | Area <br> Abundance (no.) | Var <br> Abundance |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WAG | 1980 | 543 | 11 | 4 | 19.265 | 167.693 | 12.350 | 25.059 | 236.700 | 25353.636 | 152394 | 3788715876 |
| AI | 1980 | 518+519 | 18 | 1 | 11.998 | 143.932 | 10.290 | 105.802 | 77.200 | 5966.406 | 66225 | 4385809142 |
| WAG | 1980 | 542 | 52 | 17 | 85.275 | 734.599 | 89.150 | 587.655 | 1272.500 | 163634.690 | 1330855 | 130902000000 |
| EAG | 1980 | 541 | 46 | 9 | 48.845 | 99.868 | 49.754 | 131.784 | 1053.500 | 46391.993 | 1072126 | 61217973504 |
| WAG | 1983 | 543 | 63 | 40 | 45.151 | 112.667 | 32.593 | 60.710 | 685.900 | 25997.674 | 495178 | 14008607911 |
| AI | 1983 | $518+519$ | 34 | 7 | 3.821 | 4.803 | 3.126 | 2.029 | 28.700 | 268.851 | 23398 | 113598097.1 |
| WAG | 1983 | 542 | 96 | 47 | 103.988 | 2685.281 | 64.591 | 600.455 | 1430.900 | 508491.312 | 888990 | 113704000000 |
| EAG | 1983 | 541 | 97 | 15 | 10.110 | 17.876 | 8.542 | 14.987 | 222.900 | 8672.056 | 188387 | 7270693560 |
| WAG | 1986 | 543 | 81 | 41 | 58.454 | 405.428 | 51.184 | 198.123 | 818.700 | 79544.021 | 716840 | 38871177904 |
| AI | 1986 | $518+519$ | 63 | 9 | 3.996 | 2.121 | 3.932 | 1.693 | 29.900 | 118.743 | 29451 | 94786651.87 |
| WAG | 1986 | 542 | 121 | 47 | 19.337 | 14.561 | 18.945 | 15.946 | 320.100 | 3984.618 | 313382 | 4363767214 |
| EAG | 1986 | 541 | 118 | 27 | 12.740 | 9.250 | 13.918 | 16.443 | 321.000 | 5874.043 | 350737 | 10442574472 |
| WAG | 1991 | 543 | 56 | 7 | 8.543 | 11.016 | 10.811 | 23.126 | 129.800 | 2542.017 | 164278 | 5336232548 |
| AI | 1991 | 518+519 | 55 | 3 | 2.651 | 2.858 | 2.615 | 3.011 | 19.800 | 159.989 | 19561 | 168551567.2 |
| WAG | 1991 | 542 | 91 | 19 | 15.797 | 9.875 | 17.159 | 11.551 | 261.000 | 2702.328 | 284065 | 3160909866 |
| EAG | 1991 | 541 | 129 | 17 | 5.367 | 3.753 | 4.452 | 2.012 | 135.300 | 2383.157 | 111957 | 1278027114 |
| WAG | 1994 | 543 | 69 | 16 | 12.234 | 8.721 | 17.771 | 14.258 | 185.800 | 2012.387 | 270092 | 3290012068 |
| AI | 1994 | 518+519 | 64 | 3 | 4.198 | 3.640 | 8.625 | 19.559 | 31.400 | 203.725 | 64509 | 1094811866 |
| WAG | 1994 | 542 | 114 | 30 | 31.464 | 76.864 | 27.701 | 54.770 | 520.700 | 21034.499 | 458009 | 14988180030 |
| EAG | 1994 | 541 | 133 | 37 | 15.318 | 15.525 | 19.341 | 16.161 | 386.000 | 9859.527 | 487340 | 10262963120 |
| WAG | 1997 | 543 | 92 | 16 | 15.572 | 19.626 | 17.675 | 23.512 | 236.400 | 4528.685 | 268548 | 5425278407 |
| AI | 1997 | $518+519$ | 52 | 3 | 1.617 | 0.856 | 2.620 | 2.814 | 12.100 | 47.930 | 19638 | 157529796.8 |
| WAG | 1997 | 542 | 116 | 19 | 9.048 | 5.516 | 12.285 | 7.802 | 149.600 | 1509.376 | 203159 | 2135098797 |
| EAG | 1997 | 541 | 136 | 38 | 34.274 | 25.520 | 30.520 | 17.876 | 863.700 | 16207.115 | 769185 | 11352306878 |
| WAG | 2000 | 543 | 113 | 26 | 16.454 | 30.248 | 17.236 | 32.760 | 249.800 | 6979.702 | 261830 | 7559351301 |
| AI | 2000 | 518+519 | 58 | 5 | 13.124 | 57.032 | 32.078 | 378.820 | 98.200 | 3192.318 | 239992 | 21203988926 |
| WAG | 2000 | 542 | 110 | 27 | 28.265 | 66.037 | 28.046 | 83.672 | 467.600 | 18071.623 | 463851 | 22897430722 |
| EAG | 2000 | 541 | 138 | 53 | 33.856 | 56.748 | 42.181 | 48.507 | 853.300 | 36038.604 | 1063064 | 30804930379 |
| WAG | 2002 | 543 | 107 | 39 | 33.398 | 39.411 | 33.203 | 17.977 | 507.300 | 9093.912 | 504335 | 4148202305 |


| AI | 2002 | $518+519$ | 61 | 6 | 8.439 | 23.213 | 12.957 | 42.515 | 63.200 | 1299.294 | 96917 | 2379716807 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WAG | 2002 | 542 | 114 | 25 | 34.388 | 116.472 | 43.792 | 175.112 | 568.800 | 31873.483 | 724447 | 47920683833 |
| EAG | 2002 | 541 | 132 | 39 | 29.196 | 32.510 | 36.388 | 47.325 | 735.700 | 20645.914 | 917304 | 30054356912 |
| WAG | 2004 | 543 | 124 | 25 | 22.603 | 29.164 | 16.455 | 11.404 | 343.200 | 6729.447 | 249763 | 2631445737 |
| AI | 2004 | $518+519$ | 53 | 6 | 11.464 | 56.042 | 19.011 | 132.096 | 85.800 | 3136.876 | 142259 | 7393916636 |
| WAG | 2004 | 542 | 130 | 29 | 34.812 | 64.564 | 32.933 | 53.801 | 575.900 | 17668.550 | 544857 | 14723044390 |
| EAG | 2004 | 541 | 112 | 35 | 47.289 | 132.887 | 43.661 | 80.601 | 1191.700 | 84391.707 | 1100458 | 51186919912 |
| WAG | 2006 | 543 | 112 | 38 | 29.132 | 29.122 | 31.202 | 59.603 | 442.400 | 6719.867 | 474126 | 13753283045 |
| AI | 2006 | $518+519$ | 44 | 6 | 42.705 | 310.252 | 70.652 | 1271.771 | 319.500 | 17365.938 | 528625 | 71185748167 |
| WAG | 2006 | 542 | 110 | 20 | 36.027 | 418.409 | 33.510 | 436.353 | 596.000 | 114500.967 | 554347 | 119411000000 |
| EAG | 2006 | 541 | 91 | 34 | 47.617 | 110.667 | 43.110 | 67.536 | 1199.800 | 70280.696 | 1086721 | 42889790268 |
| WAG | 2010 | 543 | 118 | 42 | 57.886 | 130.248 | 46.341 | 106.550 | 879.500 | 30054.498 | 704058 | 24586118732 |
| AI | 2010 | $518+519$ | 51 | 3 | 10.806 | 29.846 | 14.152 | 43.404 | 80.800 | 1670.570 | 105884 | 2429491153 |
| WAG | 2010 | 542 | 128 | 21 | 30.437 | 54.021 | 23.633 | 27.481 | 503.400 | 14783.401 | 390787 | 7520442999 |
| EAG | 2010 | 541 | 121 | 34 | 34.371 | 39.101 | 28.945 | 35.782 | 866.500 | 24831.388 | 729297 | 22723745641 |
| WAG | 2012 | 543 | 120 | 43 | 43.611 | 60.776 | 41.818 | 58.628 | 662.400 | 14024.029 | 635292 | 13528187569 |
| AI | 2012 | 518+519 | 55 | 7 | 13.615 | 38.741 | 15.753 | 65.204 | 101.800 | 2168.505 | 117853 | 3649691138 |
| WAG | 2012 | 542 | 113 | 29 | 42.050 | 143.618 | 35.735 | 74.063 | 695.500 | 39302.255 | 591184 | 20267923534 |
| EAG | 2012 | 541 | 132 | 46 | 50.091 | 197.620 | 65.206 | 667.864 | 1262.400 | 125501.093 | 1643226 | 424136000000 |
| WAG | 2014 | 543 | 134 | 48 | 45.987 | 44.653 | 37.885 | 49.998 | 698.700 | 10303.650 | 575577 | 11536988323 |
| AI | 2014 | 518+519 | 44 | 4 | 36.935 | 473.737 | 56.115 | 1358.509 | 276.300 | 26516.828 | 419856 | 76040808114 |
| WAG | 2014 | 542 | 110 | 14 | 20.185 | 44.188 | 53.795 | 1384.010 | 333.900 | 12092.417 | 889822 | 378745000000 |
| EAG | 2014 | 541 | 122 | 40 | 92.271 | 504.653 | 103.911 | 1821.531 | 2325.400 | 320486.541 | 2618618 | 1156790000000 |
| WAG | 2016 | 543 | 135 | 36 | 36.030 | 52.539 | 21.555 | 16.354 | 547.500 | 12123.182 | 327299 | 3773719235 |
| AI | 2016 | $518+519$ | 43 | 5 | 9.520 | 20.923 | 7.931 | 14.731 | 71.300 | 1171.165 | 59278 | 824527966.8 |
| WAG | 2016 | 542 | 114 | 17 | 11.674 | 16.990 | 11.568 | 13.615 | 193.200 | 4649.368 | 191277 | 3725777828 |
| EAG | 2016 | 541 | 127 | 37 | 34.533 | 69.894 | 36.160 | 78.286 | 869.900 | 44387.175 | 911427 | 49716540459 |
| WAG | 2018 | 543 | 129 | 33 | 27.739 | 50.241 | 18.340 | 15.318 | 421.500 | 11592.979 | 278860 | 3534573172 |
| AI | 2018 | 518+519 | 45 | 7 | 16.027 | 13.350 | 17.651 | 19.676 | 120.000 | 747.269 | 132046 | 1101315220 |
| WAG | 2018 | 542 | 120 | 19 | 21.075 | 55.306 | 16.387 | 23.418 | 348.400 | 15134.868 | 271070 | 6408654749 |
| EAG | 2018 | 541 | 126 | 39 | 156.633 | 13816.580 | 119.392 | 3756.474 | 3947.300 | 8774393.941 | 3008811 | 2385600000000 |



Figure D.1. Comparison of Aleutian Islands golden king crab index of abundance in Blocks\#1 to 10 between RACE slope survey (light brown bars) and corresponding years' observer samples (light blue bars).


Figure D.2. Comparison of Aleutian Islands golden king crab index of abundance in Blocks\#1 to 10 between RACE slope survey (light brown bars) and observer samples (light blue bars) during 1997.


Figure D.3. Comparison of EAG golden king crab index of abundance in Blocks\#1 to 4 between RACE slope survey (light brown bars) and observer samples (light blue bars) during 1997.


Figure D.4. Comparison of WAG golden king crab index of abundance in Blocks\#5 to 10 between RACE slope survey (light brown bars) and observer samples (light blue bars) during 1997.


Figure D.5. Comparison of Aleutian Islands golden king crab index of abundance in Blocks\#1 to 10 between RACE slope survey (light brown bars) and observer samples (light blue bars) during 2018.


Figure D.6. Comparison of EAG golden king crab index of abundance in Blocks\#1 to 4 between RACE slope survey (light brown bars) and observer samples (light blue bars) during 2018.


Figure D.7. Comparison of WAG golden king crab index of abundance in Blocks $\# 5$ to 10 between RACE slope survey (light brown bars) and observer samples (light blue bars) during 2018.

## Appendix E: EAG 21.1a model implementation in Gmacs

## Gmacs working session:

A working session on AIGKC model implementation in Gmacs was conducted by Andre Punt during 1-3 December 2021 in Juneau. Andre Punt, Shareef Siddeek, Katie Palof, and Cody Szuwalski participated actively in person or via google virtual. William Stockhauson, Martin Dorn, and Michael Martinez also participated occasionally via google virtual.

## Focus:

The focus was to modify eastern Aleutian Islands golden king status quo assessment model EAG21.1a to model EAG21.6 and implement it in Gmacs.

## Results:

The results are compared in tables and figures.
Differences between the base model EAG21.1a and modification of the base model EAG21.6 are highlighted yellow in the following Table 1.

Table 1. Status quo and modified EAG models' differences.

| EAG21.6 (Modification of EAG21.1a) | EAG21.1a (Original model) |
| :---: | :---: |
| Data: 1981-2020 retained, 1990-2020 total, 1989-2020 groundfish discard, 1985-1998 Fish Ticket CPUE, 1995-2020 Observer CPUE, Tag release-recaptures (6 years' returns) | Data: coequal |
| 1) Equilibrium starts of simulation in 1960 with $R_{0}=$ 1987-2017 mean of mfexp(log_mean_rec)*rec_len(1) | 1) Equilibrium starts of simulation in 1960 with $\mathrm{R}_{0}=1987$ - 2017 mean of $\operatorname{mfexp}\left(\log _{1}\right.$ mean_rec+rec_dev(t))*rec_len(l) |
| 2) Recruit distribution to first five bins by gamma, using size at lower limit of the bin <br> 3) For reference points, mean $R$ is estimated as in 1) of EAG21.1a Original model | 2) Recruit distribution to first five bins by gamma, using size at mid point of the bin <br> 3) For reference points, mean $R$ is estimated as in 1). |

4) Retained size composition likelihood is multinomial without offset for size bins 1 to 17 for 1985-2020. Francis final ESS values are used
5) Total size composition likelihood is multinomial without offset for size bins 1 to 17 for 1990 to 2020. Francis final ESS values are used
6) No groundfish size composition likelihood is used
7) Observer CPUE likelihood uses log CPUE
difference residuals for 1995-2020 and reformatted as
like $1=\log ($ stddev $)+0.5 *$ square $($ residual $/$ stddev $)$, where stddev $=\mathrm{CV}$ of CPUE+model estimated additional CV
8) Fish Ticket CPUE likelihood uses log CPUE difference residuals for 1985-1998 and reformatted as like $1=\log ($ stddev $)+0.5 *$ square $($ residual $/$ stddev $)$, where stddev $=$ CV of CPUE + model estimated additional CV
9) Retained catch likelihood uses 1981-1984 catches in number of crabs and 1985-2020 catches in biomass, all transformed into log form, and dnorm(observed catch, expected catch, gmacs CV (0.032) converted to
10) Equilibrium starts of simulation in 1960 with $\mathrm{R}_{0}=1987-2017$ mean of $\operatorname{mfexp}\left(\log _{\text {_ }}\right.$ mean_rec+rec_dev(t)) *rec_len(l)
11) Recruit distribution to first five bins by gamma, using size at mid point of the bin
3 ) For reference points, mean $R$ is estimated as in 1 ).
12) Retained size composition likelihood is robust normal for size bins 6 to 17 for 1985 -2020. Francis final ESS values are used
13) Total size composition is robust normal for size bins 1 to 17 for 1990-2020. Francis final ESS values are used
14) No groundfish size composition likelihood is used
15) Observer CPUE likelihood uses log CPUE difference residuals for 1995-2020 with CPUE variance + model estimated constant variance
16) Fish Ticket CPUE likelihood uses log CPUE difference residuals for 1985-1998 with CPUE variance + model estimated constant variance
17) Retained catch likelihood uses 1981-1984 catches in number of crabs as normal likelihood with the weight of 500 and the 1985-2020 catch biomass as lognormal likelihood with the weight of 500

STD ) function applied with the emphasis factor 4 (as weight) considered in gmacs
10) Total catch likelihood uses catch biomasses for $1990-2020$ as in 9$)$ with gmacs CV ( 0.045 ) converted to STD, and the gmacs emphasis factor 2 (as weight)
11) Groundfish bycatch likelihood uses groundfish bycatch biomasses for 1989-2020 as in 9) with gmacs CV (1.58) converted to STD, and the gmacs emphasis factor 1 (as weight)
12) likelihood for pot $F$
13) likelihood for groundfish bycatch $F$
14) likelihood for tagging data
15) Additional:
a. like_rec_dev $=\operatorname{dnorm}($ rec_dev $+0.5 * \operatorname{sigR} * \operatorname{sigR}$, sigR)
where $\operatorname{sigR}=0.3535$ (for bias correction)
b. At the end added a tst*tst to the total likelihood function?
16) Reference points:
$\mathrm{B}_{35}=6,606.73 \mathrm{t} ; \mathrm{F}_{35}=0.57 ; \mathrm{OFL}=2,165.33 \mathrm{t} ; \mathrm{B} / \mathrm{B}_{35}=$
$1.095 ; \mathrm{R}_{0}=2.17722 \mathrm{mill} ; \mathrm{B}_{0}=17031 \mathrm{t}$
10) Total catch likelihood uses catch biomasses for 1990-2020 as lognormal with the graded weight going up to a maximum of 250 . Grading of weights is by observer sampled number of pots
11) Groundfish bycatch likelihood uses groundfish bycatch biomasses for 1989-2020 as lognormal with the weight of 0.2
coequal
coequal
coequal
15) like_rec_dev= $2 *$ square $\left(r e c \_d e v(t)\right)$
16) Reference points:
$\mathrm{B}_{35}=6,767.93 \mathrm{t} ; \mathrm{F}_{35}=0.61 ; \mathrm{OFL}=2,928.87 \mathrm{t} ; \mathrm{B} / \mathrm{B}_{35}=$
$1.299 ; \mathrm{R}_{0}=2.28883$ mill; $\mathrm{B}_{0}=19,376 \mathrm{t}$

During the working session, a bridging analysis was done between models EAG21.6 and EAG21.1a. Comparison of reference points between models EAG21.6 and EAG21.1a are listed in Table 2. The comparison of MMB trends are shown in Figure1.

Table 2. Estimates of reference points for various changes of the May 2021 accepted model EAG21.1a.

| Model <br> Changes | EAG21.1a <br> Base model (May 2021 accepted model) | EAG21.6 <br> Modification of base model for gmacs | EAG21.1aSid1 <br> EAG21.1a+ <br> Retained, <br> Total, and GF <br> (by) catch <br> likelihoods <br> changed to <br> EAG21.6 form | EAG21.1aSid2 <br> EAG21.1aSid1+ <br> Retained and <br> Total size comps likelihoods changed to EAG21.6 form | EAG21.1aSid3 EAG21.1aSid2+ Rec_dev bias correction factor introduced as in EAG21.6 | EAG21.1aSid4 <br> EAG21.1aSid3+ CPUE <br> likelihoods changed to EAG21.6 form |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| M | 0.21 | 0.21 | 0.21 | 0.21 | 0.21 | 0.21 |
| $\mathrm{R}_{0}$ (millions) | 2.55756 | 2.17147 | 2.44195 | 2.46983 | 2.43102 | 2.43102 |
| $\mathrm{B}_{0}(\mathrm{t})$ | 19,376 | 17,031 | 18,581 | 18,845 | 18,577 | 18,577 |
| $\mathrm{B}_{35}(\mathrm{t})$ | 6,767.93 | 6,606.73 | 6,490.46 | 6,553.45 | 6,448.36 | 6,448.36 |
| Bcurrent/ $\mathrm{B}_{35}$ | 1.299 | 1.095 | 1.222 | 1.233 | 1.067 | 1.067 |
| $\mathrm{F}_{35}$ | 0.61 | 0.57 | 0.55 | 0.55 | 0.55 | 0.55 |
| $\mathrm{F}_{\text {ofl }}$ | 0.61 | 0.57 | 0.55 | 0.55 | 0.55 | 0.55 |
| Mean Trawl | 0.00021 | 0.00023 | 0.00022 | 0.00022 | 0.00023 | 0.00023 |
| Byc F <br> Total catch OFL (t) | 2,928.87 | 2,165.33 | 2,390.62 | 2,431.11 | 2,007.42 | 2,007.42 |



Figure 1. Comparison of MMB trends for various modifications of model EAG21.1a.

After the working session, a bridging analysis was done to assess the progress of model EAG21.6 toward model EAG 21.7. The model EAG21.7 made a few improvements to EAG21.6, one-step-at a time: SigmaR was changed from 0.3535 to 0.5 , growth parameters were estimated in the model, catch and bycatch were expressed in number of crab, and observer CPUE indices were updated following May/June 2021 CPT and SSC suggestions. Furthermore, these progressions were implemented in Gmacs models. The reference points among models EAG21.1aUpdate, EAG21.6, EAG21.7, Gmacs6b (implementation of EAG21_6 in Gmacs), Gmacs7b, Gamcs7c, and Gmacs7d are compared in Table 3. The comparison of MMB trends are shown in Figure2 and the abundance by size trends among models EAG21.6, EAG21.7, and Gmacs6b are provided in Figures 3-8.

Table 3. Progression of model EAG21.6 (developed during the December 2021 working session in Juneau) toward EAG21.7 and comparison of reference points among base, modified, and Gmacs models.

|  | EAG21.1a Update | EAG21.6 | EAG21.7 | Gmacs6b | Gmacs7b | Gmacs7c | Gmacs7d |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model Changes | Base model EAG21.1a data with updated observer CPUE indices [Gmacs version of R0 and CPUE, and CPUE likelihood] | Modification of EAG21.1a for Gmacs, EAG21.1a data with status quo observer CPUE indices [Gmacs version of R0, size comp, catch, CPUE, and bycatch likelihoods] | EAG21.6+ Use EAG21.1a data with status quo CPUE indices | Convert <br> EAG21.6 <br> estimated <br> par. values <br> for input to <br> Gmacs6b.ctl, <br> use <br> Gmacs6b.dat | Convert <br> EAG21.7 <br> estimated par. <br> values for input to <br> Gmacs7b.ctl, use <br> Gmacs6b.dat | Gmacs7b+ change retained, total, and bycatch from tons to number of crab (in 1000s) in Gmacs6b.dat | Gmacs7c+ <br> Run EAG21.7 <br> with updated observer CPUE, convert EAG21.7 par. values for input to Gmacs7d.ctl |
| Additional Changes | SigmaR= 0.5 bias correction | sigmaR $=0.3535$, Growth parameters fixed to previously estimated values | sigmaR $=0.5$, <br> Growth <br> parameters <br> estimated |  |  |  |  |
| M | 0.21 | 0.21 | 0.21 | 0.21 | 0.21 | 0.21 | 0.21 |
| $\begin{aligned} & \mathrm{R}_{0} \\ & \text { (millions) } \end{aligned}$ | 2.83536 | 2.15772 | 2.12642 | 2.60732 | 3.58235 | 2.65106 | 2.69891 |
| $\mathrm{B}_{0}(\mathrm{t})$ | 25,937 | 20,058 | 19,871 | 20,280 | 24,838 | 16,322 | 16,895 |
| $\mathrm{B}_{35}(\mathrm{t})$ | 9,297.68 | 6,553.5 | 6,600.21 | 7,097.9 | 8,693.13 | 5,712.58 | 5,913.28 |
| Bcurrent/ $B_{35}$ | 1.187 | 1.132 | 1.317 | 1.265 | 1.407 | 1.345 | 1.427 |
| $\mathrm{F}_{35}$ | 0.64 | 0.56 | 0.53 | 0.59 | 0.97 | 1.67 | 1.59 |
| $\mathrm{F}_{\text {ofl }}$ | 0.64 | 0.56 | 0.53 | 0.59 | 0.97 | 1.67 | 1.59 |
| Mean <br> Trawl Byc | 0.00018 | 0.00022 | 0.00022 | 0.00022 | 0.00014 | 0.00014 | 0.00013 |
| F <br> Total catch OFL (t) | 3,795.0 | 2,240.17 | 2,714.16 | 2,876.88 | 5,309.06 | 4,349.77 | 4,912.36 |



Figure 2. Comparison of MMB trends for various modifications of EAG golden king crab model and Gmacs runs. EAG21.1a refers to the model accepted at the May/June 2021 CPT/SSC meeting whereas EAG21.1aUpdate refers to the updated model following CPT/SSC suggestions (mostly improving observer CPUE standardization).


Figure 3. Model predicted abundance by size. N matrix plot 1: 1960-1971.


Figure 4. Model predicted abundance by size. N matrix plot 2: 1972-1983.


Figure 5. Model predicted abundance by size. N matrix plot 3: 1984-1995.


Figure 6. Model predicted abundance by size. N matrix plot 4: 1996-2007.


Figure 7. Model predicted abundance by size. N matrix plot 5: 2008-2019.


Figure 8. Model predicted abundance by size. N matrix plot 6: 2020. Color key for all plots is provided here.


[^0]:    Data
    Male golden king crab carapace lengths (CL) were measured to the nearest mm and chela height (CH) measured to the nearest one-tenth of a mm by observers, and biologists from Alaska Fisheries Science Center (AFSC) and Alaska Department of Fish and Game (ADF\&G) during the commercial fishery and special surveys in the Aleutian Islands. Crab were inspected for abnormal growth due to limb loss or diseases and disregarded from measurements. There were 14,615 measurements taken during 1984, 1991, 2018 to 2021. This analysis restricts the data to the 2018/19-2020/21 fishing period with 10,815 measurements for the whole Aleutian Islands region, comprising 5,454 measurements for EAG and 5,361 measurements for WAG (Table B.2).

